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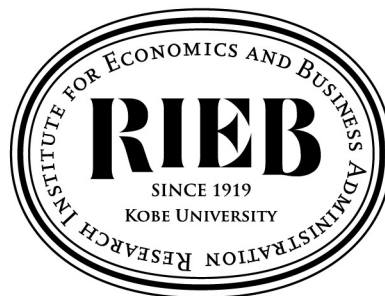
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**Hotter Days, Wider Gap:
The Distributional Impact of Heat
on Student Achievement**

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Hotter Days, Wider Gap: The Distributional Impact of Heat on Student Achievement*

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Abstract

This study demonstrates that heat disproportionately impairs human capital accumulation among low-performing students compared with their high-performing peers, using nationwide examination data from 22 million students in Japan. Given the strong correlation between academic performance and socioeconomic background, this suggests that heat exposure exacerbates pre-existing socioeconomic disparities among children. However, access to air conditioning in schools significantly mitigates these adverse effects across all achievement levels, with particularly pronounced benefits for lower-performing students. These findings suggest that public investment in school infrastructure can help reduce the unevenly distributed damage caused by heat to student learning, thereby promoting both efficiency and equity.

Keywords: Heat, Distributional impact, Student achievement, Adaptation, Air conditioning, Children, Climate change

JEL code: I21, I24, Q54

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1. Introduction

Climate change affects a wide range of societal outcomes, including human health, agriculture, labor productivity, income, cognition, and conflict, among many others (Carleton and Hsiang 2016). However, most studies are limited to measuring the *average* impact of heat, and thus its *distributional* impacts—identifying which individuals disproportionately bear the burden of heat-related damage—and the implications for inequality remain largely unexplored.

This study examines how cumulative exposure to extreme heat affects student achievement differently based on their socioeconomic statuses (SES). Cognitive ability constitutes an essential component of human capital and is closely associated with future labor market performance and the emergence of economic inequality (Cunha and Heckman 2007). Due to their physiological and neurocognitive immaturity, children’s cognitive functions may be particularly vulnerable to environmental stressors (Rowland 2008). Prolonged heat exposure can disrupt students’ learning through distractions and loss of concentration, leading to a lasting impact on their accumulation of human capital. Therefore, the warming climate underscores the importance of improving children’s learning environments.

There are two main challenges in studying the distributional consequences of heat exposure. First, assessing distributional impacts requires a representative sample. However, prior studies have been unable to assess these impacts due to their *specific* samples or the lack of individual-level data. For example, Cho (2017) and Park et al. (2020) examined the effects of cumulative heat exposure on test scores but focused on high school students taking college entrance exams in Korea and the Preliminary Scholastic Aptitude Test (PSAT) in the United States (US). These students are likely to come from higher socioeconomic backgrounds than the general population, making them unsuitable for a distributional analysis. Despite their representativeness, Park et al. (2021) used aggregated achievement data at the school district level in the US, which did not allow for distributional analysis.

Second, even if the heat damage on test scores is greater for disadvantaged students than for advantaged students, two different (but not necessarily mutually exclusive) mechanisms are consistent with the finding: students from disadvantaged households live in warmer regions and experience extreme heat (“exposure”), or because they have limited resources, which makes them more susceptible to the same heat exposure (“vulnerability”). This distinction is crucial because different policy responses are required to improve heat resilience among the poor (Hsiang et al. 2019).

To overcome these challenges, we analyze individual-level test scores from nationwide exams in Japan between 2007 and 2019 for all public-school students in grades six and nine during their compulsory education period, encompassing an extensive sample of approximately 22.8 million students. Importantly, the combination of individual-level data and the nationally

representative nature of the exams provides an ideal setting for examining, for the first time, the distributional impact of temperature on test scores.

Specifically, we analyze test scores at percentile ranks (10th, 25th, 50th, 75th, and 90th percentiles) *within* schools over time, comparing students who experienced hotter summer (colder winter) school days with those who experienced milder summer (milder winter) school days. Since a student's rank correlates with SES, such as household income and parental education, it serves as a reasonable proxy for socioeconomic background, allowing us to assess whether cumulative exposure to extreme temperatures differentially impacts student performance based on SES. Importantly, by holding the "exposure" constant for all students in the school, we can isolate the role of vulnerability—if low-SES students experience greater score declines than their high-SES peers under the same heat exposure, the disparity is likely to reflect differences in "vulnerability" by socioeconomic backgrounds (e.g., differential access to private tutoring).

This study has two major findings. Our first major finding is that the negative effects of heat are *regressive*, that is, they are far greater for low-performing students than for high-achieving students. Each additional day with the maximum temperature exceeding 34°C lowers scores by 0.09% SD for students in the top 10th percentile, but by 0.30% SD for those in the bottom 10th percentile, an impact approximately three times larger. This disproportionate effect of heat by student rank contributes to widening academic inequality, highlighting how average effects mask substantial disparities in heat-related damage between low- and high-performing students.

Importantly, by comparing temperature effects within schools, we hold school-level "exposure" constant, eliminating influences by school-level resources such as staffing ratio, teacher quality, or access to air conditioning (AC). Instead, the results likely reflect differences in "vulnerability"—how advantaged and disadvantaged students differentially adapt to the same school heat exposure. Indeed, advantaged students tend to study longer after school, spend more money on education, and are more likely to attend cram school. Given the strong link between academic performance and SES, this finding suggests that *without* any further intervention, climate change will widen pre-existing socioeconomic disparities among children.

This finding speaks to the emerging literature linking environmental and economic inequalities (Burke et al. 2015; Diffenbaugh and Burke 2019; Gilli et al. 2024). As educational attainment and earnings are positively correlated (Chetty et al. 2011), our findings suggest that environmental inequality, which exacerbates inequality in academic performance, may be a pathway through which global warming accelerates economic inequality. Specifically, our sample (grades six and nine) comprises younger students compared with samples from other studies that focused more on students nearing high school graduation (Cho 2017; Park et al. 2020). Dynamic complementarities, in which human capital investment in early childhood may complement later investments (Cunha and Heckman 2007; Johnson and Jackson 2019), indicate

that earlier heat shocks could have a more lasting impact on future economic outcomes.

Our second major finding is that adaptation through school AC can largely offset the adverse effects of heat on learning, and importantly, these benefits are *progressive*, benefiting low-performing students more than high-performing ones. In schools without AC, the negative effects of extreme temperatures are significantly more pronounced for low-performing students, who are more likely to come from low-SES backgrounds. However, in schools with AC, extreme heat has little impact on test scores across all achievement levels, resulting in a larger benefit for low-achieving students. Specifically, without AC, one extra day with the maximum temperature exceeding 34°C widens the 90th–10th score gap by 0.71% SD, but school AC reduces this widening gap by 0.55% SD. This finding suggests that public school investment, such as installing AC, can largely mitigate unevenly distributed heat damage. This is particularly encouraging because both primary and secondary education are compulsory in Japan, as in many other countries, and thus, public investment should play a critical role in improving children’s learning environments.

This second finding speaks to the literature on temperature adaptation, addressing whether environmental hazards are unavoidable or can be mitigated using current technology (See Carleton et al. 2024 for a review). While evidence supports adaptation for heat-related mortality (Barreca et al. 2016; Cohen and Dechezleprêtre 2022) and violence (Colmer and Doleac 2023), findings on workplace injuries are mixed (Dillender 2021; Park et al. 2021b). Regarding educational outcomes, a seminal study by Park et al. (2020) demonstrated that school AC reduces the cumulative impact of heat on learning. Moving beyond the average impact, we examine its distributional effects, revealing how adaptation unequally benefits students.

Furthermore, this finding contributes to the debate on the effectiveness of resource-based education policies in fostering human capital accumulation (Baron 2022; Cellini et al. 2010; Jackson et al. 2015; Lafortune et al. 2018), focusing on investments in school facilities (Lafortune and Schönholzer 2022; Martorell et al. 2016; Neilson and Zimmerman 2014). With few exceptions—such as mold remediation and ventilation (Stafford 2015) and school AC (Park et al. 2020)—prior research has rarely examined the effects of specific facility upgrades. While existing studies primarily assess the average effectiveness of such policies, we show that public investment in school infrastructure, particularly in AC, not only improves overall test scores but also disproportionately benefits lower-achieving students, thereby promoting both *efficiency* and *equity*.

2. Conceptual framework

This section outlines a simple conceptual framework for the distributional impact of extreme temperatures, based on Hsiang et al. (2019) and Behrer et al. (2021). We discuss how an

empirical observation (i.e., climate impacts are often greater for poor individuals) can mask two different explanations: differing *exposure* and/or *vulnerabilities*. Exposure refers to the degree to which individuals are subjected to environmental stressors (e.g., heat, cold, and pollution), while vulnerability denotes the extent of their susceptibility to these stressors.

Damage from environmental stressors is defined as a function of two factors: exposure and vulnerability. Here, we aim to estimate the marginal damage, which is the slope of the damage function. Importantly, it can vary by SES (e.g., income, education, and occupation) for two reasons, as shown in Figure 1. We assume that exposure is higher for low-SES individuals than for high-SES ones because the poor tend to live in hotter places, both within and across countries (Park et al. 2018).

First, as shown in panel A of Figure 1, a convex damage function with respect to exposure can lead to greater marginal damage for low-SES individuals who experience more exposure than their high-SES counterparts (i.e., differential exposure).¹ Alternatively (or additionally), as illustrated in panel B of Figure 1, the damage function itself may differ by SES due to factors such as baseline health or defensive investments correlating with SES (i.e., differential vulnerability).

Distinguishing whether SES-based disparities in heat impact stem from a single nonlinear damage function with differential exposure or differing vulnerabilities is crucial for policy design. The former requires reducing direct contact with extreme heat (e.g., urban cooling, housing interventions, and warning systems). Conversely, the latter requires targeted support to enhance adaptive capacity (e.g., subsidizing AC and expanding medical programs to address heat-related illnesses) or promoting broader poverty reduction to strengthen the heat resilience of low-SES individuals.

However, distinguishing between exposure and vulnerability is challenging because the poor tend to live in hotter locations. Thus, even if heat damage is greater among the poor, this may simply result from differential exposure (panel A) rather than differential vulnerability (panel B) by SES. This study is the first to rigorously isolate the impact of vulnerability from exposure. Using individual-level data from nationally representative exams, we analyze the distributional impact of temperature *within* schools, while holding exposure constant, at least in the school environment where most learning is supposed to occur. If the reductions in test scores are greater for low-SES students than for high-SES counterparts in the same school with identical heat exposure, the difference in marginal damages likely arises from varying vulnerabilities between these groups.

¹ Park et al. (2018) demonstrated that the poor tend to live in hotter locations both within and across countries.

3. Data

We combine temperature data with nationwide test data of nearly 22.8 million students in Japan. Appendix B provides details of the data sources. We discuss the school AC penetration data in Section 6.

3.1. Test scores

We use data on the nationwide exams, called the National Assessment of Academic Ability (hereinafter “NAAA”), conducted by the Ministry of Education, Culture, Sports, Science and Technology (MEXT). The NAAA aims to monitor the academic performance and progress of students nationwide and contribute to the improvement of educational policies (MEXT 2024a). The NAAA has been conducted annually since 2007, except in 2011, when the NAAA was completely canceled because of the Great East Japan Earthquake, and in 2010 and 2012, when the NAAA was administered to a random subset of schools.²

The NAAA is administered to students in their final years of public primary (grade six) and secondary school (grade nine).³ Both primary and secondary education are compulsory in Japan. Nearly 100% of public primary and secondary schools participate in the NAAA (NIER 2024). Although the subjects assessed vary slightly over time, we focus on reading and mathematics, which were consistently tested throughout our sample period.

The NAAA is held on the 3rd or 4th Tuesday of April⁴, the month when the academic year begins in Japan. Consequently, the NAAA is designed to assess students’ understanding of the material covered until the previous academic year (NIER 2021). This timing aligns well with our research design on learning disruptions from the past summer and winter. Since the exam date is predetermined and the NAAA is centrally administered and graded, no room exists for endogenous choice in the timing of test-taking or score manipulation by schools and students.

The NAAA is not a high-stakes exam for students or schools. Students’ scores do not affect their promotion to higher grades or better schools. Furthermore, school performance has no direct consequences, such as reduced federal funding, unlike test-based accountability systems such as the No Child Left Behind Act in the US. The only potential stakes are reputation concerns for schools (Morozumi and Tanaka 2023); however, publication of school-level scores is not allowed for years before 2014, and very few school councils do so in our sample period.

We use 2007–2019 NAAA data with MEXT’s permission for the secondary use of confidential information. Table A.1 details the number of participating schools and students each

² This is entirely due to political reasons. In 2009, a change of government occurred, and the new administration chose to cut the NAAA’s budget.

³ In 2022, 1.3% of primary and 7.7% of secondary students attended private schools (MEXT 2022).

⁴ From 2019, it was held on Thursday instead of Tuesday.

year. From 2007 to 2019, approximately 22.8 million students took the exams, with approximately 30,000 schools participating annually (excluding 2010 and 2012). See Figure A.1 for the school locations across the country. For statistical power, we combine both grades in the main analysis unless stated otherwise. Since the exams are administered to both grades on the same day, all students experience identical conditions, including cumulative heat exposure and test-day weather.

Our primary outcome is the combined reading and math scores, although we also separately analyze each subject. Since exam difficulty varies by year, we calculate z-scores for each year and grade and multiply them by 100 for interpretation as percentage changes. Student-level data include limited demographics such as gender. The NAAA also conducts student surveys in every round and parental surveys in 2013 and 2017. Student surveys capture behaviors (e.g., after-school study and study habits), while parental surveys (administered to about 4.8% of randomly selected schools)⁵ collect household information such as household income, father's occupation, and parental education. Table A.2 (panel A) provides descriptive statistics of the individual characteristics.

3.2. Temperature

We use daily temperature data for 2006–2018 from the Japan Automated Meteorological Data Acquisition System (AMeDAS) operated by the Japan Meteorological Agency. We utilize AMeDAS data from a subset of 899 weather stations that have daily temperature information available for at least 99% of the days from 2006 to 2018. To create a balanced panel, missing daily observations were imputed using the nearest station with complete data. Each school was then assigned to its nearest weather station to ensure that our estimates remain unaffected by changes in the number or location of the stations.

Figure A.2 displays the locations of all 899 weather stations as of 2018 and the cumulative distribution of the distance from the nearest station to each school. The density of stations is high, given the country's size, and consequently, the mean (median) distance is 6.95 (6.48) km, compared with 15.6 km in the US (Park et al. 2020).

Our primary measure of cumulative exposure to extreme temperatures is the number of hot and cold school days that a student experienced in the year leading up to the test in April (i.e., from April of the previous year to March of the test year). We utilize the daily maximum temperature as a measure, since schooling occurs during daytime hours when peak temperatures are typically observed. Following Park et al. (2020), we focus on temperatures during terms as school days and treat school break days and weekends during terms as separate non-school

⁵ In 2013 and 2017, parental surveys covered 2,821 of 59,734 schools (4.72%) and 203,023 of 4,255,669 students (4.77%), with an 84.9% response rate.

days.⁶

We also use weather station data to construct both cumulative and test-date measures for rainfall, wind speed, and relative humidity. We also include pollution data from the nearest monitoring station, as it is known to impact short-term cognition (e.g., Ebenstein et al. 2016). Table A.2 (panel B) presents the descriptive statistics of the weather conditions.

4. Econometric model

4.1. Estimation of the average of marginal damages

We exploit year-to-year variations in the number of hot and cold school days to identify the causal impact of exposure to extreme temperatures on human capital accumulation. Specifically, we compare the test scores of students in the same school who experienced hotter summers or cooler winters with those exposed to milder conditions.

Figure A.4 shows both the spatial and temporal variations in the daily maximum temperature that students experienced from last April to March of the test year, highlighting significant climate differences across the region and considerable year-to-year variations in both cold and hot school days.

To reduce the computational burden, we collapse the data into school-year cells and weigh all estimates by the number of students in each cell. Specifically, we estimate the following specifications:

$$\text{Average_Z-score}_{st} = \sum_k \beta^k T_{st}^k + \rho_s + \theta_t + \delta X'_{st} + \varepsilon_{st}, [1]$$

where the dependent variable is the average z-score for school s in year t . T_{st}^k represents the number of school days in the prior year where the maximum temperature falls into one of nine bins k : below 6°C, 6–10°C, 10–14°C, 14–18°C, 22–26°C, 26–30°C, 30–34°C, and above 34°C, with 18–22°C as the reference, the optimal range for test performance.

This specification enabled us to flexibly capture the nonlinear temperature effects. The coefficient of interest are β^k . ρ_s and θ_t are school FE and year FE, respectively. X'_{st} includes other time-varying school-level controls, such as precipitation, humidity, and pollution. Standard errors are clustered at the weather station level (N=889) to account for potential serial correlations reflecting the underlying variations in our treatment variable (Abadie et al. 2023). The underlying assumption for β^k to reflect the causal impact of temperature is that the temporal and geographic variations in prior-year temperature are uncorrelated with unobserved determinants of student learning.

⁶ School days, school break days, and weekends during the terms are mutually exclusive, averaging 212.6, 85.1, and 67.6 days, respectively. Lacking a comprehensive national school calendar dataset, we assign each school a probable start and end date using the 2018 calendar of its prefectural capital (Figure A.3). Colder regions tend to have shorter summer and longer winter breaks, while warmer regions exhibit a reverse pattern.

To visualize the identifying variations underlying the baseline specification, we plot residuals from a regression of the number of school days $<6^{\circ}\text{C}$ and $>34^{\circ}\text{C}$ against school fixed effects. Figure A.5 illustrates the interquartile and interdecile ranges of the residual variations by prefecture and year. These distributions confirm ample variations in the number of extreme-temperature school days within each prefecture and each year, ensuring that our estimates are not driven by variations in a specific region or year.

4.2. Estimating heterogeneous marginal damages

This study's main contribution is that it moves beyond the effect of temperature on average test scores (Equation [1]) and examines its distributional impacts. Using individual test scores linked to school IDs, we assess the effect of the temperature by the score rank within schools. Specifically, for each school, we compute the z-scores at the 10th, 25th, 50th, 75th, and 90th percentiles within schools. We then run each value separately as the outcome as follows:

$$\begin{aligned} & \text{Z-score at } X \text{ percentile } (X = 10, 25, 50, 75, \text{ and } 90)_{st} \\ &= \sum_k \beta^k T_{st}^k + \rho_s + \theta_t + \delta X'_{st} + \varepsilon_{st}. \quad [2] \end{aligned}$$

We occasionally use the test score gap at different percentiles (e.g., the 90th-10th test gap) as the outcome.

What does the student's within-school rank capture? Using the 2013 and 2017 NAAA surveys of parents in a subset of schools, Figure 2 illustrates a strong positive correlation between student's rank and socioeconomic background, namely household income (panel A), and father's education (panel B). The income gap between the 90th and 10th percentiles is 1.79 million yen (approximately 17.9K USD), while the gap in fathers' university education is 30.4 percentage points. Overall, we posit that a student's within-school rank is largely indicative of their socioeconomic status.

Finally, we demonstrate that the variations in scores within schools reflect most of the variation in scores at the national level. Figure A.6 shows the within-school score distribution by school rank, grouping schools into ventiles based on each year's average scores. While higher-ranked schools have more compressed score distributions, considerable within-school variations exist across all ranks. This addresses the concern that within-school test score variations are small and potentially missing larger national-level variations in test scores.⁷

⁷ The decomposition of the variation in test scores shows that as much as 91–93% occurs *within* schools rather than *between* schools over the years, likely due to the relatively uniform quality of public schools compared with private ones. Furthermore, the school curriculum is uniformly determined by the MEXT's Course of Study.

5. Baseline results

5.1. Average impacts

First, we present graphical evidence of the average impacts of cumulative exposure to heat and cold on test scores. Figure 3A shows β^k from equation [1] with 95% confidence intervals. The test scores are measured in 0.01σ . Highlighting the nonlinear effect of temperature on learning, the figure shows that test scores decline as the number of hot or cold school days increases, especially for the extremely hot days at the right end of the figure ($>34^\circ\text{C}$) and the extremely cold days at the left end ($<6^\circ\text{C}$).

This aligns with the well-documented “U-shaped” mortality-temperature relationship (or “inverse-U” in our case, as damage is negative), where both hot and cold days increase mortality globally (e.g., Barreca et al. 2016; Carleton et al. 2022; Cohen and Dechezleprêtre 2022; Heutel et al. 2021). While some studies have examined the cumulative effects of heat (Cho 2017; Park et al. 2020, 2021a) and cold (Johnston et al. 2021) on test scores separately, we are the first to show that both extremes in the same country impair students’ learning environments and hinder teachers’ abilities to teach by causing distractions and a loss of concentration.

In terms of magnitude, one additional school day $<6^\circ\text{C}$ or $>34^\circ\text{C}$ in the previous year (compared with $18\text{--}22^\circ\text{C}$) reduces test scores by 0.13% SD and 0.19% SD, respectively ($p < 0.01$). These estimates align with prior research on the effects of cumulative exposure to heat or cold on test scores, as shown in Table A.3.

5.2. Distributional impacts

Next, we examine whether the negative impacts of extreme temperatures significantly vary among students across different score distributions. Figure 3B presents β^k from Equation [2], which clearly indicates that the negative effects of extreme temperatures are significantly greater for lower-performing students (see Table A.4 for corresponding estimates).

One additional hot day $>34^\circ\text{C}$ lowers scores by 0.09% SD for students in the top 10th percentile, while the impact on the bottom 10th percentile is 0.30% SD, which is approximately three times larger. Adverse effects consistently increase as the rank decreases. Similarly, an extra cold day $<6^\circ\text{C}$ leads to a negligible reduction of 0.03% SD for students in the top 10th percentile (not statistically significant), while the bottom 10th percentile experiences a decline of 0.26% SD. Consequently, both the extremely hot and cold conditions widens the 90th–10th score gap by 0.22% and 0.23% SD, respectively. Given the strong link between academic performance and SES (Figure 2), these results suggest that exposure to extreme temperatures exacerbates pre-existing academic inequality by SES among children.

Source of varying vulnerability—. Importantly, since we compare temperature effects within schools, keeping “exposure” at school constant, our results are not driven by school

resources (e.g., class size, teacher quality, or AC). Instead, they are consistent with those in panel B of Figure 1, likely reflecting “vulnerability”—individual or household adaptations outside school (e.g., private tutoring). This study’s main goal is to uncover the *presence* of socioeconomic disparities in vulnerability to extreme temperatures. Consequently, it is beyond the scope of this study to fully explore the underlying *sources* of such heterogeneity in vulnerability owing to limited data on detailed student and household behaviors during the hot and cold days of the previous summer and winter.

Nevertheless, Figure A.7 shows that higher-SES students tend to study longer after school, spend more money on education, and are more likely to attend cram school. Additionally, Table A.5 suggests that longer after-school study hours may mitigate the negative effects of heat exposure.⁸ However, other factors such as better baseline health among higher-SES students (Case et al. 2002), may also contribute to the observed heterogeneity. Understanding the specific sources of these unequal vulnerabilities is an avenue for future research.

6. The impact of AC

6.1. AC penetration

AC is the main technology for adapting to heat (Barreca et al. 2016), but its widespread adoption in public primary and secondary schools in Japan has occurred only recently. During the sample period from 2006 to 2018, AC coverage in public primary and secondary schools increased from approximately 10% to 50%, reaching nearly 100% by 2022.⁹

Unfortunately, the government began reporting the penetration rates of school AC in public primary and secondary schools at the municipal level only in 2017 (MEXT 2024b). The school council of each municipality determines the installation of AC in public schools within the municipality.¹⁰ Using this data in 2018, the last year of the sample period, we categorize schools into municipalities with 0% (“schools without AC”), 100% (“schools with AC”), and intermediate AC penetration. Thus, schools without AC had no AC throughout the *entire* period of 2006–2018 without any measurement error. Conversely, schools with AC only indicate full availability *at some point* during the sample period, likely leading to an underestimation of the positive impact of AC on test scores.

Figure 4A maps municipalities with 100% (“schools with AC”), 0% (“schools without AC”), and partial (>0% and <100%) AC penetration. Clearly, schools without AC are more common in the cooler northern region, while both with and without AC are widely distributed in

⁸ Educational spending and cram school attendance data are limited to parental surveys from 2014 and 2017, covering only 4.7% of students. Thus, unlike after-school hours from student surveys available for all years, they cannot be included as mediators.

⁹ Source: https://www.mext.go.jp/content/20240930-mxt_sisetujo01-000013462_02.pdf

¹⁰ A total of 1,724 municipalities exist as of April 1, 2019.

central Honshu, Japan’s main and largest island.

One concern is that school AC penetration may correlate with many factors at the school or municipal levels that could directly impact test scores. However, Figure A.8 shows that after controlling for the average temperature, the AC penetration rate in 2018 is not strongly linked to taxable income per capita or the student-to-teacher ratio, a measure of per-pupil educational expenditure at school. As school AC could still correlate with other adaptive technologies or resources that may independently enhance student learning, the results below should be interpreted with caution.

Furthermore, Figure A.9 shows that in 2007 (the first year of the our sample period), when AC penetration was only 10.2%, the test score distributions between schools with and without AC were nearly identical, suggesting that the observed differential patterns are unlikely to be due to systematic differences in test scores between the two groups (e.g., the 10th percentile score with AC matches the 90th percentile score without AC).

6.2. Average impacts

We now examine the average impact of access to school AC on test scores. Figure 4B illustrates β^k from Equation [1] separately for schools with and without AC. Strikingly, most of the negative effects of heat are concentrated in schools that lack AC throughout the sample period. Conversely, AC largely mitigates the adverse impact on learning if taken causally.

To assess how effectively school AC mitigates the impact of heat on learning, we conduct a formal regression analysis. Specifically, we interact the cross-sectional measure of AC penetration in 2018 (“school AC” dummy; AC_s) with the number of school days in each temperature bin and include them in our baseline specification [1]. To highlight the effect of AC availability, we focus on schools in municipalities with either 0% or 100% AC (56.9% of school-year observations). However, as Figure A.10 shows, results remain robust when including schools from municipalities with partial AC ($>0\%$ & $<100\%$) in the “with AC” category. Specifically, we estimate

$$\text{Average_Z-score}_{st} = \sum_k \beta^k T_{st}^k + \sum_k \gamma^k T_{st}^k * AC_s + \rho_s + \theta_t + \tau_t * AC_s + \delta X'_{st} + \varepsilon_{st}, [3]$$

where β^k now measures the impact of heat on a school without AC, while γ^k represents the difference in that impact compared to a fully air-conditioned school. Column (1) of Table 1 shows that school AC largely offsets the negative effects of extreme heat ($>34^\circ\text{C}$). Without AC, test scores drop by 0.56% SD, but the interaction with the AC dummy reduces this by 0.41% SD, suggesting that AC mitigates approximately 73% of the adverse impact of heat on learning.

The offsetting effect of school AC may reflect other factors that correlate with AC availability. To address this concern, Table A.6 controls for interactions between temperature

bins, municipality-level taxable income per capita, the student-teacher ratio, and home AC share, but the results remain robust.

Other robustness—Table A.7 presents additional robustness checks. The estimates for the days with the maximum temperature exceeding 34°C and their interaction with the school AC dummy are reported due to their greatest relevance to global warming. The estimates remain largely unaffected by controlling for test-day weather conditions (temperature, precipitation, wind speed, and humidity), test-day air pollution, cumulative weather conditions other than temperature, and hot days during non-school periods.¹¹

Heterogeneity—Figure A.12 and Table A.8 explore the heterogeneous effects of heat >34°C and the mitigating role of school AC across grades (6th vs. 9th), subjects (math vs. reading), gender (girls vs. boys), question difficulty (basic vs. advanced),¹² and climate (cool vs. warm regions). Overall, the impact of temperature >34°C and offsetting effect of AC appear to be consistent across contexts, with a few notable exceptions that extreme heat affects 6th graders and boys more than 9th graders and girls by approximately 50%, suggesting their greater vulnerability to heat. Notably, school AC offsets the effect on basic but not advanced questions, aligning with its stronger benefit for lower-achieving students, as shown next.

6.3. Distributional impact

Finally, we analyze how the impact of school AC availability differs among students across various score distributions. Figures 4C and 4D present β^k from Equation [2], separately for schools without AC and with AC. In schools without AC, heat disproportionately harms lower-ranked students, whereas in schools with AC, nearly all the negative effects disappear across ranks, resulting in a larger benefit for low-achieving students. As expected, school AC does not affect performance under extremely cold conditions (<6°C). However, as shown in Figure 4C, without school AC, heat is much more likely to exacerbate pre-existing academic inequalities than cold, without any intervention.

To formally assess how school AC mitigates heat-driven inequality in learning, we estimate a variant of Equation [3], where the outcomes are z-scores at the 10th, 25th, 50th, 75th, and 90th percentiles within schools. The estimates from other percentiles can be found in Table A.9.

Column (2) shows that high temperatures (>34°C) without AC reduce scores at the 10th

¹¹ Figure A.11, which uses minimum rather than maximum temperatures, shows no discernible effect on test scores, suggesting that sleep disruption due to nighttime temperatures (Mullins and White 2019) is not the channel through which cumulative heat exposure negatively affects academic performance.

¹² Both math and reading included basic and advanced questions (until 2018), with basic skills practically applied to advanced ones. For example, in 6th grade math, a basic question asks for simple multiplication, while an advanced one requires using it to find a square's area (Figure A.13). The two scores are highly correlated, with correlations of 0.90 (average), 0.83 (math), and 0.85 (reading) for 6th graders.

percentile by 0.93% SD. Since these schools lack AC, the estimates reflect the “pure” negative impact of heat without reflecting any offsetting effects.¹³ However, the interaction term is positive and as large as 0.69% SD ($p < 0.01$), indicating that school AC significantly offsets the damage from heat exposure. Conversely, column (3) shows that high temperatures ($> 34^{\circ}\text{C}$) reduce scores at the 90th percentile only by 0.22% SD ($p < 0.01$), while the offsetting effect of AC is 0.14% SD, albeit not statistically significant. Consequently, column (4) indicates that without AC, extreme heat widens the 90th–10th score gap by 0.71% SD, whereas school AC reduces this widening gap by 0.55% SD, suggesting that the benefit of school AC is *progressive*.¹⁴

We demonstrate that school facilities help reduce the widening test score gap between advantaged and disadvantaged students caused by heat. This suggests that the widening gap in the absence of school AC is not primarily caused by differences in *outside-of-school* heat exposure (e.g., longer commutes for disadvantaged students); if this were the case, we would not expect school AC to counteract the widening of the achievement gap. Simultaneously, school AC did not fully offset the growing gap, likely because of measurement errors in the AC penetration measure and/or remaining outside-of-school adaptations by socioeconomic background (e.g., access to clam school).¹⁵

This finding suggests that public investment in school AC, rather than household-level adaptation, can largely reduce heat’s inequality-enhancing negative effects. Thus, adequate investment in school infrastructure can mitigate unevenly distributed damage caused by heat to student learning, thereby promoting both efficiency and equity. This is encouraging because both primary and secondary education are mandatory in Japan, as in many other countries, where public investment plays a vital role. Moreover, Dechezleprêtre et al. (2025) demonstrate that climate policies that are both environmentally effective and distributionally progressive are more likely to garner public support. However, it should be emphasized that while school AC largely offsets the widening of socioeconomic inequalities, pre-existing socioeconomic disparities *persist*.

7. Conclusion

Many studies have investigated the average impact of extreme temperatures but their

¹³ Conversely, the distributional impact of a cold day (temperature $< 6^{\circ}\text{C}$) is similar for schools with and without school AC.

¹⁴ Table A.10 confirms that the impact of school AC on the 90th–10th score gap remains robust when controlling for interactions with municipality-level taxable income per capita, the student-teacher ratio, and prefecture-level home AC share.

¹⁵ We cannot entirely dismiss the possibility that this persistent widening gap stems from differing *outside-of-school* exposure, resulting in varying *in-school* vulnerabilities. For example, limited access to AC at home deteriorates sleep quality (outside-of-school exposure), which in turn leads to diminished focus and concentration at school (in-school vulnerability), even within the same classroom environment.

distributional impact across different SES remains poorly understood. Even less explored is how different socioeconomic groups adapt to environmental stressors such as heat. Using nationwide exam data from Japan for 2007–2019, we find that extreme temperatures disproportionately hinder the human capital accumulation of low-achieving students, deepening academic and social inequalities. However, school AC largely offsets these negative effects, highlighting the potential for public infrastructure investments to reduce heat-related learning disparities.

This study offers several avenues for future research. First, it is essential to determine whether the inequality-enhancing effects of heat exposure on learning persist across different contexts and environments. Second, although we focus on heat damage because of its relevance to global warming, understanding how to mitigate the adverse effects of cold exposure, although smaller, may be important in specific situations. Third, while we highlight the presence of social disparities in vulnerabilities, understanding the sources of these differential vulnerabilities, supported by more comprehensive data on individual and household behaviors, is essential for addressing social disparities. Finally, it is also important to examine whether the inequality-enhancing effects of heat exposure on learning translate into inequalities in long-term economic outcomes such as wages and income.

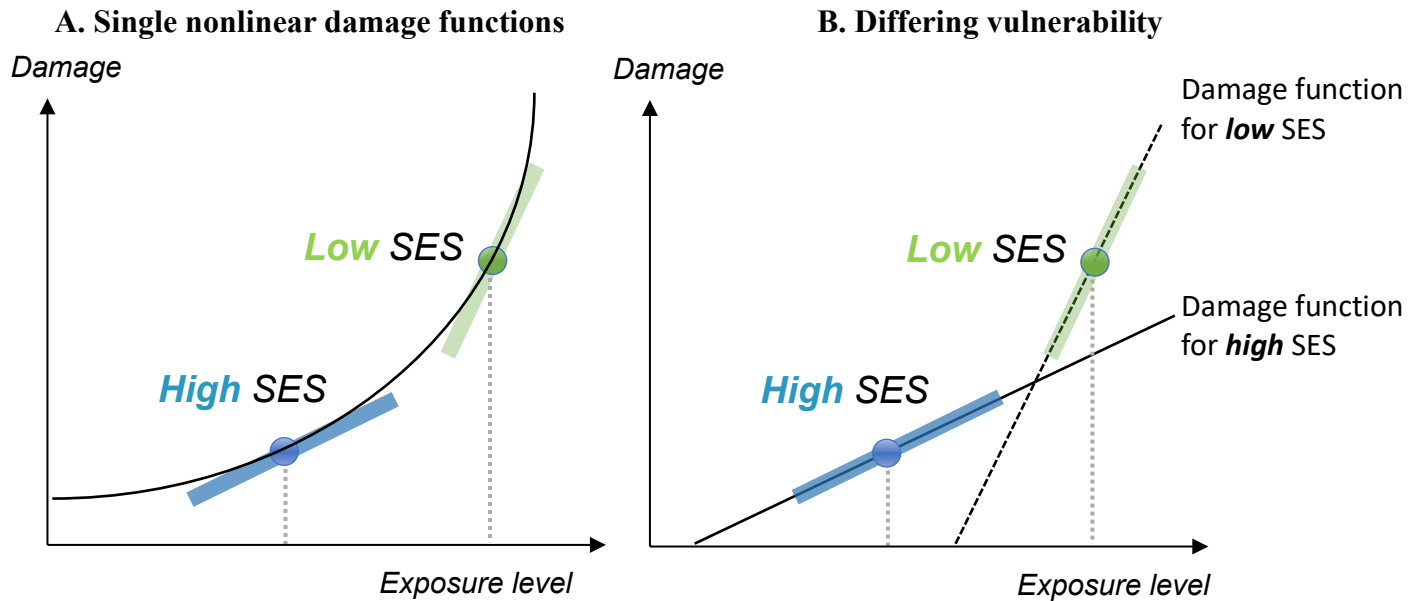
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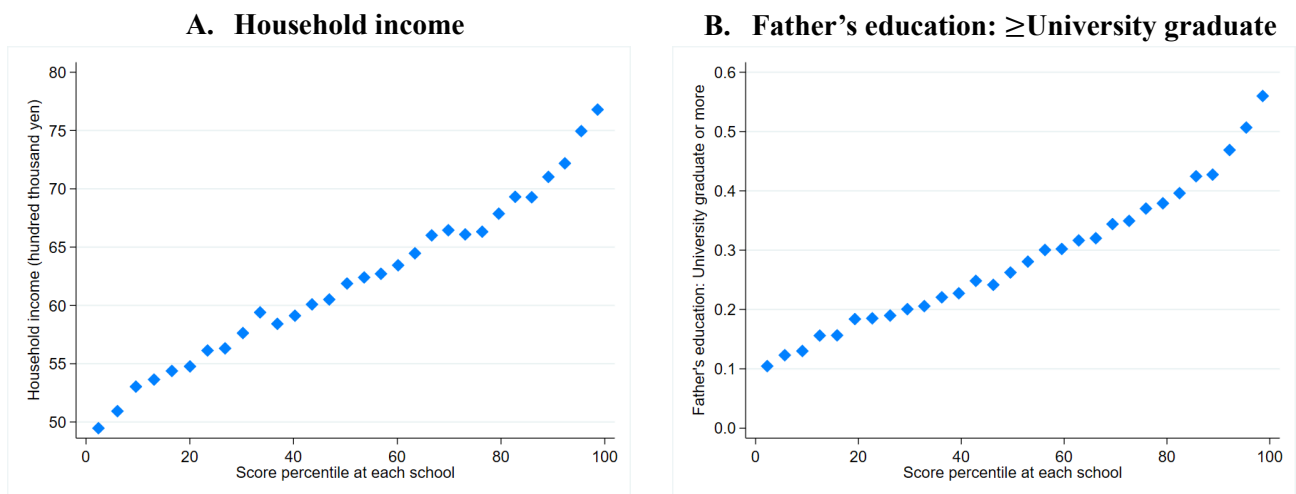
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Figure 1—Heterogeneity in marginal damages from two different explanations



Notes: Adapted from Hsiang et al. (2019, Figure 1), this figure presents two different explanations for the empirically observed heterogeneity in marginal damages between high and low socioeconomic status (SES): a single nonlinear damage function, illustrated in panel A, or different damage functions (i.e., differential vulnerability) related to SES that correlate with exposure levels, as shown in panel B.

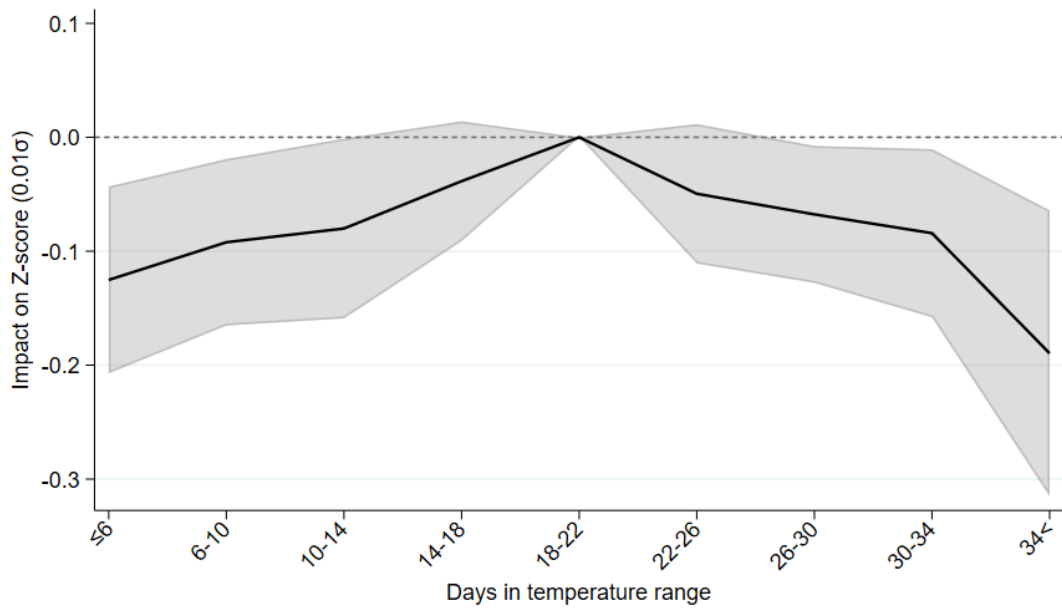
Figure 2—Within-school student rank and socioeconomic status



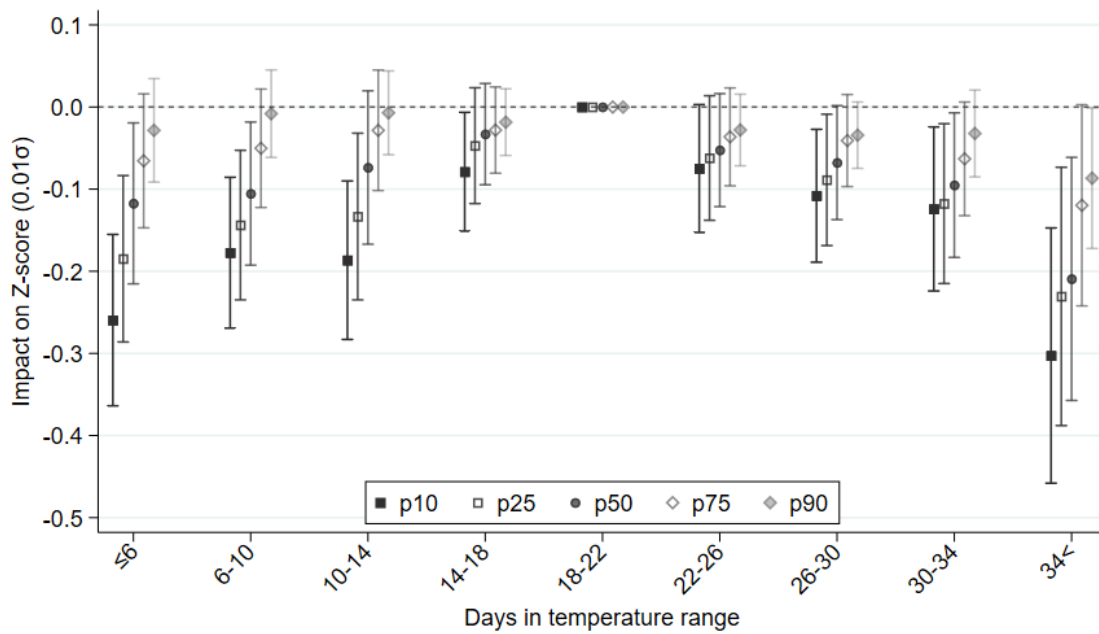
Notes: The data are obtained from parent surveys in 2013 and 2017 NAAA. The bin scatter plot illustrates the relationship between within-school student rank and various measures of students' socioeconomic status, net of school fixed effects, specifically household income (panel A), and the proportion of fathers with education at or above a 4-year university/college degree (panel B). Household income (panel A) is reported in hundreds of thousands of yen, with US\$1 equal to approximately 100 yen. We transform the median of each household income bin into a continuous variable.

Figure 3—Cumulative heat/cold exposure and test performance

A. Average impacts



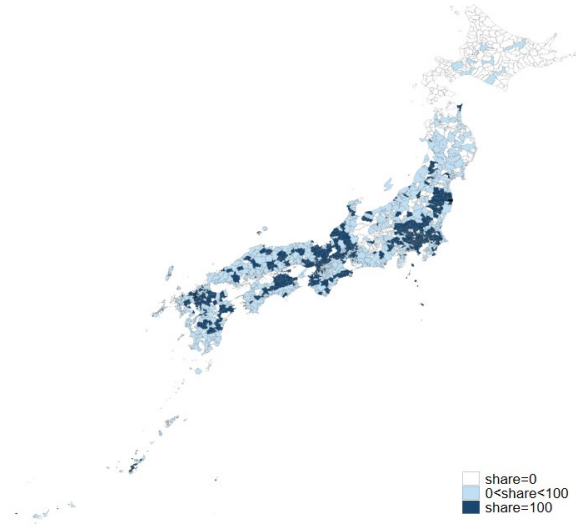
B. Distributional impacts



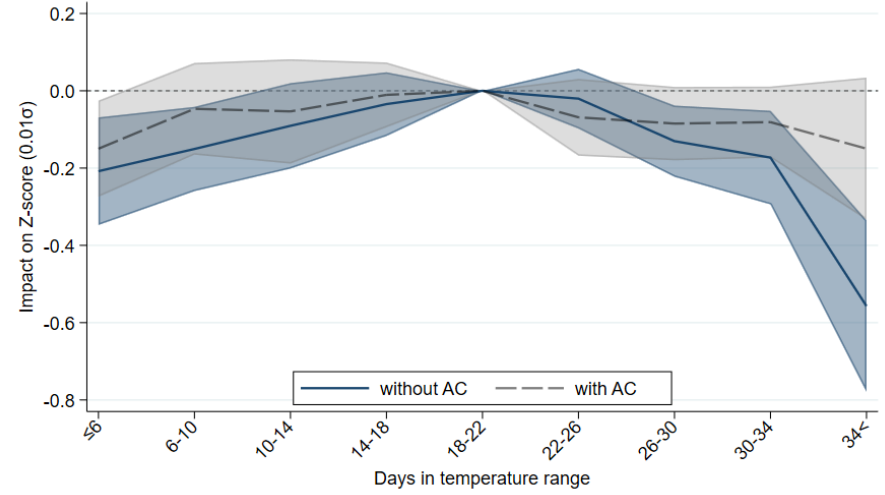
Notes: Panel A plots β^k from an estimating Equation [1], where the average z-score (measured in 0.01σ) is regressed on the number of school days within a given maximum temperature bin in the year prior to the test date, along with the 95% confidence intervals. Panel B plots β^k from an estimating Equation [2], where z-scores at the 10th, 25th, 50th, 75th, and 90th percentiles within schools (measured in 0.01σ) are regressed separately on the number of school days within a given maximum temperature bin from the year prior to the test date, along with the 95% confidence intervals. The omitted category is the temperature range between 18–22°C.

Figure 4—The impact of school AC

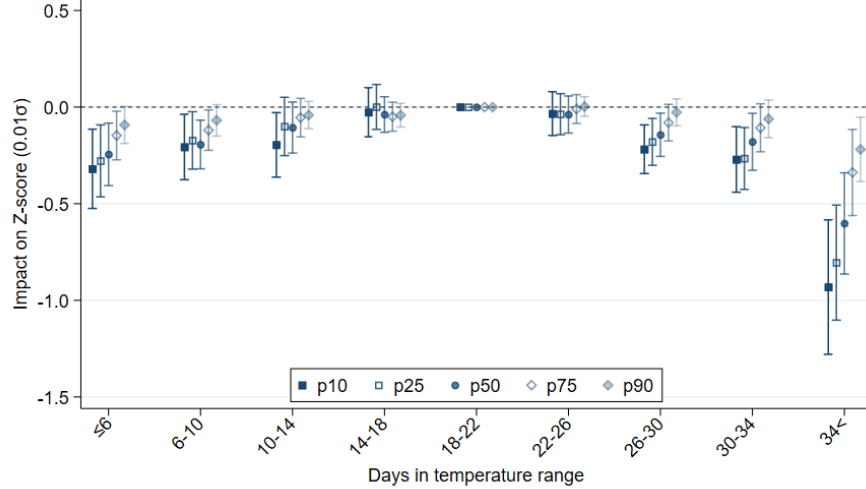
A. Map of the school AC penetration



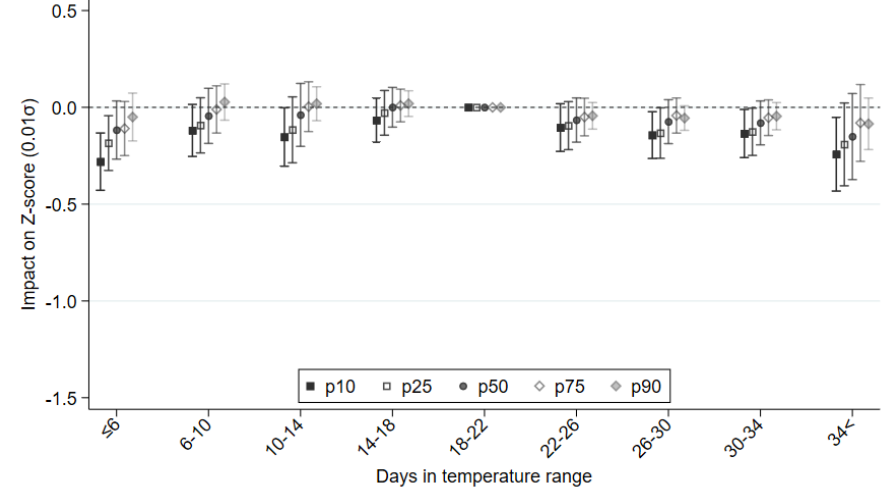
B. Average impacts



C. Distributional impact *without* school AC



D. Distributional impact *with* school AC



Notes: Panel A displays the locations of municipalities according to the degree of school AC penetration rate. Using school AC penetration rates for public primary and secondary schools at the municipal level in 2018 (the last year of the sample period), schools are categorized into municipalities with a 0% share (in white), a 100% share (in dark blue), and the remaining (in light blue) of school AC penetration as of 2018. Panel B plots β^k from an estimating Equation [1]. Panels C and D plot β^k from an estimating Equation [2], separately for schools without AC and with AC in 2018, respectively, along with the 95% confidence intervals. The omitted category is the temperature range between 18–22°C.

Table 1—The impact of school AC

| Outcomes: | (1) | | (2) | | (3) | | (4) | |
|--------------|----------------------|---------------------|-----------------------------------|---------------------|-----------------------------------|-------------------|--|----------------------|
| | Average Z-score | | 10 th percentile score | | 90 th percentile score | | 90 th -10 th score gap | |
| | × school AC | | × school AC | | × school AC | | × school AC | |
| Days 6°C≤ | -0.216*** (0.069) | 0.068 (0.090) | -0.320*** (0.101) | 0.040 (0.124) | -0.099** (0.047) | 0.050 (0.076) | 0.222** (0.086) | 0.011 (0.110) |
| Days 6-10°C | -0.158*** (0.053) | 0.114 (0.076) | -0.207** (0.083) | 0.088 (0.103) | -0.075* (0.040) | 0.104* (0.059) | 0.132* (0.078) | 0.016 (0.089) |
| Days 10-14°C | -0.098* (0.054) | 0.046 (0.084) | -0.196** (0.082) | 0.043 (0.108) | -0.047 (0.035) | 0.067 (0.053) | 0.149** (0.074) | 0.025 (0.088) |
| Days 14-18°C | -0.040 (0.041) | 0.031 (0.057) | -0.027 (0.062) | -0.039 (0.083) | -0.046 (0.030) | 0.067 (0.042) | -0.019 (0.052) | 0.106 (0.078) |
| Days 22-26°C | -0.024 (0.039) | -0.043 (0.063) | -0.034 (0.057) | -0.070 (0.083) | -0.000 (0.026) | -0.042 (0.043) | 0.034 (0.050) | 0.028 (0.068) |
| Days 26-30°C | -0.134*** (0.046) | 0.050 (0.066) | -0.218*** (0.063) | 0.075 (0.087) | -0.030 (0.035) | -0.024 (0.046) | 0.188*** (0.053) | -0.099 (0.075) |
| Days 30-34°C | -0.177*** (0.061) | 0.097 (0.073) | -0.271*** (0.085) | 0.136 (0.102) | -0.065 (0.049) | 0.020 (0.058) | 0.207*** (0.076) | -0.116 (0.094) |
| Days 34°C> | -0.562*** (0.112) | 0.413*** (0.145) | -0.932*** (0.176) | 0.690*** (0.201) | -0.223*** (0.085) | 0.139 (0.108) | 0.709*** (0.170) | -0.551*** (0.184) |
| R-squared | 0.751 | | 0.669 | | 0.603 | | 0.552 | |
| Observations | 190,210 | | 190,210 | | 190,210 | | 190,210 | |

Notes: The data are from the 2007–2019 NAAA and 2006–2018 AMeDAS. The unit of observation is the school-year. Column (1) presents the estimates from Equation [3] where the outcome is the average test score at the school-year level (measured in 0.01σ). Columns (2) and (3) present the estimates from the variant of Equation [3], where the outcomes are z-scores at the 10th and 90th percentiles within schools (measured in 0.01σ). School AC is a dummy variable that equals one if an air conditioner was available at the school in 2018. Figure 4A shows the locations of the schools within each AC penetration category. The complete table showing the results for the other percentiles is presented in Table A.9. Column (4) presents the estimate of the score gap between the 90th and 10th percentiles within the school (measured in 0.01σ). Standard errors are clustered at the weather station level in parentheses. The estimates are weighted by the number of students in each school-year. The omitted category is the temperature range between 18 and 22°C. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Online Appendix (Not for Publication)

Hotter Days, Wider Gap: The Distributional Impact of Heat on Student Achievement

By Mika Akesaka and Hitoshi Shigeoka

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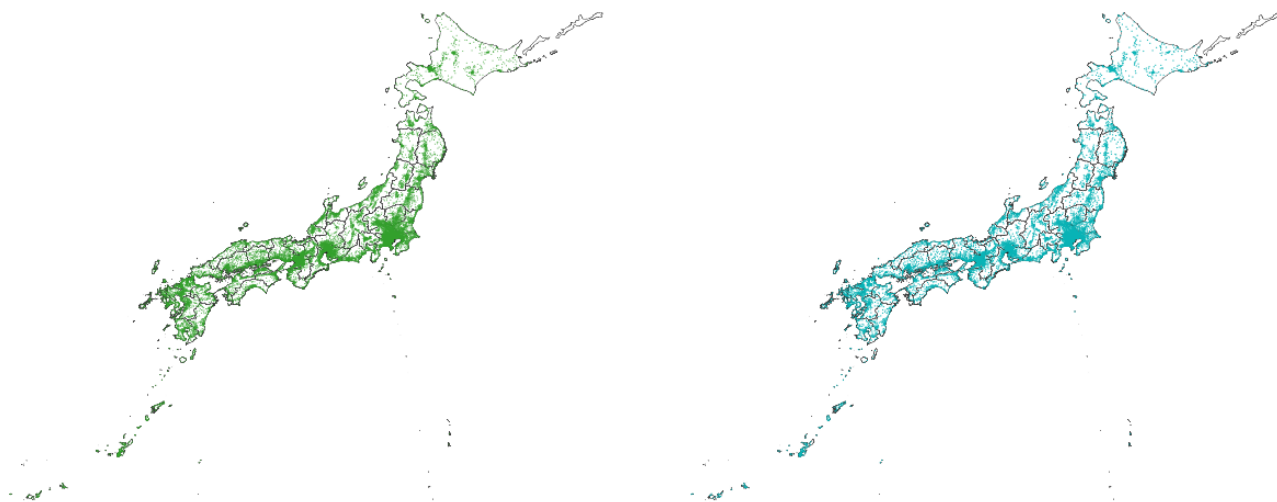
| | |
|-----------|--|
| Section A | <u>Additional Figures and Tables</u> |
| Section B | <u>Data Appendix</u> |

Appendix A: Additional figures and tables

Figure A.1—Location of schools

A. Primary schools (grade 6)

B. Secondary schools (grade 9)

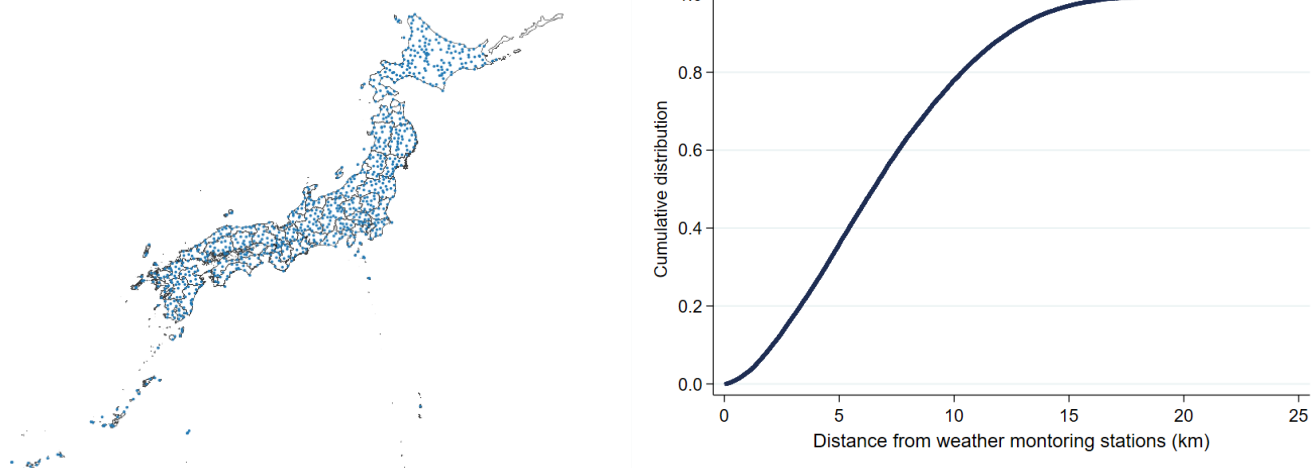


Notes: Panels A and B illustrate the locations of primary (grade 6) and secondary schools (grade 9) as of April 2019. There are 19,304 primary schools and 9,776 secondary schools.

Figure A.2—Weather stations

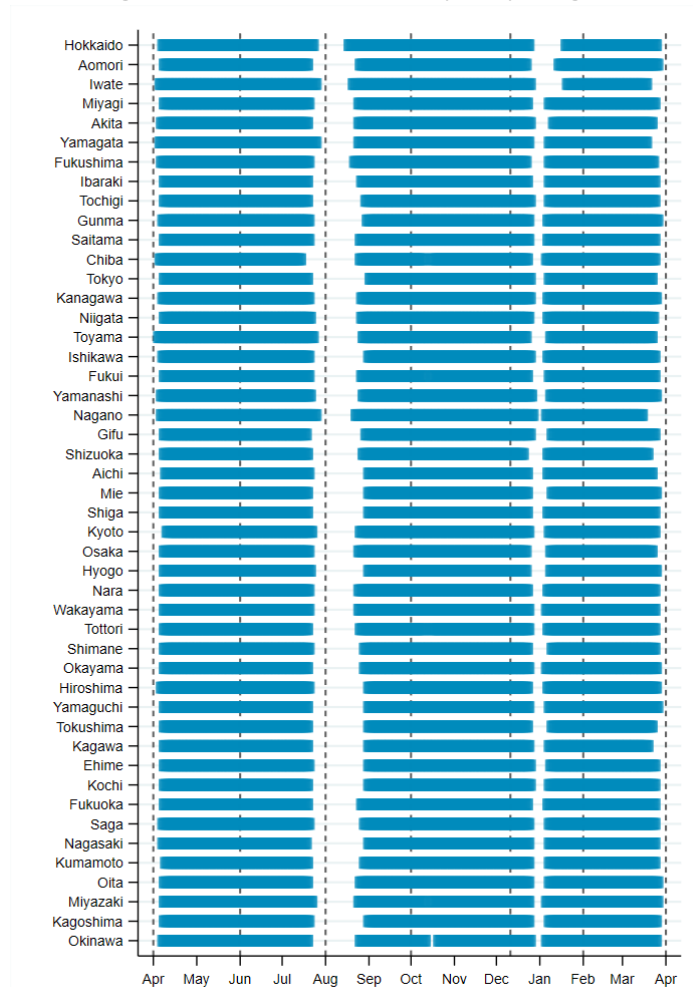
A. Location of weather stations

B. Distance to the weather stations



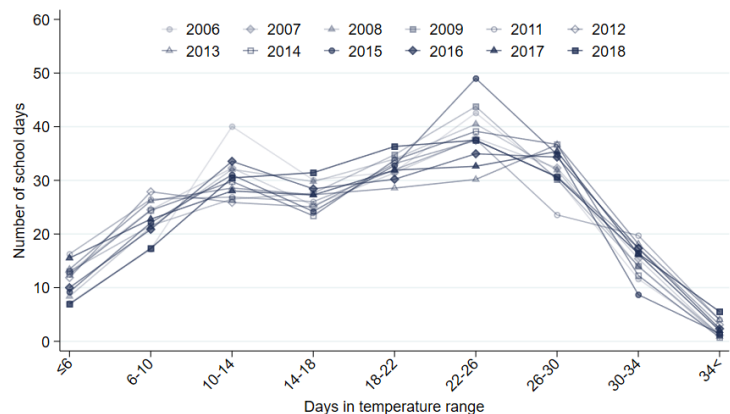
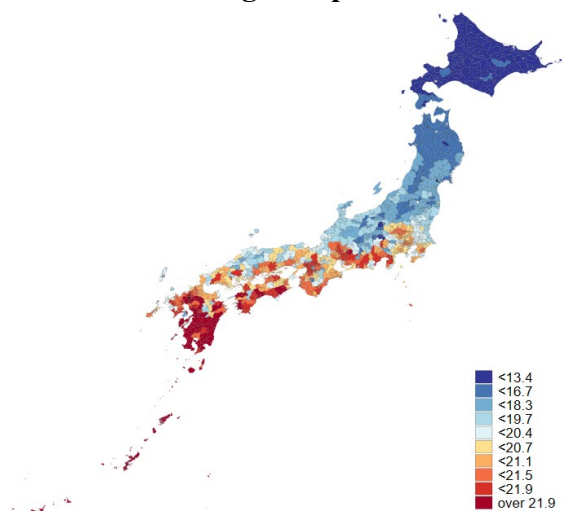
Notes: Panel A displays the locations of all 899 weather stations as of 2019. Panel B shows the cumulative distribution of the distances from schools to the nearest weather stations. The mean (median) distance from the weather stations is 6.95 (6.48) km.

Figure A.3—School days by region



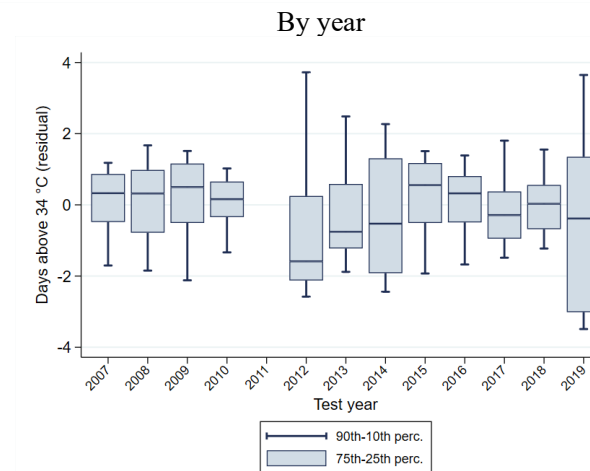
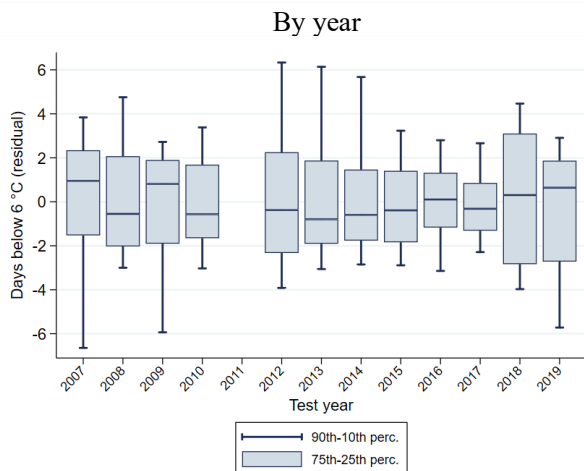
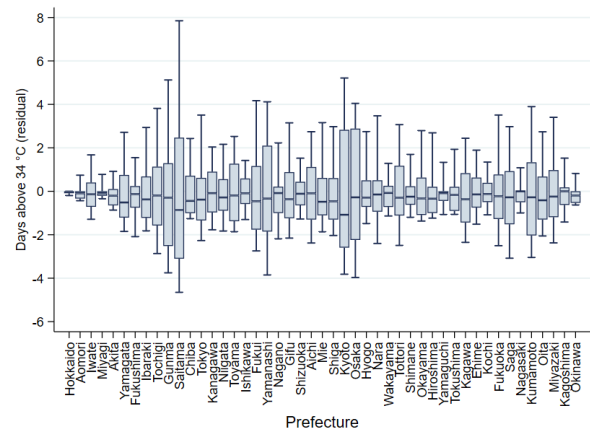
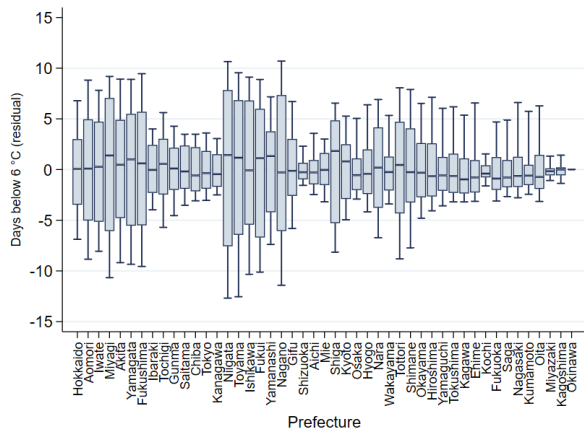
Notes: The figure displays the academic calendar of the prefectural capital in the school's prefecture for 2018. Japan is divided into 47 prefectures. The academic calendar mostly comprises three terms: spring, fall, and winter.

Figure A.4—Spatial and temporal variations in prior year temperature
A. Average temperature **B. Number of school days**



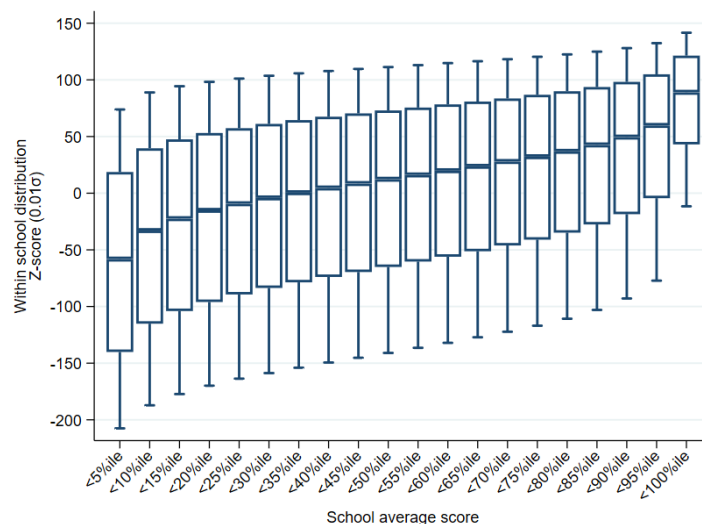
Notes: The figures illustrate the spatial variations in the mean daily maximum temperature in the year preceding the test year (panel A) and temporal variations in the number of school days within a given maximum temperature bin from last April to March of the test year, as experienced by students on school days (panel B).

Figure A.5—Identifying variations in prior year temperature
A. Number of days below 6°C
B. Number of days above 34°C
 By prefecture



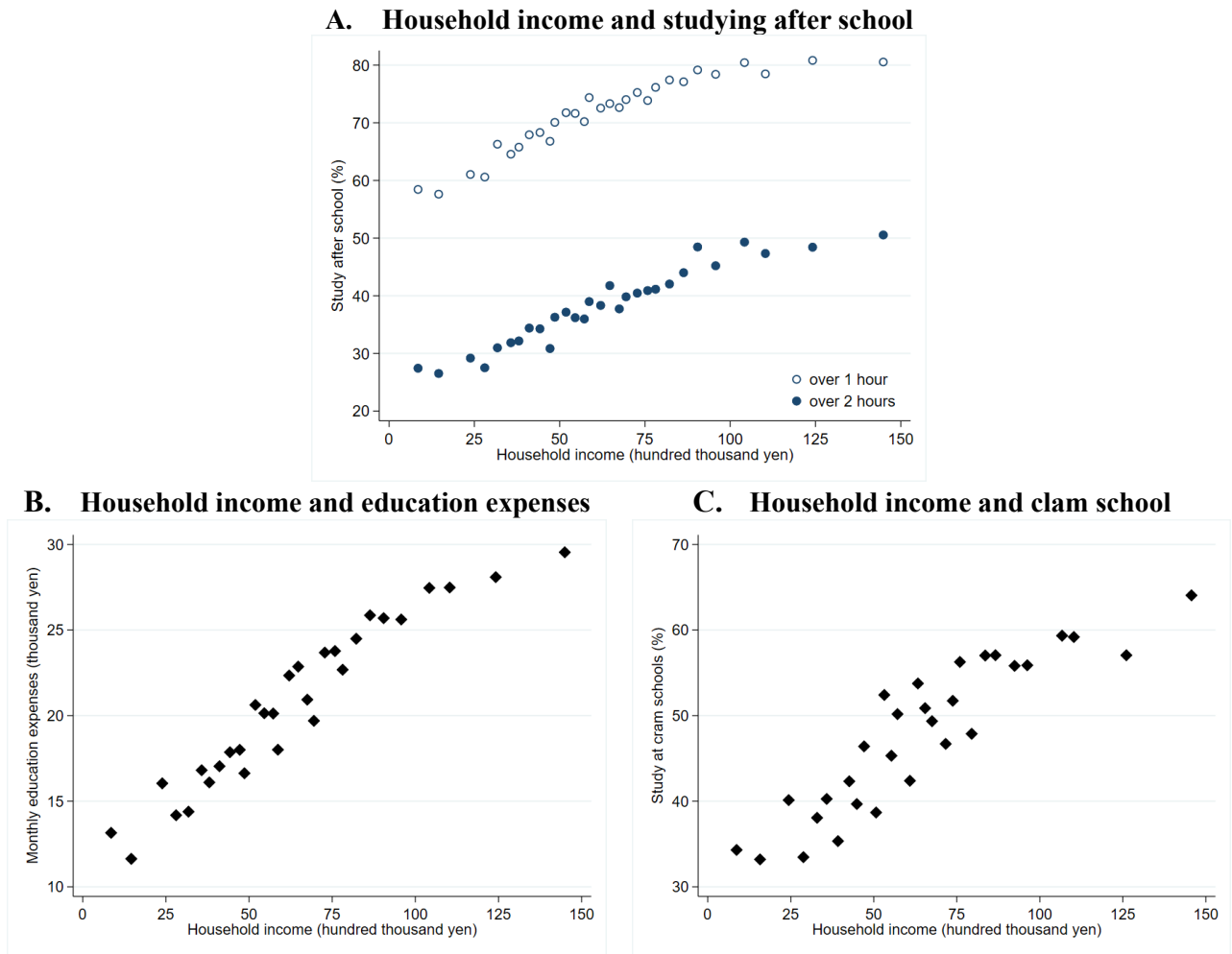
Notes: This figure illustrates the interquartile and interdecile ranges of the residual variation, net of school fixed effects, in the number of school days below 6°C in the year prior to the test date (panel A) and the number of school days above 34°C in the year prior to the test date (panel B), by prefecture and year. Japan has a total of 47 prefectures. The estimates are weighted by the number of students in each school.

Figure A.6—Within-school score distribution across school ranks



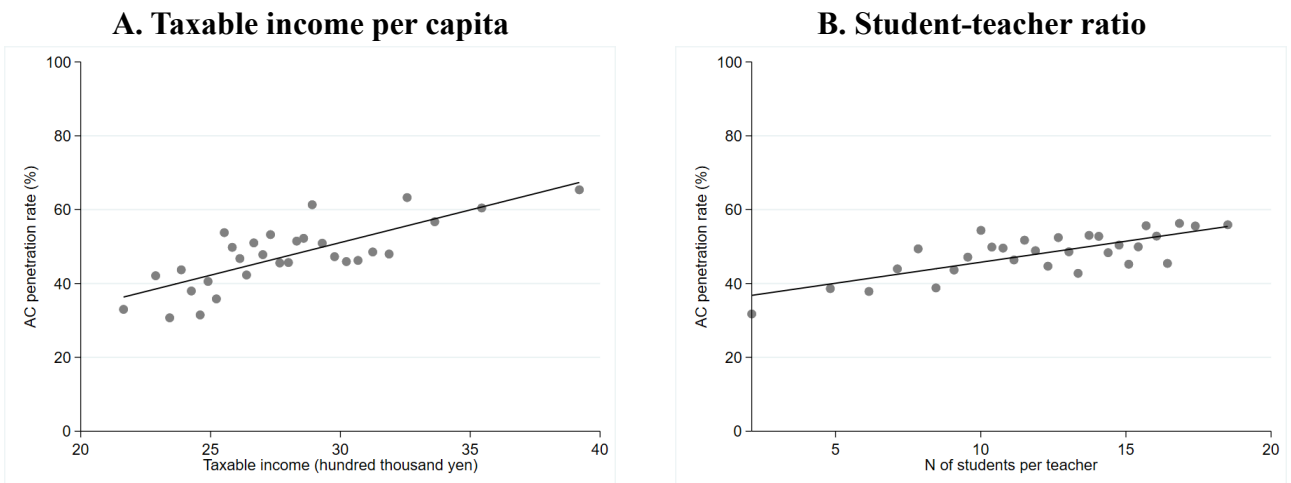
Notes: This figure illustrates the variations in *within*-school score distribution across school ranks based on the average school scores. Specifically, we group schools into ventiles based on their average scores each year and plot the average interquartile and interdecile ranges of the within-school score distribution for every ventile.

Figure A.7—Socioeconomic status and studying after school



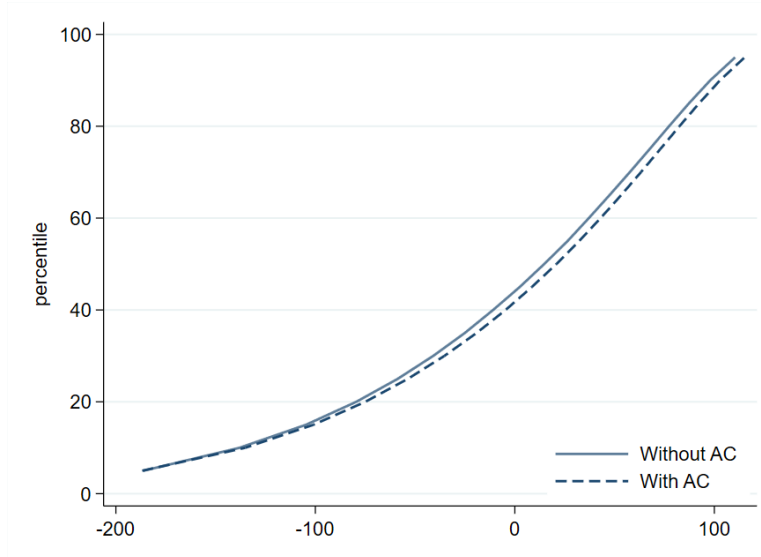
Notes: The data are obtained from parent surveys in the 2013 and 2017 NAAA, except for the fraction of students who study for more than 1 hour or 2 hours in panel A, which is obtained from student surveys in the 2013 and 2017 NAAA. The binscatter plot illustrates the relationship between students' socioeconomic status, as indicated by household income, and various study-related variables after school, net of school fixed effects. Specifically, it shows the proportion of students studying after school for more than 1 hour or 2 hours (panel A), monthly education expenses (panel B), and the proportion of students attending cram schools (panel C). Household income (panels A-C) is presented in hundreds of thousands of yen, while monthly education expenses (panel B) are presented in thousands of yen, with US\$1 being approximately equal to 100 yen. For both variables, we use the median of each household income/monthly education expense bin to transform them into continuous variables.

Figure A.8—Correlation with school AC penetration rates



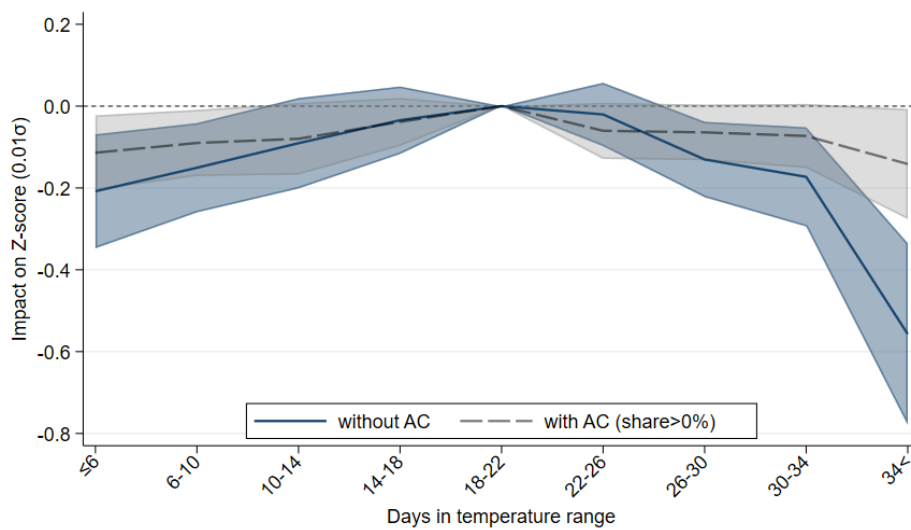
Notes: The binscatter plot illustrates the cross-sectional relationship between school AC penetration rates at the municipality level and taxable income per capita (panel A) as well as the student-teacher ratio (panel B) for 2018, after controlling for the average temperature between 2006 and 2018. Both taxable income per capita and the student-teacher ratio were averaged over the period from 2006 to 2018.

Figure A.9—Test score distributions between schools with and without AC in 2007



Notes: This figure plots the average test scores at 5-percentile increments in 2007 (the first year of our sample period), when AC penetration was only 10.2%, separately for schools with and without AC. These averages were obtained by first calculating test scores at each 5th percentile within schools, then averaging across schools within each group.

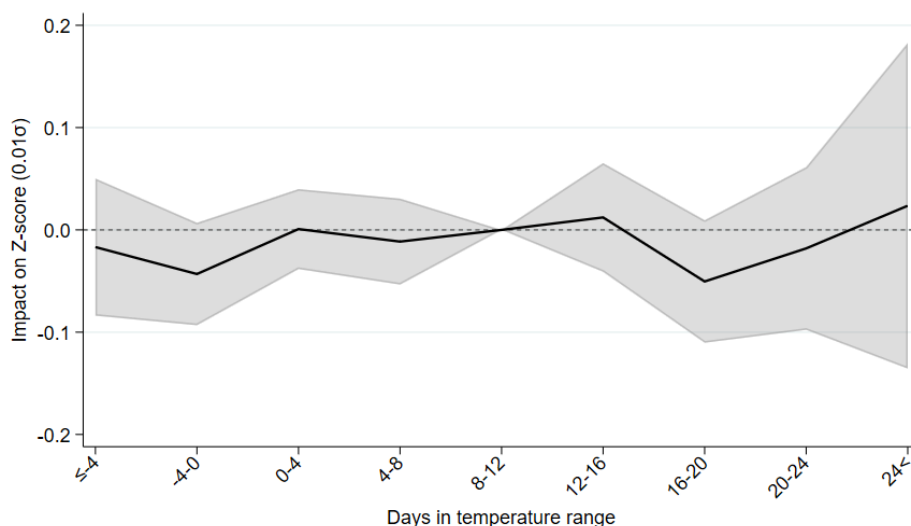
**Figure A.10—Cumulative heat/cold exposure and test performance
(school AC 0% vs. AC>0%)**



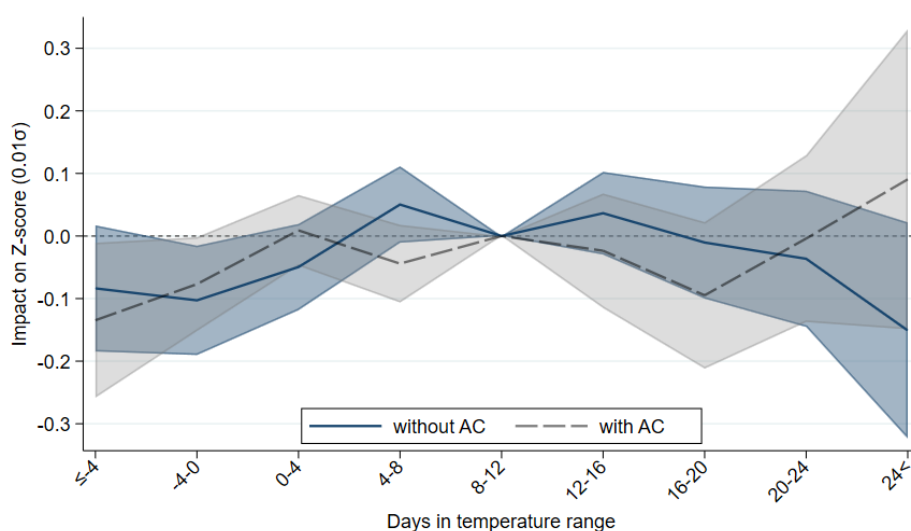
Notes: The figure displays β^k from an estimating Equation [1], separately for schools in municipalities with a positive share of AC and those in municipalities with 0% AC availability in 2018, along with the 95% confidence intervals. The omitted category is the temperature range between 18 and 22°C. Figure 4A shows the locations of schools for each school AC penetration category.

Figure A.11—Minimum temperature and test performance (average impacts)

A. All schools



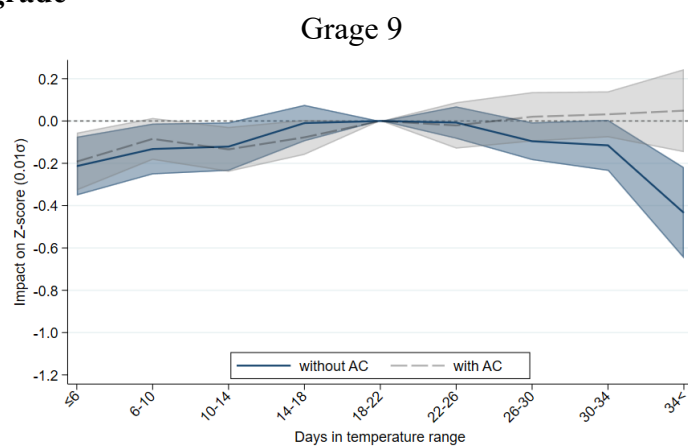
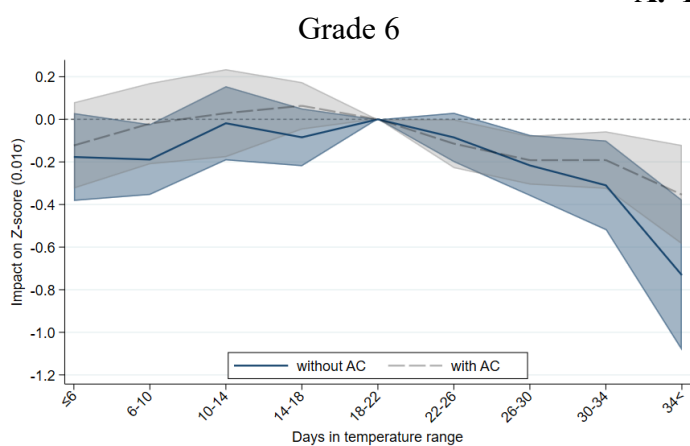
B. Schools with vs. without AC



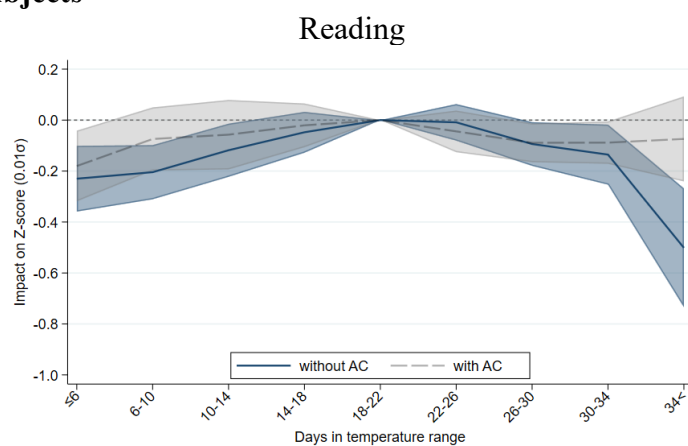
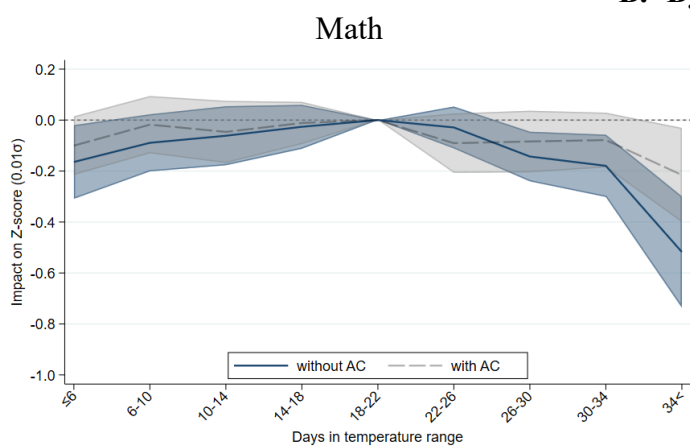
Notes: Panel A plots β^k from an estimating Equation [1], where the average z-score (measured in 0.01σ) is regressed on the number of school days within a given *minimum* temperature bin in the year prior to the test date, and panel B plots β^k separately for schools with and without AC in 2018, along with the 95% confidence intervals. The omitted category is the temperature range (8–12°C).

Figure A.12—Heterogeneity: Cumulative heat/cold exposure and test performance (with and without school AC)

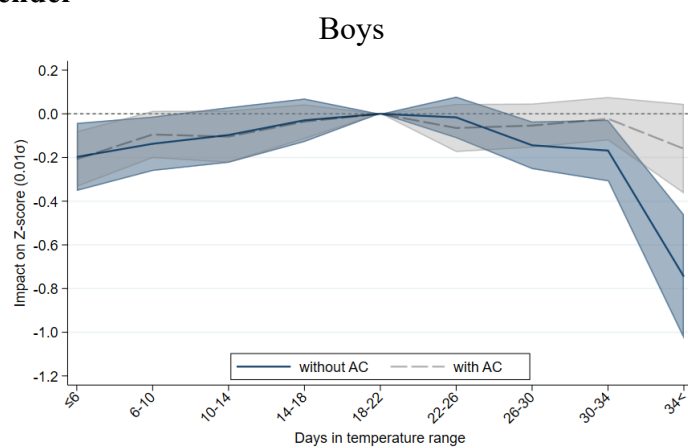
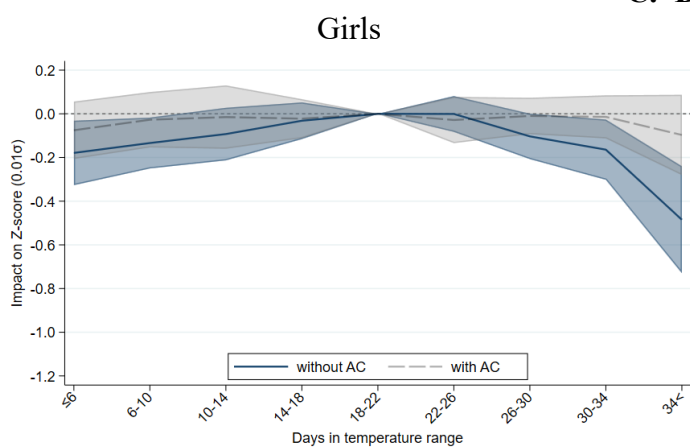
A. By grade



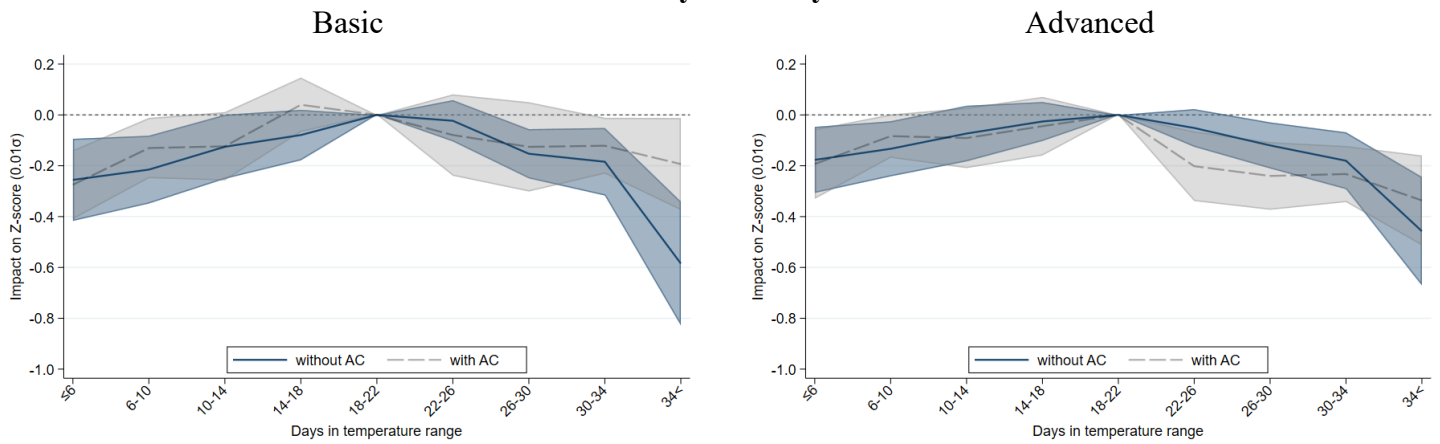
B. By subjects



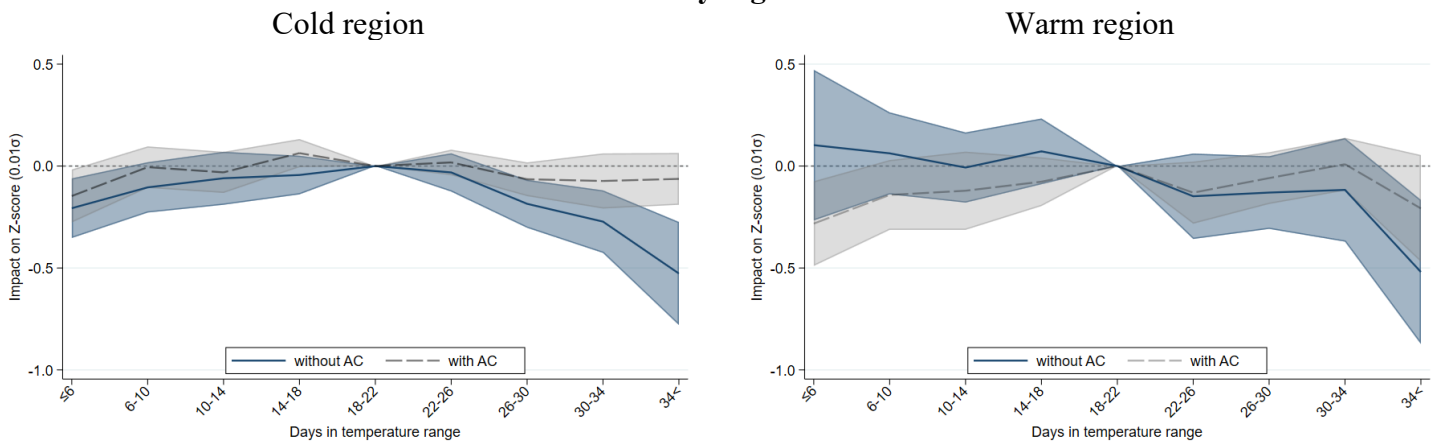
C. By gender



D. By difficulty



E. By region



Notes: The figures plot β^k from an estimating Equation [1], separately for schools with and without AC in 2018, along with the 95% confidence intervals. Figure 4A shows the locations of schools for each school AC penetration category. The omitted category is the temperature range (18–22°C). Panels A-D divide the sample by grade (grade 6 vs. grade 9), subject (math vs. reading), student gender (girls vs. boys), and the difficulty of the test questions (basic vs. advanced). Finally, panel E divides the sample into cool and warm regions by the national median temperature between 2006 and 2018.

Figure A.13—Examples of basic and advanced questions (math for grade 6)

| Basic | Advanced |
|--|--|
| <div style="border: 1px solid black; padding: 2px; width: 30px; margin: 0 auto;">1</div> <p>次の計算をしましょう。</p> <p>(1) $28 + 72$</p> <p>(2) 27×3.4</p> <p>(3) 9.3×0.8</p> <p>(4) $12 \div 0.6$</p> | <div style="border: 1px solid black; padding: 2px; width: 30px; margin: 0 auto;">1</div> <p>図アのような、たてが6 m、横が9 mの長方形の形をした花だんがあります。この中に、たてが3 m、横が5 mの長方形の の部分があります。</p> <div style="text-align: center; margin: 10px 0;"> <p>図ア</p> </div> <p>(1) の部分のまわりにロープをはります。 の部分のまわりにはるロープの長さは、どのような式で求められますか。</p> |

Notes: The examples are from mathematics for grade 6 in the NAEE.

Table A.1—Number of participating schools and students in NAAA

| Year | N of schools | | | N of students | | |
|-------|--------------|---------|---------|---------------|------------|------------|
| | Total | Grade 6 | Grade 9 | Total | Grade 6 | Grade 9 |
| 2007 | 31,899 | 21,523 | 10,376 | 2,203,309 | 1,115,808 | 1,087,501 |
| 2008 | 32,095 | 21,670 | 10,425 | 2,243,391 | 1,162,311 | 1,081,080 |
| 2009 | 31,835 | 21,498 | 10,337 | 2,264,473 | 1,153,059 | 1,111,414 |
| 2010 | 9,866 | 5,421 | 4,445 | 708,995 | 271,004 | 437,991 |
| 2011 | - | - | - | - | - | - |
| 2012 | 9,545 | 5,177 | 4,368 | 703,244 | 262,114 | 441,130 |
| 2013 | 30,560 | 20,468 | 10,092 | 2,207,777 | 1,124,018 | 1,083,759 |
| 2014 | 30,233 | 20,221 | 10,012 | 2,162,765 | 1,097,584 | 1,065,181 |
| 2015 | 29,962 | 20,030 | 9,932 | 2,136,316 | 1,076,832 | 1,059,484 |
| 2016 | 29,125 | 19,397 | 9,728 | 2,076,404 | 1,037,066 | 1,039,338 |
| 2017 | 29,174 | 19,375 | 9,799 | 2,047,892 | 1,018,505 | 1,029,387 |
| 2018 | 29,248 | 19,431 | 9,817 | 2,012,527 | 1,041,474 | 971,053 |
| 2019 | 28,989 | 19,252 | 9,737 | 2,025,844 | 1,046,722 | 979,122 |
| Total | 322,531 | 213,463 | 109,068 | 22,792,937 | 11,406,497 | 11,386,440 |

Notes: This table shows the number of schools and students participating in the National Assessment of Academic Ability (NAAA) each year. We exclude schools that are observed only once during the sample period, along with their corresponding students, and those without math and reading scores (0.24% of schools and 1.59% of students). The NAAA has been conducted annually across the nation by the Ministry of Education, Culture, Sports, Science, and Technology (MEXT) since 2007. Exceptions occurred in 2011, when the NAAA was entirely canceled because of the Great East Japan Earthquake, and in 2010 and 2012, when it was administered to a random subset of schools: approximately 25% of sixth graders and 40% of ninth graders.

Table A.2—Descriptive statistics

| Variable: | N of schools | Mean | Std. dev. | Min | Max | N of station | Available period |
|-------------------------------------|--------------|-------|-----------|-------|--------|--------------|------------------|
| Panel A. Student information | | | | | | | |
| Student survey: | | | | | | | |
| Female | 301,821 | 0.48 | 0.08 | 0 | 1 | - | 2007-2019 |
| Study time after school: >1 hour | 323,144 | 0.65 | 0.13 | 0 | 1 | - | 2007-2019 |
| Study time after school: >2 hours | 323,144 | 0.31 | 0.13 | 0 | 1 | - | 2007-2019 |
| Parent survey: | | | | | | | |
| Household income | 2,624 | 62.26 | 31.68 | 10 | 150 | - | 2013, 2017 |
| Father's educ: ≥University graduate | 2,779 | 0.31 | 0.46 | 0 | 1 | - | 2013, 2017 |
| Education expenses | 2,628 | 17.03 | 14.48 | 0 | 50 | - | 2013, 2017 |
| Attending a cram school | 1,952 | 0.33 | 0.47 | 0 | 1 | - | 2017 |
| Regional information: | | | | | | | |
| School AC | 322,962 | 0.60 | 0.46 | 0 | 1 | - | 2018 |
| Taxable income per capita | 323,153 | 32.29 | 5.86 | 18.89 | 126.67 | - | 2007-2019 |
| Student-teacher ratio | 321,263 | 15.63 | 3.06 | 0.09 | 25.05 | - | 2007-2019 |
| Home AC | 323,153 | 0.90 | 0.15 | 0.27 | 0.99 | - | 2014 |
| Panel B. Weather condition | | | | | | | |
| Number of school days | | | | | | | |
| 6°C≤ | 322,531 | 11.02 | 16.83 | 0 | 194 | 891 | 2007-2019 |
| 6-10°C | 322,531 | 22.59 | 8.37 | 0 | 53 | 891 | 2007-2019 |
| 10-14°C | 322,531 | 30.93 | 8.78 | 0 | 60 | 891 | 2007-2019 |
| 14-18°C | 322,531 | 27.20 | 5.80 | 0 | 58 | 891 | 2007-2019 |
| 18-22°C | 322,531 | 32.42 | 6.34 | 0 | 71 | 891 | 2007-2019 |
| 22-26°C | 322,531 | 38.37 | 8.51 | 0 | 88 | 891 | 2007-2019 |
| 26-30°C | 322,531 | 32.95 | 9.58 | 0 | 97 | 891 | 2007-2019 |
| 30-34°C | 322,531 | 14.87 | 8.90 | 0 | 79 | 891 | 2007-2019 |
| 34°C> | 322,531 | 2.25 | 2.95 | 0 | 20 | 891 | 2007-2019 |
| Mean precipitation (mm) | 322,531 | 4.53 | 1.37 | 0.82 | 21.48 | 1,165 | 2007-2019 |
| Mean wind speed (m/s) | 322,531 | 2.51 | 0.91 | 0.26 | 8.75 | 887 | 2007-2019 |
| Mean relative humidity | 322,531 | 68.44 | 4.92 | 58.39 | 82.94 | 153 | 2007-2019 |

Notes: Panel A provides descriptive statistics of student information aggregated at the school level. Gender information for grade 6 was not collected in 2015. Household income is presented in hundreds of thousands of yen, while monthly education expenses are shown in thousands of yen, with US\$1 being approximately equal to 100 yen. For both variables, we calculate the median household income and monthly education expense bin to convert them into continuous variables. For school and home AC, data from 2018 and 2014, respectively, are applied to all years. Panel B displays the descriptive statistics of the cumulative weather conditions from last April to March of the test year, as experienced by students on school days.

Table A.3—Comparison with previous studies of cumulative exposure to heat or cold on test scores

| Study | Country (period) | Exam type | Stakes | Grades | Representation | Exam Days | Effect size by one additional day |
|------------------------|---------------------|-----------------------|--------------|-------------------|---|---|--|
| Our Study | Japan ('07-'19) | Achievement test | Low | G6 and G9 | All students in public schools | 3 rd or 4 th Tuesday in April | <u>Reference: 18–22°C</u> Above 34°C ↓ 0.19% SD Below 6° ↓ 0.13% SD |
| Cho (2017) | Korea ('09-'13) | College entrance exam | High | G12 | Takers of university entrance exam | 2 nd Thursday in November | <u>Reference: 28–30°C</u> Above 34°C ↓ 0.42% SD (Math) ↓ 0.64% SD (English) |
| Park et al. (2020) | US ('01-'14) | PSAT | Intermediate | G10 or G11 | Takers of PSAT at least twice | 3 rd week of October | <u>Reference: 60–69°F (15.6-20.6°C)</u> Above 100°F (37.8°C) ↓ 0.07% SD Above 90°F (32.2°C) ↓ 0.05% SD |
| Park et al. (2021) | US ('09-'15) | State-specific exams | Intermediate | G3 to G8 | 12,000 US school districts | Spring (differ by state) | <u>Reference: 60–69°F (15.6-20.6°C)</u> Above 80°F (26.7°C) ↓ 0.10% SD (G3–G5) ↓ 0.03% SD (G6–G8) |
| Johnston et al. (2021) | Australia ('09-'18) | Achievement test | Low | G3, G5, G7 and G9 | All students in public schools in New South Wales | 2 nd week of May | <u>Reference: 65–75°F (18.3-23.9°C)</u> Below 60°F (15.6°C) ↓ 0.15% SD |

References:

- Cho, Hyunkuk. 2017. “Effect of Summer Heat on Test Scores: A Cohort Analysis.” *Journal of Environmental Economics and Management*, 83: 185–196.
- Park, R. Jisung, Joshua Goodman, Michael Hurwitz, and Jonathan Smith. 2020. “Heat and Learning.” *American Economic Journal: Economic Policy*, 12(2): 306–339.
- Park, R. Jisung, A. Patrick Behrer, and Joshua Goodman. 2021. “Learning is inhibited by heat exposure, both internationally and within the United States.” *Nature Human Behaviour*, 5: 19–27.
- Johnston, David W., Rachel Knott, Silvia Mendolia, and Peter Siminski. 2021. “Upside-Down Down-Under: Cold Temperatures Reduce Learning in Australia.” *Economics of Education Review*, 85: 102172.

Table A.4—Distributional impact of cumulative heat/cold exposure

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
|--------------|----------------------|----------------------|----------------------|--------------------|---------------------|------------------------------------|------------------------------------|------------------------------------|
| Outcomes: | 10 th | 25 th | 50 th | 75 th | 90 th | Resulting score gap | | |
| | percentile | percentile | percentile | percentile | percentile | 90 th -10 th | 90 th -50 th | 50 th -10 th |
| Days 6°C≤ | -0.260*** (0.053) | -0.185*** (0.052) | -0.117** (0.050) | -0.066 (0.042) | -0.028 (0.032) | 0.231*** (0.042) | 0.089*** (0.026) | 0.142*** (0.036) |
| Days 6-10°C | -0.177*** (0.047) | -0.144*** (0.046) | -0.105** (0.044) | -0.050 (0.037) | -0.008 (0.027) | 0.169*** (0.037) | 0.097*** (0.023) | 0.072** (0.032) |
| Days 10-14°C | -0.187*** (0.049) | -0.133** (0.052) | -0.074 (0.048) | -0.028 (0.037) | -0.007 (0.026) | 0.180*** (0.036) | 0.067*** (0.026) | 0.113*** (0.027) |
| Days 14-18°C | -0.079** (0.037) | -0.047 (0.036) | -0.033 (0.031) | -0.028 (0.027) | -0.018 (0.021) | 0.060* (0.036) | 0.015 (0.019) | 0.046* (0.027) |
| Days 22-26°C | -0.075* (0.040) | -0.062 (0.039) | -0.052 (0.035) | -0.036 (0.030) | -0.028 (0.022) | 0.047 (0.031) | 0.024 (0.019) | 0.022 (0.023) |
| Days 26-30°C | -0.108*** (0.041) | -0.089** (0.041) | -0.068* (0.035) | -0.041 (0.029) | -0.034* (0.021) | 0.074** (0.034) | 0.033 (0.021) | 0.040* (0.024) |
| Days 30-34°C | -0.124** (0.051) | -0.118** (0.050) | -0.095** (0.045) | -0.063* (0.035) | -0.032 (0.027) | 0.092** (0.043) | 0.063** (0.027) | 0.029 (0.029) |
| Days 34°C> | -0.303*** (0.079) | -0.231*** (0.080) | -0.209*** (0.075) | -0.120* (0.062) | -0.087** (0.044) | 0.216*** (0.062) | 0.123*** (0.043) | 0.093** (0.047) |
| R-squared | 0.648 | 0.687 | 0.691 | 0.647 | 0.573 | 0.531 | 0.532 | 0.359 |
| Observations | 322,531 | 322,531 | 322,531 | 322,531 | 322,531 | 322,531 | 322,531 | 322,531 |

Notes: The data are from the 2007–2019 NAAA and 2006–2018 AMeDAS. The unit of observation is the school-year. Columns (1)–(5) present the estimates from Equation [2], where the outcome is the z-scores at the 10th, 25th, 50th, 75th, and 90th percentiles within school (measured in 0.01σ), along with standard errors clustered at the weather station level in parentheses. Columns (6), (7), and (8) present the estimate of the score gap between the 90th and 10th percentiles, 90th and 50th percentiles, and 50th and 10th percentiles within the school, respectively, measured in 0.01σ . The estimates are weighted by the number of students in each school-year. The omitted category is the temperature range (18–22°C). Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table A.5—The impact of studying after school

| Outcomes: | (1) | | (2) | |
|--------------|--------------------------------|---------------------|---------------------------------|---------------------|
| | Average Z-score | | Average Z-score | |
| | × study more than 1 hour | | × study more than 2 hours | |
| Days 6°C≤ | -0.117*** (0.041) | 0.151*** (0.038) | -0.126*** (0.041) | -0.042 (0.049) |
| Days 6-10°C | -0.064* (0.036) | 0.376*** (0.047) | -0.095** (0.037) | 0.599*** (0.060) |
| Days 10-14°C | -0.080** (0.039) | 0.331*** (0.053) | -0.084** (0.039) | 0.316*** (0.058) |
| Days 14-18°C | -0.052** (0.026) | 0.415*** (0.063) | -0.047* (0.025) | 0.188** (0.080) |
| Days 22-26°C | -0.053 (0.032) | 0.039 (0.040) | -0.047 (0.031) | 0.007 (0.045) |
| Days 26-30°C | -0.075** (0.030) | 0.288*** (0.042) | -0.077** (0.030) | 0.212*** (0.051) |
| Days 30-34°C | -0.094** (0.037) | 0.248*** (0.069) | -0.101*** (0.038) | 0.149** (0.070) |
| Days 34°C> | -0.186*** (0.060) | 0.403** (0.170) | -0.206*** (0.062) | 0.150 (0.201) |
| R-squared | 0.748 | | 0.737 | |
| Observations | 322,523 | | 322,523 | |

Notes: The data are from the 2007–2019 NAAA and 2006–2018 AMeDAS. The unit of observation is the school-year. The dependent variable is the average test score at the school-year level, measured in 0.01σ . Estimates from a variant of Equation [1], which additionally includes the interaction between the fraction of students studying after school for more than 1 hour (column 1) and for 2 hours (column 2), with the number of days in each temperature bin during school days from the previous year, are reported along with standard errors clustered at the weather station level in parentheses. Note that both fractions of students studying after school for more than 1 (column 1) and 2 hours (column 2) are demeaned by the average for the NAAA between 2007 and 2019. The interaction terms in columns (1) and (2) reflect the offsetting effect of studying after school, as the fraction of students studying for more than 1 or 2 hours after school increased from 0% to 100%. The estimates are weighted by the number of students in each school-year. The omitted category is the temperature range (18–22°C). Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table A.6—The average impact of school AC

| Outcomes: | (1) | | (2) | | (3) | |
|------------------------------------|----------------------|---------------------|----------------------|---------------------|----------------------|---------------------|
| | Average Z-score | | Average Z-score | | Average Z-score | |
| | × school AC | | × school AC | | × school AC | |
| Days 6°C≤ | -0.216*** (0.069) | 0.068 (0.090) | -0.235*** (0.077) | 0.061 (0.097) | -0.080 (0.061) | -0.109 (0.082) |
| Days 6-10°C | -0.158*** (0.053) | 0.114 (0.076) | -0.111** (0.056) | -0.007 (0.074) | -0.063 (0.052) | 0.010 (0.074) |
| Days 10-14°C | -0.098* (0.054) | 0.046 (0.084) | -0.082 (0.059) | -0.059 (0.073) | -0.041 (0.047) | -0.044 (0.082) |
| Days 14-18°C | -0.040 (0.041) | 0.031 (0.057) | -0.020 (0.050) | -0.040 (0.064) | 0.043 (0.043) | -0.080 (0.053) |
| Days 22-26°C | -0.024 (0.039) | -0.043 (0.063) | -0.042 (0.043) | 0.019 (0.057) | -0.115*** (0.042) | 0.069 (0.081) |
| Days 26-30°C | -0.134*** (0.046) | 0.050 (0.066) | -0.140*** (0.050) | 0.125* (0.067) | -0.243*** (0.048) | 0.187*** (0.069) |
| Days 30-34°C | -0.177*** (0.061) | 0.097 (0.073) | -0.209*** (0.071) | 0.202* (0.113) | -0.298*** (0.065) | 0.310*** (0.084) |
| Days 34°C> | -0.562*** (0.112) | 0.413*** (0.145) | -0.565*** (0.121) | 0.476*** (0.141) | -0.623*** (0.107) | 0.501*** (0.153) |
| Interaction with taxable income | | | X | | | |
| student-teacher ratio | | | X | | | |
| home AC share | | | | | X | |
| R-squared | 0.751 | | 0.752 | | 0.751 | |
| Observations | 190,210 | | 188,911 | | 190,210 | |

Notes: The data are from the 2007–2019 NAAA and 2006–2018 AMeDAS. The unit of observation is the school-year. The dependent variable is the average test score at the school-year level, measured in 0.01σ . The estimates from Equation [3] are reported, along with standard errors clustered at the weather station level in parentheses. School AC is a dummy variable that equals one if an air conditioner was available at the school in 2018. Figure 4A shows the locations of the schools within each AC penetration category. Column (1) replicates column (1) of Table 1 for reference. Column (2) adds to column (1) the interaction of municipality-level taxable income per capita and the student-teacher ratio in 2018 with the number of school days within a given maximum temperature bin in the year prior to the test date. Column (3) adds to column (1) the interaction of prefecture-level home AC shares in 2014 with the number of school days within a given maximum temperature bin in the year before the test date. The estimates are weighted by the number of students in each school-year. The omitted category is the temperature range (18–22°C).

Table A.7—Robustness of the impact of heat

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
|-------------------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| Outcomes: | Average Z-score | | | | | | | |
| Days above 34°C | -0.562*** (0.112) | -0.568*** (0.113) | -0.557*** (0.111) | -0.278*** (0.105) | -0.503*** (0.113) | -0.493*** (0.105) | -0.482*** (0.130) | -0.551*** (0.123) |
| Days above 34°C × school AC | 0.413*** (0.145) | 0.425*** (0.145) | 0.421*** (0.139) | 0.401*** (0.126) | 0.363** (0.146) | 0.386*** (0.130) | 0.409** (0.164) | 0.440*** (0.150) |
| R-squared | 0.751 | 0.751 | 0.751 | 0.772 | 0.751 | 0.751 | 0.751 | 0.759 |
| Observations | 190,210 | 190,210 | 190,210 | 145,769 | 190,210 | 190,210 | 190,210 | 141,733 |
| Sample period | Full | Full | Full | 2009-2019 | Full | Full | Full | Full |
| Temperature (test day) | | X | | | | | | |
| Weather (test day) | | | X | | | | | |
| Pollution (test day) | | | | X | | | | |
| Weather (cumulative) | | | | | X | | | |
| Temperature (school holidays) | | | | | | X | | |
| Temperature (weekend) | | | | | | | X | |
| Stations within 10 km | | | | | | | | X |

Notes: The data are from the 2007–2019 NAAA and 2006–2018 AMeDAS. The unit of observation is the school-year. The dependent variable is the average test score at the school-year level, measured in 0.01σ . The estimates are obtained from Equation [3], along with the standard errors clustered at the weather station level in parentheses. The estimates for the number of school days above 34°C and their interaction with the school AC dummy are reported, while the estimates for days in other temperature ranges are omitted for expositional purposes. School AC is a dummy variable that takes the value of one if an AC was available at the school in 2018. Figure 4A shows the locations of the schools within each AC penetration category. The estimates are weighted by the number of students in each school year. The omitted category is the temperature range (18–22°C). The full sample period is from 2007 to 2019. Column (1) presents the baseline estimate without any controls other than school and year fixed effects, as reported in column (1) of Table 1. Column (2) adds the test-day temperature and column (3) includes additional test-day weather conditions (precipitation, wind speed, and relative humidity). Column (4) includes test-day air pollution (SO₂, NO, NO₂, CO, OX, and PM₁₀) for the 2009–2019 period, as pollution data are only available for this period. Column (5) includes other cumulative weather conditions (precipitation, wind speed, and relative humidity). Columns (6) and (7) control the number of days during school break days and weekends, respectively, within a given maximum temperature bin from the year prior to the test date. Finally, column (8) restricts the sample to schools located within 10 km of the weather stations. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table A.8—Heterogeneous impacts of heat

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
|-----------------------------|------------------------|----------------------|------------------------|----------------------|------------------------|----------------------|------------------------|----------------------|------------------------|----------------------|
| Outcomes: | Average Z-score | | Average Z-score | | Average Z-score | | Average Z-score | | Average Z-score | |
| | <i>By grade</i> | | <i>By subject</i> | | <i>By gender</i> | | <i>By difficulty</i> | | <i>By region</i> | |
| | 6th | 9th | Math | Reading | Girls | Boys | Basic | Advanced | Cool | Warm |
| Days above 34°C | -0.731*** (0.180) | -0.443*** (0.109) | -0.523*** (0.111) | -0.504*** (0.118) | -0.467*** (0.123) | -0.755*** (0.145) | -0.588*** (0.123) | -0.465*** (0.108) | -0.530*** (0.128) | -0.520*** (0.178) |
| Days above 34°C × school AC | 0.378* (0.210) | 0.495*** (0.148) | 0.309** (0.145) | 0.431*** (0.142) | 0.356** (0.153) | 0.581*** (0.176) | 0.395*** (0.153) | 0.132 (0.141) | 0.470*** (0.142) | 0.312 (0.220) |
| R-squared | 0.682 | 0.807 | 0.742 | 0.712 | 0.676 | 0.679 | 0.741 | 0.753 | 0.721 | 0.778 |
| Observations | 115,312 | 74,898 | 190,210 | 190,210 | 176,297 | 176,169 | 173,005 | 173,005 | 105,672 | 84,538 |

Notes: The data are from the 2007–2019 NAAA and 2006–2018 AMeDAS. The unit of observation is the school-year. The dependent variable is the average test score at the school-year level, measured in 0.01σ . The estimates are obtained from Equation [3], along with the standard errors clustered at the weather station level in parentheses. The estimates for the number of school days above 34°C and their interaction with the school AC dummy are reported, while the estimates for days in other temperature ranges are omitted for expositional purposes. School AC is a dummy variable that takes the value of one if an AC was available at the school in 2018. Figure 4A shows the locations of the schools within each AC penetration category. The estimates are weighted by the number of students in each school-year. The omitted category is the temperature range (18–22°C). Columns (1) and (2) present the estimates by grade (grade 6 vs. grade 9). Columns (3) and (4) show estimates by subject area (math vs. reading). Columns (5) and (6) show estimates by student gender (girls vs. boys). Columns (6) and (7) present the estimates based on the difficulty of the test questions (basic vs. advanced). Finally, Columns (9) and (10) divide the sample into cool and warm regions based on the national median of the average temperature from 2006 to 2018. Note that the number of observations is at the school-year level; therefore, we observe the average test score of each school-year for each subject, gender, and question difficulty, while we observe only one test score for each grade and each region, as they are mutually exclusive. Thus, the sum of the observations in columns (1) and (2) and the sum of the observations in columns (9) and (10) is 190,210, which is equal to the number of school-years in columns (3) and (4). The slightly smaller observations for columns (5) and (6), compared with columns (3) and (4), are because gender information was not collected for grade 6 in 2015. Similarly, the slightly smaller observations in columns (7) and (8) compared to those in columns (3) and (4) are due to the absence of such a distinction in 2019. Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table A.9—Distributional impact of school AC (full)

| | (1) | | (2) | | (3) | | (4) | | (5) | |
|------------------|---|---------------------|---|---------------------|---|---------------------|---|-------------------|---|-------------------|
| Outcomes: | 10th percentile score | | 25th percentile score | | 50th percentile score | | 75th percentile score | | 90th percentile score | |
| | × school AC | | × school AC | | × school AC | | × school AC | | × school AC | |
| Days 6°C≤ | -0.320*** (0.101) | 0.040 (0.124) | -0.293*** (0.092) | 0.112 (0.113) | -0.257*** (0.080) | 0.143 (0.107) | -0.152** (0.062) | 0.044 (0.090) | -0.099** (0.047) | 0.050 (0.076) |
| Days 6-10°C | -0.207** (0.083) | 0.088 (0.103) | -0.187*** (0.072) | 0.098 (0.097) | -0.206*** (0.061) | 0.166* (0.091) | -0.124** (0.051) | 0.115 (0.076) | -0.075* (0.040) | 0.104* (0.059) |
| Days 10-14°C | -0.196** (0.082) | 0.043 (0.108) | -0.113 (0.074) | 0.001 (0.109) | -0.117* (0.065) | 0.081 (0.102) | -0.059 (0.049) | 0.064 (0.077) | -0.047 (0.035) | 0.067 (0.053) |
| Days 14-18°C | -0.027 (0.062) | -0.039 (0.083) | -0.009 (0.057) | -0.017 (0.080) | -0.047 (0.045) | 0.050 (0.067) | -0.053 (0.037) | 0.064 (0.053) | -0.046 (0.030) | 0.067 (0.042) |
| Days 22-26°C | -0.034 (0.057) | -0.070 (0.083) | -0.044 (0.054) | -0.049 (0.082) | -0.045 (0.049) | -0.019 (0.075) | -0.013 (0.037) | -0.036 (0.061) | -0.000 (0.026) | -0.042 (0.043) |
| Days 26-30°C | -0.218*** (0.063) | 0.075 (0.087) | -0.187*** (0.061) | 0.056 (0.089) | -0.150*** (0.056) | 0.077 (0.079) | -0.083* (0.047) | 0.041 (0.065) | -0.030 (0.035) | -0.024 (0.046) |
| Days 30-34°C | -0.271*** (0.085) | 0.136 (0.102) | -0.274*** (0.081) | 0.149 (0.097) | -0.187** (0.074) | 0.108 (0.090) | -0.110* (0.062) | 0.057 (0.074) | -0.065 (0.049) | 0.020 (0.058) |
| Days 34°C> | -0.932*** (0.176) | 0.690*** (0.201) | -0.813*** (0.151) | 0.624*** (0.186) | -0.610*** (0.133) | 0.461*** (0.171) | -0.341*** (0.113) | 0.262* (0.149) | -0.223*** (0.085) | 0.139 (0.108) |
| R-squared | 0.669 | | 0.708 | | 0.714 | | 0.673 | | 0.603 | |
| Observations | 190,210 | | 190,210 | | 190,210 | | 190,210 | | 190,210 | |

Notes: The data are from the 2007–2019 NAAA and 2006–2018 AMeDAS. The unit of observation is the school-year. Columns (1)–(5) present the estimates from the variant of Equation [3], where the outcomes are z-scores at the 10th, 25th, 50th, 75th, and 90th percentiles within schools (measured in 0.01σ), along with standard errors clustered at the weather station level in parentheses. School AC is a dummy variable that equals one if an air conditioner is available at the school in 2018. Figure 4A shows the locations of the schools within each AC penetration category. The estimates are weighted by the number of students in each school-year. The omitted category is the temperature range (18–22°C). Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table A.10—Robustness: The impact of school AC on academic inequality

| Outcomes: | (1) | | (2) | | (3) | |
|------------------------------------|--|----------------------|--|---------------------|--|----------------------|
| | 90 th -10 th score gap | | 90 th -10 th score gap | | 90 th -10 th score gap | |
| | × school AC | | × school AC | | × school AC | |
| Days 6°C≤ | 0.222** (0.086) | 0.011 (0.110) | 0.225*** (0.086) | 0.025 (0.110) | 0.154 (0.094) | 0.107 (0.123) |
| Days 6-10°C | 0.132* (0.078) | 0.016 (0.089) | 0.121 (0.077) | -0.015 (0.089) | 0.052 (0.086) | 0.122 (0.102) |
| Days 10-14°C | 0.149** (0.074) | 0.025 (0.088) | 0.134* (0.074) | 0.065 (0.088) | 0.115 (0.079) | 0.081 (0.099) |
| Days 14-18°C | -0.019 (0.052) | 0.106 (0.078) | -0.054 (0.056) | 0.130* (0.078) | -0.051 (0.063) | 0.146 (0.089) |
| Days 22-26°C | 0.034 (0.050) | 0.028 (0.068) | 0.029 (0.053) | 0.037 (0.070) | 0.063 (0.057) | -0.008 (0.078) |
| Days 26-30°C | 0.188*** (0.053) | -0.099 (0.075) | 0.185*** (0.054) | -0.137* (0.072) | 0.233*** (0.058) | -0.154* (0.081) |
| Days 30-34°C | 0.207*** (0.076) | -0.116 (0.094) | 0.213*** (0.078) | -0.143 (0.098) | 0.257*** (0.082) | -0.213** (0.107) |
| Days 34°C> | 0.709*** (0.170) | -0.551*** (0.184) | 0.629*** (0.172) | -0.402** (0.191) | 0.714*** (0.169) | -0.584*** (0.199) |
| Interaction with taxable income | | | X | | | |
| student-teacher ratio | | | X | | | |
| home AC share | | | | | X | |
| R-squared | 0.552 | | 0.553 | | 0.552 | |
| Observations | 190,210 | | 188,911 | | 190,210 | |

Notes: The data are from the 2007–2019 NAAA and 2006–2018 AMeDAS. The unit of observation is the school-year. Columns (1)–(3) present the estimates from the variant of Equation [2], which additionally includes the interaction of the number of school days within a given maximum temperature bin in the year prior to the test date and the school AC dummy, along with standard errors clustered at the weather station level in parentheses. School AC is a dummy variable that equals one if an air conditioner is available at the school in 2018. Figure 4A shows the locations of the schools within each AC penetration category. The outcome is the gap between the 90th and 10th percentile scores within the school, measured at 0.01σ . Column (1) replicates the estimates in column (3) of Table 2. The estimates are weighted by the number of students in each school-year. The omitted category is the temperature range (18–22°C). Significance levels: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Appendix B: Data Appendix

| Data | Source |
|--|--|
| National Assessment of Academic Ability (NAAA) | <p>Years: 2007–2019</p> <p>Data description: Reading and math scores for grades 6 and 9, after-school study participation (student survey), and students' socioeconomic status (parent survey, conducted only in 2013 and 2017)</p> <p>Source: The National Institute for Educational Policy Research https://www.nier.go.jp/kaihatsu/zenkokugakuryoku.html</p> |
| Weather | <p>Years: 2006–2019</p> <p>Data description: daily temperature (maximum, average, minimum)</p> <p>Source: Japan Automated Meteorological Data Acquisition System (AMeDAS) operated by the Japan Meteorological Agency (JMA) https://www.data.jma.go.jp/obd/stats/etrn/</p> |
| Pollution | <p>Years: 2009 April–2019 March</p> <p>Data description: hourly SO₂, NO, NO₂, CO, OX, PM₁₀</p> <p>Source: National Institute for Environmental Studies https://tenbou.nies.go.jp/download/</p> |
| Taxable income | <p>Years: 2006–2018</p> <p>Data description: taxable income per capita at the municipality level</p> <p>Source: Survey on Municipal Taxation Status (Shichōsonzei kazeijōkyō tou no shirabe) https://www.soumu.go.jp/main_sosiki/jichi_zeisei/czaisei/czaisei_seido/ichiran09.html</p> |
| Student-teacher ratio | <p>Years: 2006–2018</p> <p>Data description: student-teacher ratio at municipality level</p> <p>Source: School Basic Survey https://www.mext.go.jp/b_menu/toukei/chousa01/kihon/1267995.htm</p> |
| School AC penetration rate | <p>Year: 2018</p> <p>Data description: school AC penetration rate for public primary and secondary schools at the municipality level</p> <p>Source: Survey of Air Conditioning Installation Status in Public School Facilities https://www.mext.go.jp/a_menu/shotou/zyosei/mext_01278.html</p> |
| Home AC share | <p>Year: 2014</p> <p>Data description: home AC share at the prefecture level</p> <p>National Survey of Family Income and Expenditure https://www.stat.go.jp/data/zensho/2014/index.html</p> |