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Rural and Urban Poverty Estimates for Developing Countries: Methodologies

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Abstract

This paper is to set out the backgrounds for the construction of new rural and urban poverty and inequality estimates using the World Bank Living Standard Measurement Survey (LSMS) data of developing countries with focus on methodological details as well as on their advantages or disadvantages. First, we have reviewed recent regional estimates based on the US\$1.25 per day poverty line as well as those based on Multidimensional Poverty Index (MPI) for both rural and urban areas. It has been found that the level of poverty is much higher in rural areas than in urban areas across different regions regardless of the definitions of poverty. Second, we have summarised estimates of poverty and inequality for Tanzania and Uganda based on recent panel data constructed by LSMS.

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Rural and Urban Poverty Estimates for Developing Countries: Methodologies

1. Backgrounds

This paper is to set out methodological backgrounds for the construction of new rural and urban poverty estimates for over 40 countries¹ mainly drawing upon the methodology used by the World Bank. The Global Monitoring Report 2013 (World Bank, 2013) disaggregates poverty estimates based on the international poverty line (US\$1.25 a day) into rural and urban poverty as in Table 1. It is not clear exactly which countries World Bank (2013) used to derive regional estimates of rural and urban poverty rates. World Bank (2013, p.18) noted "Data and insights have been drawn from countries in Asia, Sub-Saharan Africa, and Latin America, including Bangladesh, China, India, Indonesia, Kyrgyzstan, Thailand, Ghana, Kenya, South Africa, Tanzania, Uganda, Zambia, Brazil, and Ecuador", implying that only 14 countries were used to derive these estimates. However, as World Bank (2013) does not provide details of how regional aggregates are calculated, or report country-level estimates of rural and urban poverty, it is difficult to assess the extent to which these estimates are reliable or accurate. It should be noted that regional estimates in Table 1 are not likely to represent all the countries in each region. Our research project aims to derive rural and urban poverty estimates using LSMS data.²

¹ The exact number of countries for which decomposition of poverty is possible will be confirmed shortly by the research team. Our earlier study (Imai et al., 2014) presented recent estimates of rural and urban poverty based on the secondary study for 31 countries (see Appendix 1). The data are available for most of these countries, though for a few countries LSMS data are unavailable. For instance, in the case of India, National Sample Survey Data were used, but this is much more time consuming to process than LSMS data.

 $^{^2}$ The research team is now in the middle of processing LSMS data for the period 2000-2010 for a number of countries for which rural and urban poverty rates can be calculated. This is a highly labour-intensive work as it involves a large amount of data generation and management based on raw household data.

| | 19 | 90 | 19 | 96 | 20 | 02 | 20 | 08 | - |)15 liction) | | 015 0G1) |
|--|-------|-------|-------|-------|-------|-------|-------|-------|-------|-----------------|-------|-------------|
| | Rural | Urban | rural | Urban |
| East Asia and Pacific | 67.5 | 24.4 | 45.9 | 13 | 39.2 | 6.9 | 20.4 | 4.3 | 2.1 | 0.0 | 33.8 | 12.2 |
| Europe and Central Asia Latin America and the | 2.2 | 0.9 | 6.3 | 2.8 | 4.4 | 1.1 | 1.2 | 0.2 | 0.8 | 0.0 | 1.1 | 0.5 |
| Caribbean Middle East and North | 21 | 7.4 | 20.3 | 6.3 | 20.3 | 8.3 | 13.2 | 3.1 | 10.2 | 1.4 | 10.5 | 3.7 |
| Africa | 9.1 | 1.9 | 5.6 | 0.9 | 7.5 | 1.2 | 4.1 | 0.8 | 2.2 | 0.4 | 4.6 | 1.0 |
| South Asia Sub-Saharan Africa | 50.5 | 40.1 | 46.1 | 35.2 | 45.1 | 35.2 | 38 | 29.7 | 33.1 | 25.7 | 25.3 | 20.1 |
| (SSA) | 55 | 41.5 | 56.8 | 40.6 | 52.3 | 41.4 | 47.1 | 33.6 | 44.0 | 30.5 | 27.5 | 20.8 |
| Total | 52.5 | 20.5 | 43 | 17 | 39.5 | 15.1 | 29.4 | 11.6 | 20.4 | 8.1 | 26.3 | 10.3 |

Table 1. Poverty rates in urban and rural areas: Share of the population below \$1.25 a day.

Source: The figures for 1990-2008 are based on World Bank (2013, p.87). '2015 (prediction)' is based on the simple linear prediction using the estimates in 1990 and 2015, while '2015 (MDG1)' shows the target, 50% of poverty headcount ratios in 1990. In the column titled as '2015 (prediction)', the bold numbers show the cases where MDG1 will not be achieved while those in Italics the cases where MDG1 is expected to be met.

Given the limitations, a few points are noted about the aggregate and regional estimates of rural and urban poverty in Table 1. With regard to the aggregate estimates (in the last row of Table 1), a steady reduction in poverty has been achieved in both rural and urban areas. In 18 years between 1990 and 2008, rural poverty rate reduced from 52.5% to 29.4%, while urban poverty rate changed from 20.5% to 11.6%. In both cases, MDG1 is expected to be met.

Table 1 also indicates that MGD1 of halving poverty is unlikely to be met for both rural and urban areas in South Asia and Sub-Saharan Africa. In other areas (*i.e.* East Asia and Pacific, Europe and Central Asia, Latin America and the Caribbean, and Middle East and North Africa), MDG1 will be achieved by 2015 for both rural and urban areas. As pointed out by World Bank (2013), the pace of poverty reduction is faster in urban areas than in rural areas. However, it should also be noted that the pace of poverty reduction in rural areas was accelerated in the period between 2002-2008 than in earlier periods in East Asia and Pacific and South Asia. It can be, however, safe to argue that reduction of poverty in rural areas in South Asia and Sub-Saharan Africa is likely to key to overall poverty reduction given the large population in these regions. Imai et al. (2014) used the secondary estimates of rural, urban and aggregate poverty rates for 31 developing countries and statistically examined the extent to which the rural sector contributes to aggregate poverty reduction (see Appendix 1). After adjusting for the effect of rural-urban migration, they have shown that the rural sector makes a substantial contribution to aggregate poverty reduction across all five regions. Recent studies giving greater priority to urbanisation, especially small and secondary towns (e.g. including World Bank, 2013, Collier and Dercon, 2014, Christiaensen and Todo, 2014) are likely to be misleading as discussed in details in Gaiha (2014).

As a supplementary analysis, we report summary statistics of "Multidimensional Poverty Indices (MPI)" separately for rural and urban areas. Because LSMS data do not have information necessary for constructing MPI, it is necessary to use Demographic and Health Surveys (DHS) data (http://dhsprogram.com/Data/) which are available for over 100 developing countries. There has been in the recent empirical literature a search for new poverty measures as the current MDGs do not fully reflect the multidimensional aspect of human development. The UNDP has proposed a MPI to capture micro-level deprivation. Alkire and Foster (2011) also develop an MPI aggregated from individuals to communities.

Oxford Poverty & Human Development Initiative (OPHI) has recently provided estimates of MPI separately for rural and urban areas at country levels. Table 2 summarises rural and urban MPI based on country estimates for 105 countries.³ It should be noted that Table 1 and Table 2 cannot be compared as Table 1 is based on estimates for much fewer countries. MPI (national, rural and urban) is an index ranging from 0 to 1 based on the share of people who are considered to be poor in multidimensional aspects (H), covering education (years of schooling and child school attendance, with the weight of 1/3), health (child mortality and nutrition, 1/3) and living standard (access to electricity, sanitation, drinking water, flooring,

³ The data are available from <u>http://www.ophi.org.uk/multidimensional-poverty-index/mpi-2014/mpi-data/</u>. Country-level estimates are reported in Appendix 2.

cooking fuel, TV, telephone etc. 1/3) and the intensity (A).⁴ The MPI reflects both the incidence, or headcount ratio (H) of poverty (or the proportion of the population that is multidimensionally poor)⁵ and the average intensity (A) of their poverty (or the average proportion of indicators in which poor people are deprived) (Alkire et al., 2014).

Table 2 reports regional averages of MPI for rural and urban areas. The last row shows that the share of households which are multidimensionally poor is much higher in rural areas (38.0%) than in urban areas (16.1%), which is broadly consistent with the estimates in 2008 based on the US\$1.25 poverty line in Table 1 (29.4% for rural areas and 11.6% for urban areas). Intensity of deprivation among the poor (A) is also higher in rural areas than in urban areas (0.075). The pattern - in which MPI, headcount ratio (H) and intensity (A) are higher in rural than in urban areas - is consistently found across different regions.⁶

⁴ See Alkire et al. (2014) for technical details.

⁵ A person is identified as poor if he or she is deprived in at least one third of the weighted indicators. Those identified as 'Vulnerable to Poverty' are deprived in 20% - 33.33% of weighted indicators and those identified as in 'Severe Poverty' are deprived in 50% or more of the dimensions (Alkire et al., 2014, p.4).

⁶ While the national estimates for MPI are more or less plausible (i.e., the highest MPI for SSA (0.329), to be followed by 0.187 (South Asia), East Asia and the Pacific (0.115), Latin America and Caribbean (0.055) and Europe and Central Asia (0.011)), the estimates for rural and urban MPI will have to be interpreted with caution because they are not consistent with the national estimates. For instance, it is counter-intuitive to find the highest value of rural MPI for East Asia and the Pacific, but this reflects the high MPI estimates for the Pacific countries (e.g. Vanuatu. See Appendix 2). The estimates for Sub-Saharan African (SSA) countries will also have to be scrutinised as there are a few countries with counter-intuitively low estimates for MPI (e.g. Benin, Burkina Faso, Central African Republic, Cote d'Ivoire. See Appendix 2). This will raise the issues of comparability of MPI estimates across different countries in particular for their rural and urban estimates. It is also counter-intuitive to find that rural MPI estimates for South Asia (0.126) than that for Europe and Central Asia (0.299) or for Latin America and Caribbean (0.169). Finding these inconsistencies is to some extent inevitable as MPI is derived by a number of qualitative questions and the data can be influenced by e.g. social or cultural norms.

| | | Mu | Itidimensional poverty | y - Rural | Mu | Itidimensional poverty | - Urban | Population | |
|--------------------------------|------------------------------------|-----------------|---|----------------------------------|-----------------|------------------------------------|----------------------------------|------------|-------------|
| | Multidimensional | | Headcount Ratio: | Intensity of deprivation | | Headcount ratio: Population in | Intensity of deprivation | shares | (%) |
| W orld region (average) | Poverty Index (MPI) National | MPI | Population in multidimensional poverty (H) | among the poor (A) | MPI | multidimensional poverty (H) | among the poor (A) | Rural | Urban |
| | | Range 0 to 1 | % Population | % of weighted deprivations | Range 0 to 1 | % Population | % of weighted deprivations | i turui | indi Ofbari |
| East Asia and the Pacific | 0.115 | 0.299 | 53.3 | 50.9 | 0.113 | 23.5 | 44.3 | 58.8 | 41.2 |
| Europe and Central Asia | 0.011 | 0.188 | 34.9 | 45.7 | 0.070 | 15.1 | 40.4 | 50.5 | 49.5 |
| Latin America and Caribbean | 0.055 | 0.169 | 31.5 | 45.0 | 0.061 | 13.5 | 40.2 | 55.9 | 44.1 |
| South Asia | 0.187 | 0.126 | 26.1 | 39.5 | 0.032 | 7.6 | 34.7 | 50.4 | 49.6 |
| Sub-Saharan Africa (SSA) | 0.329 | 0.223 | 41.3 | 46.8 | 0.084 | 17.7 | 41.9 | 56.4 | 43.6 |
| Total | 0.168 | 0.205 | 38.0 | 46.0 | 0.075 | 16.1 | 40.9 | 54.6 | 45.4 |

Table 2. Multidimensional Poverty Index (MPI) for Rural and Urban Areas

Data source: Calculated based on OPHI data (2014). Regional averages are derived as simple averages of country estimates for the countries for which the data area available (see Appendix 2 for the list of countries and estimates).

2. Methodological Issues

2.1. Backgrounds for the international poverty lines

This study will apply the international poverty lines based on \$1.25 and \$2 per capita per day (2005 PPP) separately for rural and urban areas using LSMS data. Before presenting the methodological details, we will provide a few background issues focusing on advantages and disadvantages using these poverty lines.

Obvious advantages include the comparability across different geographical areas/categories (e.g. across different countries as well as between rural and urban areas) as well as over different time periods on the grounds that the consumption or income is adjusted by using PPP (purchasing power parity) in 2005 (Chen and Ravallion, 2008; 2010).

In 2008 the World Bank revised the international poverty line for the extreme poverty and revised it from the US\$1.08 per day (1993 PPP) poverty line to US\$1.25 (2005 PPP) to increase the coverage of countries as detailed in Chen and Ravallion (2008), which provides revised poverty trends since 1981. The US\$1.25 line was adopted as the average of the

national poverty line of the poorest 15 countries in the new dataset in such a way that the new methodology is consistent with earlier ones (World Bank, 2014). The new method has an advantage as the new sample of national lines is representative of low income countries. That is, by using the US\$1.25 a day line, we can assess poverty situations broadly in terms of average nutritional requirement in low income countries or the equivalent levels in income or consumption. Using the revised poverty line, Chen and Ravallion (2010, p.1621) concluded that "25% of the population of the developing world, 1.4 billion people, were poor in 2005, which is 400 million more for that year 2005 than implied by our old international poverty line based on national lines for the 1980s and the 1993 ICP". The overall pattern of the results is unchanged if other alternative poverty lines are used (*ibid.*, 2010). Building on this, we propose to disaggregate the poverty estimates into rural and urban poverty.

Disadvantages of using the international poverty lines in the context of the present study include: (i) the lower poverty lines, such as the US\$ 1.25 line, may be inappropriate for some middle income countries – in particular for urban areas, (ii) the income or consumption poverty may not capture the actual nutritional situations of the countries, and (iii) nutritional requirements are likely to be different for rural and urban areas. To partially overcome these limitations, we proposed to use both the US\$ 1.25 and US\$2.00 lines. We also propose to convert detailed food expenditure data into nutrients' intakes for Malawi and for Indonesia.⁷

2.2. FGT Poverty Measure

The methodological framework to derive urban and rural poverty rate is founded in the wellknown Foster-Greer-Thorbecke (FGT) measure (Foster et al., 1984) denoted as P_{α} . The formula is given by:

$$P_{\alpha} = \frac{1}{N} \sum_{i=1}^{H} \left(\frac{z - y_i}{z} \right)^{\alpha} \tag{1}$$

⁷ The team is still in the middle of acquiring the datasets.

where z is US\$1.25 or US\$2.00 poverty line, N is the number of households in either rural or urban areas in the country, H is the number of poor households whose per capita income or consumption is below z, α is the sensitivity parameter which reflects the weight given to poorer households farther below z. We examine the cases where α is 0, 1 and 2.

Headcount ratio ($\alpha = 0$) is defined as:

$$P_0 = \frac{H}{N} \tag{2}$$

which indicates the fraction of households which are poor in either rural or urban area.

Poverty gap ($\alpha = 1$) is defined as:

$$P_1 = \frac{1}{N} \sum_{i=1}^{H} \left(\frac{z - y_i}{z} \right) \tag{3}$$

which indicates the amount of income necessary to bring everyone in poverty right up to the poverty line, divided by total population in either rural or urban area.

 P_2 measure ($\alpha = 2$) is defined as:

$$P_{2} = \frac{1}{N} \sum_{i=1}^{H} \left(\frac{z - y_{i}}{z} \right)^{2}$$
(4)

which reflects the degree of poverty as well as the inequality among the poor households in either rural or urban areas.

2.3. World Bank Estimates versus Our Estimates for International Poverty

As detailed in Ravallion et al. (2008), the World Bank has made the historical update of the international "\$1 a day" poverty line, which was first proposed by the World Bank in 1990 for measuring absolute poverty by the standards of the world's poorest countries (Ravallion et al. (1991). Since then poverty researchers have used two major thresholds, "\$1.25 a day" and "\$2 a day". As discussed in Ravallion et al. (1991, 2008) and World Bank (2015), the World has used only the distributional information of household/individual data or national census aggregated at sub-regional categories, such as, the proportion of population of each

category as well as that of income – which could be defined as cumulative ranked by income, or not cumulative (or alternatively, percentage of the population in a given class interval of incomes as well as the mean income of that class interval). ⁸ On the other hand, we use directly micro-level household data, such as LSMS data, to derive the estimates of poverty indices. Hence, at national levels, the World Bank estimates of poverty may be different from ours.

2.4. Inequality measures

We also derive estimates for inequality separately rural and urban areas for each country. Among many inequality measures, we will use the Gini coefficient and the Generalised Entropy (GE) measure.

The Gini coefficient derives from the Lorenz curve, a cumulative frequency curve that compares the distribution of a resource (here, per capita consumption expenditure) with the uniform distribution representing equality. It ranges from 0 to 1, representing perfect inequality with 1 (i.e. one member in a society holds all of the resource) and no inequality with 0 (all members shares same level of the resource).

The Gini coefficient can be formally defined as follows:

$$Gini = \frac{1}{2N^2 \bar{y}} \sum_{i=1}^{N} \sum_{j=1}^{N} |y_i - y_j|$$
(5)

⁸More specifically, the World Bank used one of the following types of data of distributional information to derive poverty measures (see <u>http://iresearch.worldbank.org/PovcalNet/index.htm?0,5</u>).

[•]Type 1: p=cumulative proportion of population (ranked by the poverty indicator, which we will call "income"), L=cumulative proportion of income held by that proportion of the population.

[•]Type 2: q=proportion of population (as in p, but not cumulative), r=proportion of income (as in L, but not cumulative).

[•]Type 3: p (as in 1), r (as in 2).

[•]Type 4: q (as in 2), L (as in 1).

[•]Type 5: f(x)=percentage of the population in a given class interval of incomes, X=the mean income of that class interval.

[•]Type 6: upper bound of a class interval, f(x) (as in 5), X (as in 5).

[•]Type 7: upper bound of a class interval, p (as in 5), X (as in 5).

[•]Type 8: upper bound of a class interval, f(x) (as in 5).

where \overline{y} is mean income (or consumption expenditure) and $|y_i - y_j|$ is all the pair-wise absolute differences in income. N denotes the number of population. The Gini coefficient is a full information measure drawing up all parts of distribution and the most widely used inequality measure in literature. It allows comparison in the distribution of a variable of interest across different populations. Besides, it also can be used for comparison of distribution over time regardless of whether inequality increases or decreases. However, it has a shortcoming that it often fails to meet decomposability (i.e. the sum of the Gini coefficients of population sub-groups is not equal to the total Gini coefficient of the population).

To supplement Gini coefficient, we also compute the GE measure, which is defined as:

$$GE(\theta) = \frac{1}{\theta(\theta - 1)} \left[\frac{1}{N} \sum_{i=1}^{N} \left(\frac{y_i}{\overline{y}} \right)^{\theta} - 1 \right]$$
(6)

where θ is a discretionary parameter that represents the weight given to distances between incomes at different parts of the income distribution, and can take any real value.

The value of GE measure ranges from zero to infinity, representing higher inequality with higher value. It is more sensitive to changes at the lower (upper) tail of the distribution (i.e. the poorest) for lower (higher) values of θ , and is equally sensitive to changes across the distribution for θ equal to 1.

For example, with θ =0, we obtain Theil's L index, often referred to as the Mean Log Deviation:

$$GE(0) = \frac{1}{N} \sum_{i=1}^{N} \log \frac{\overline{y}}{y_i}$$

$$\tag{7}$$

With $\theta = 1$, we obtain Theil's T index:

$$GE(1) = \frac{1}{N} \sum_{i=1}^{N} \frac{y_i}{\overline{y}} \log \frac{y_i}{\overline{y}}$$
(8)

Further supplement Gini coefficient and GE measures, the percentile ratios p90/p10 (or p75/p25) (the ratio of income at 90% (75%) to that at 10%(25%)).

2.5. Derivation of Rural and Urban Poverty Estimates using LSMS data

Deriving aggregate consumption

Using the World Bank Living Standard Measurement Survey (LSMS) data, we propose to use the consumption data to measure rural and urban poverty. If consumption data are not available, income data will be used. In the case where consumption data are used, consumption from own production will be included for each household. The total household consumption or income is then divided by the number of persons living in the household to derive a per capita measure.

More specifically, Deaton and Grosh (2000) and Deaton and Zaidi (2002) detail how the expenditure data are derived by LSMS data. First, it is necessary to add up all reported expenditures on individual goods and services or on groups of goods and services. Then a value for consumption out of home production or in kind received from employers will be added up. Then the estimates must be converted to real terms by adjusting them by a price index to account for differences in prices among different regions or interview dates.⁹

Deaton and Grosh (2000) have also noted limitations when we use consumption or expenditure data from LSMS. These include (i) recall errors associated with the fading of people's memories; (ii) the "telescoping" of reported events by incorrect dating; (iii) reporting errors associated with respondents being overwhelmed either by the length of the survey or by the number of items covered; (iv) "prestige" errors, in other words, misreporting due to various social pressures; (v) conditioning effects from being in the survey; (vi)

⁹ Examples of Stata codes aggregating expenditure are provided in the appendix of Deaton and Zaidi (2002).

respondent effects where the identity of the respondent affects the answers that they give; (vii) interviewer effects; and (viii) effects associated with the design of the instrument. Similar errors are expected for income data. These may result in errors in classifying households as poor or non-poor and thus the final results should be interpreted with caution.

Price adjustments

As discussed in Deaton and Grosh (2000), unlike other similar surveys, LSMS is supposed to provide an estimate of annual expenditures at the household level, implying that consumer price index can be constructed either at the household level, or at the community level. While there are both advantages and disadvantages in these two, in light of the main purpose of the project (i.e. classifying households as poor or non-poor), it is appropriate to use price index at the community or PSU (primary sampling unit) averages because measurement errors for quantity and price for each item can be correlated at the household level.

It is necessary for the researchers to impute values of all the items if they are not purchased at the market (e.g. gifts, own farm production, home-made clothes or wood and water fetched by children or women) (Deaton and Grosh, 2000). Ideally, we need to derive the imputed values for all these items, but as this is a time consuming work, we will focus only on the home-produced food as well as food received as gifts or payment in kind, which are likely to be the largest and non-negligible in LSMS data. The respondent is asked to report the value of any home-produced food consumed by the household during the reference period, and then the sum of these items is added in to the consumption total (*ibid.*, 2000).

While the community-level price index should ideally be used to derive the real values of household consumption over different time periods, as these variables are not necessarily available in LSMS data, we will use consumption price index (CPI, available from World Development Indicators (WDI) in 2014) to convert the final private household consumption expenditure per capita per day in local currency unit (LCU) in the survey year (e.g. 2011) to

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2005, the base year for international poverty indices. Then they will be converted to US\$ in 2005 using the Purchasing Power Parity (PPP) conversion factor which is available in WDI in 2014. The household is defined as poor if per capita per day annual household consumption is below US\$1.25, or US\$2, in terms of PPP in 2005 and as non-poor otherwise. The same procedure is used to derive poverty rate for the US\$2 line.

3. Country Case Studies¹⁰

Tanzania

First we will report poverty and inequality estimates for Tanzania based on LSMS data in 2008 and 2010. The results on inequality and poverty are summarised in Tables 3 and 4 and the regression results about determinants of poverty and vulnerability are reported in Table 5.

In the first panel of Table 3, percentile ratios for distribution are shown. The figures imply that the magnitude of inequality is large in Tanzania. This is reflected in Gini coefficient which remained relatively large -0.4557 in 2008 and 0.4467 in 2010. If we decompose the Gini coefficient in rural and urban areas, we will find that Gini remained higher in rural area than urban area, with the gap having narrowed from 2008 to 2010. GE measures have been decomposed into within-inequality (the inequality within rural or urban area) and between-inequality (the inequality between rural and urban areas). The results show that the former contributes more to the overall inequality.

Table 4 summarises poverty in rural and urban areas in Tanzania. The first panel shows the results for the \$1.25 poverty line in 2008 and 2010. Poverty headcount ratio – as well as poverty gap and squared poverty gap- is much higher in rural area than urban area. The pattern of the results is similar for \$2.00, which are shown in the second panel.

¹⁰ More countries will be covered by Part II to be prepared by Dr Gordon Abekah-Nkrumah and Dr Purnima Purohit.

Table 3 Inequality in Tanzania

| 2008 | | | | | | | 2010 | | | | | | |
|--|--------------------|--|--------------------------------------|--------------------------------------|-----------------------------|-----------------|--|-------------------------|--|--------------------------------------|--------------------------------------|-----------------------------|-----------------|
| Percentile ra | tios for distrib | bution of exp | om_pc_2005 | all valid obs. | | | Percentile ra | atios for distril | bution of exp | om_pc_2005 | all valid obs. | | |
| p90/p10 | p90/p50 | p10/p50 | p75/p25 | p75/p50 | p25/p50 | | p90/p10 | p90/p50 | p10/p50 | p75/p25 | p75/p50 | p25/p50 | |
| 7.375 | 3.201 | 0.434 | 2.841 | 1.766 | 0.622 | | 7.443 | 3.17 | 0.426 | 2.812 | 1.774 | 0.631 | |
| | 15 | | | me difference | | | | 1. | | | me difference | | |
| 5 1 | parameter, and | | | | | ~· · | 21 | parameter, and | | | | | <i>~</i> |
| All | Obs | GE(-1) | GE(0) | GE(1) | GE(2) | Gini | All | obs | GE(-1) | GE(0) | GE(1) | GE(2) | Gini |
| | | 0.42886 | 0.34561 | 0.37495 | 0.56056 | 0.45567 | | | 0.41763 | 0.33344 | 0.35881 | 0.53681 | 0.44671 |
| Subgroup su | nmary statist | ics, foreach | subgroup k = | = 1,,K: Income | | | Subgroup su | mmary statist | ics, foreach | subgroup k = | = 1,,K: Income | | |
| rural/urban | Pop. share | Mean | Rel.mean | share | log(mean |) | rural/urban | Pop. share | Mean | Rel.mean | share | log(mean |) |
| Rural | 0.63185 | 1.84619 | 0.64055 | 0.40474 | 0.61312 | | Rural | 0.67196 | 1.99567 | 0.70763 | 0.4755 | 0.69098 | |
| Urban | 0.36815 | 4.66024 | 1.61692 | 0.59526 | 1.53907 | | Urban | 0.32804 | 4.50916 | 1.59888 | 0.5245 | 1.50611 | |
| | | | | | | | | | | | | | |
| Subgroup | indices: | GE_k(a) | And | Gini_k | | | Subgroup | indices: | GE_k(a) | and | Gini_k | | |
| 0 1 | | GE_k(a) GE(-1) | And GE(0) | Gini_k GE(1) | GE(2) | Gini | Subgroup rural/urbai | | GE_k(a) GE(-1) | and GE(0) | Gini_k GE(1) | GE(2) | Gini |
| rural/urba | | , | | | GE(2) 0.31594 | Gini 0.35195 | 0 1 | | , | | | GE(2) 0.43419 | Gini 0.38314 |
| rural/urbai Rural | | GE(-1) | GE(0) | GE(1) | | | rural/urba | | GE(-1) | GE(0) | GE(1) | | |
| rural/urbai Rural Urban Within- | 1 | GE(-1) 0.23192 0.40864 | GE(0) 0.20374 0.30515 | GE(1) 0.22299 | 0.31594 | 0.35195 | rural/urban Rural Urban Within- | 1 | GE(-1) 0.28011 0.36196 | GE(0) 0.24293 0.27975 | GE(1) 0.27364 | 0.43419 | 0.38314 |
| rural/urban Rural Urban Within- group | inequality, | GE(-1) 0.23192 0.40864 GE_W(a) | GE(0) 0.20374 0.30515 | GE(1) 0.22299 0.30061 | 0.31594 0.38193 | 0.35195 | ru ral/urban Ru ral Urban Within- group | n inequality, | GE(-1) 0.28011 0.36196 GE_W(a) | GE(0) 0.24293 0.27975 | GE(1) 0.27364 0.28024 | 0.43419 0.36134 | 0.38314 |
| Subgroup rural/urban Rural Urban Within- group All | 1 | GE(-1) 0.23192 0.40864 | GE(0) 0.20374 0.30515 | GE(1) 0.22299 | 0.31594 | 0.35195 | rural/urban Rural Urban Within- | 1 | GE(-1) 0.28011 0.36196 | GE(0) 0.24293 0.27975 | GE(1) 0.27364 | 0.43419 | 0.38314 |
| rural/urban Rural Urban Within- group | inequality, | GE(-1) 0.23192 0.40864 GE_W(a) | GE(0) 0.20374 0.30515 | GE(1) 0.22299 0.30061 | 0.31594 0.38193 | 0.35195 | ru ral/urban Ru ral Urban Within- group | n inequality, | GE(-1) 0.28011 0.36196 GE_W(a) | GE(0) 0.24293 0.27975 | GE(1) 0.27364 0.28024 | 0.43419 0.36134 | 0.38314 |
| rural/urban Rural Urban Within- group All | inequality, | GE(-1) 0.23192 0.40864 GE_W(a) GE(-1) 0.32181 | GE(0) 0.20374 0.30515 GE(0) | GE(1) 0.22299 0.30061 GE(1) | 0.31594 0.38193 GE(2) | 0.35195 | rural/urban Rural Urban Within- group All | n inequality, | GE(-1) 0.28011 0.36196 GE_W(a) GE(-1) 0.34026 | GE(0) 0.24293 0.27975 GE(0) | GE(1) 0.27364 0.28024 GE(1) | 0.43419 0.36134 GE(2) | 0.38314 |
| rural/urban Rural Urban Within- group All | inequality, Obs | GE(-1) 0.23192 0.40864 GE_W(a) GE(-1) 0.32181 | GE(0) 0.20374 0.30515 GE(0) | GE(1) 0.22299 0.30061 GE(1) | 0.31594 0.38193 GE(2) | 0.35195 | rural/urban Rural Urban Within- group All | n inequality, obs | GE(-1) 0.28011 0.36196 GE_W(a) GE(-1) 0.34026 | GE(0) 0.24293 0.27975 GE(0) | GE(1) 0.27364 0.28024 GE(1) | 0.43419 0.36134 GE(2) | 0.38314 |

Table 4 Poverty in Tanzania

| 2008 | | | | | 2010 | | | | |
|--------------|-----------------|--------------|-------------|------------------|--------------|-----------------|----------------|-------------|------------------|
| Based on \$1 | 1.25 Poverty | Line | | | Based on \$ | 1.25 Poverty | Line | | |
| Foster-Gree | r-Thorbecke | poverty indi | ces, FGT(a) | | Foster-Gree | r-Thorbecke | poverty indic | ces, FGT(a) | |
| All | obs | a=0 | a=1 | a=2 | All | obs | a=0 | a=1 | a=2 |
| national | | 0.28484 | 0.08217 | 0.03333 | national | | 0.282 | 0.08325 | 0.03496 |
| FGT(0): hea | dcount ratio | (proportion | poor) | | FGT(0): hea | adcount ratio | (proportion p | boor) | |
| FGT(1): ave | erage normali | sed poverty | gap | | FGT(1): ave | erage normali | ised poverty | gap | |
| FGT(2): ave | erage squared | normalised | poverty gap | | FGT(2): ave | erage squared | normalised | poverty gap | |
| Summary sta | atistics for su | bgroup k = | 1,,K | | Summary st | atistics for su | ıbgroup k = 1 | ,,K | |
| 1/ 1 | Pop. | | | Mean gap | 1/ 1 | Pop. | X | | Mean gap |
| rural/urban | share | Mean | Meanpoor | poor | rural/urban | share | Mean | Meanpoor | poor |
| rural | 0.63185 | 1.84619 | 0.88176 | 0.36824 | 0 | 0.67196 | 1.99567 | 0.87256 | 0.37744 |
| urban | 0.36815 | 4.66024 | 0.94008 | 0.30992 | 1 | 0.32804 | 4.50916 | 0.96035 | 0.28965 |
| Subgroup FC | 3T index esti | mates, FGT | (a) | | Subgroup FC | GT index esti | imates, FGT | (a) | |
| rural | | a=0 | a=1 | a=2 | rural | | a=0 | a=1 | a=2 |
| rural | | 0.39166 | 0.11538 | 0.04735 | 0 | | 0.3794 | 0.11456 | 0.04861 |
| urban | | 0.1015 | 0.02516 | 0.00927 | 1 | | 0.08247 | 0.01911 | 0.00698 |
| 2008 | | | | | 2010 | | | | |
| Based on \$2 | 2 Poverty Li | ne | | | Based on \$2 | 2 Poverty Li | ne | | |
| Foster-Gree | r-Thorbecke | poverty indi | ces, FGT(a) | | Foster-Gree | r-Thorbecke | poverty indic | es, FGT(a) | |
| All | obs | a=0 | a=1 | a=2 | All | obs | a=0 | a=1 | a=2 |
| | | 0.53752 | 0.21024 | 0.10532 | | | 0.5398 | 0.21012 | 0.10594 |
| FGT(0): hea | dcount ratio | (proportion | poor) | | FGT(0): hea | adcount ratio | (proportion p | boor) | |
| FGT(1): ave | erage normali | sed poverty | gap | | FGT(1): ave | erage normali | ised poverty g | gap | |
| FGT(2): ave | erage squared | normalised | poverty gap | | FGT(2): ave | erage squared | normalised | poverty gap | |
| Summary sta | | bgroup k = | l ,,K | | Summary st | | ıbgroup k = 1 | ,,K | |
| rural/urban | Pop. share | Mean | Meanpoor | Mean gap poor | rural/urban | Pop. share | Mean | Meanpoor | Mean Gap Poor |
| Rural | 0.63185 | 1.84619 | 0.88176 | 0.36824 | rural | 0.67196 | 1.99567 | 1.18778 | 0.81222 |
| Urban | 0.36815 | 4.66024 | 0.94008 | 0.30992 | urban | 0.32804 | 4.50916 | 1.41198 | 0.58802 |
| Subgroup FC | 3T index esti | mates, FGT | (a) | | Subgroup FC | GT index esti | imates, FGT | (a) | |
| rural a= | :0 | a=0 | a=1 | a=2 | rural a= | =0 | a=0 | a=1 | a=2 |
| Rural | | 0.70868 | 0.28622 | 0.1455 | rural | | 0.68254 | 0.27719 | 0.14274 |
| | | 0.24376 | 0.07984 | 0.0363 | urban | | 0.24742 | 0.07275 | 0.03056 |

In Table 5 the determinants of poverty and vulnerability in Tanzania are presented. Regardless of whether poverty or vulnerability is adopted, the pattern of the results is similar. For instance, better education of head is associated with lower poverty or vulnerability. Household headed by older person tends to be poorer and more vulnerable. Female headed households are more likely to be poorer and more vulnerable. Larger household size and higher dependency burden are associated with poverty and vulnerability.

| | Case (1) Random- | Case (2) Random- | Case (3) Fixed- | Case (4) Random- | Case (5) Fixed- | Case (6) Random- |
|-----------|---------------------|---------------------|----------------------|-----------------------|----------------------|-----------------------|
| | Effects | Effects | Effects | Effects | Effects | Effects |
| | Probit | Probit | Model | Model | Model | Model |
| Dep. Var | Poverty | Poverty | Vulnerability | Vulnerability | Vulnerability | Vulnerability |
| • | \$1.25 | \$2.00 | \$1.25 | \$1.25 | \$2.00 | \$2.00 |
| VARIABLES | | | | | | |
| educ_hhh | -0.0806*** | -0.104*** | -0.00582*** | -0.00468*** | -0.00640*** | -0.00530*** |
| | (0.00969) | (0.00928) | (0.000308) | (0.000190) | (0.000359) | (0.000196) |
| age_hhh | 0.00752*** | 0.00653*** | 0.000600** | 0.000361*** | 0.000608* | 0.000566*** |
| | (0.00239) | (0.00244) | (0.000272) | (8.61e-05) | (0.000317) | (8.51e-05) |
| male_hhh | -0.207*** | -0.0511 | -0.0348*** | -0.0264*** | -0.0417*** | -0.0328*** |
| | (0.0788) | (0.0798) | (0.00547) | (0.00260) | (0.00637) | (0.00262) |
| Hhsize | 0.122*** | 0.177*** | 0.0212*** | 0.0222*** | 0.0291*** | 0.0285*** |
| | (0.0129) | (0.0145) | (0.000582) | (0.000374) | (0.000678) | (0.000390) |
| dep_ratio | 0.268*** | 0.395*** | 0.0193*** | 0.0192*** | 0.0199*** | 0.0201*** |
| | (0.0322) | (0.0390) | (0.00124) | (0.000946) | (0.00145) | (0.00101) |
| Mainland | -0.601** | 0.0198 | | 0.483*** | | 0.456*** |
| | (0.259) | (0.279) | | (0.0107) | | (0.0105) |
| D2008 | 0.0265 | 0.0199 | -0.00497*** | -0.00702*** | -0.00643*** | -0.00920** |
| | (0.0501) | (0.0498) | (0.00100) | (0.000845) | (0.00117) | (0.000975) |
| Rural | -0.722*** | -0.764*** | 0.0133*** | 0.0173*** | 0.0156*** | 0.0188*** |
| | (0.0779) | (0.0698) | (0.00243) | (0.00190) | (0.00283) | (0.00205) |
| 2.region | -0.986*** | -0.913*** | -0.0462 | 0.0521*** | -0.0688* | 0.0580*** |
| U | (0.254) | (0.264) | (0.0316) | (0.00979) | (0.0368) | (0.00968) |
| 3.region | -0.791*** | -0.793*** | 0.00634 | 0.134*** | -0.00113 | 0.150*** |
| 8 | (0.240) | (0.255) | (0.0354) | (0.00970) | (0.0413) | (0.00956) |
| 4.region | -0.582** | -0.231 | 0.244*** | 0.519*** | -0.0856 | 0.498*** |
| | (0.226) | (0.256) | (0.0462) | (0.00971) | (0.0538) | (0.00955) |
| 5.region | -0.502** | -0.336 | -0.104*** | -0.0542*** | -0.179*** | -0.102*** |
| o negion | (0.225) | (0.249) | (0.0265) | (0.00914) | (0.0309) | (0.00910) |
| 6.region | -1.306*** | -1.075*** | -0.00510 | -0.0425*** | -0.0604* | -0.101*** |
| ollegion | (0.303) | (0.286) | (0.0299) | (0.0105) | (0.0349) | (0.0104) |
| 7.region | -1.985*** | -2.082*** | -0.0690*** | -0.0369*** | -0.151*** | -0.0987*** |
| , | (0.250) | (0.240) | (0.0243) | (0.00789) | (0.0284) | (0.00786) |
| 8 region | -0.0740 | 0.0826 | (0.0243) 0.517*** | (0.00789) 0.466*** | (0.0284) 0.432*** | (0.00780) |
| 8.region | (0.216) | (0.246) | $(0.0517)^{+++}$ | (0.00931) | (0.432^{+++}) | (0.00917) |
| 0 region | -0.369* | -0.277 | (0.0311) 0.601*** | (0.00931) 0.510*** | (0.0393) 0.648*** | (0.00917) 0.596*** |
| 9.region | | | | | | |
| 10 marian | (0.212) | (0.237) | (0.0335) | (0.00885) | (0.0390) | (0.00875) |
| 10.region | 0.141 | 0.425* | 0.199*** | 0.169*** | 0.189*** | 0.182*** |
| 11 | (0.210) | (0.248) | (0.0380) | (0.00911) | (0.0443) | (0.00898) |
| 11.region | -0.661*** | -0.498** | 0.168*** | 0.333*** | 0.328*** | 0.488*** |
| 10 | (0.228) | (0.249) | (0.0312) | (0.00939) | (0.0363) | (0.00929) |
| 12.region | -0.502** | -0.141 | 0.234*** | 0.318*** | 0.272*** | 0.346*** |
| | (0.214) | (0.242) | (0.0408) | (0.00910) | (0.0476) | (0.00897) |
| 13.region | -0.0231 | -0.163 | 0.233*** | 0.233*** | 0.296*** | 0.297*** |
| | (0.251) | (0.291) | (0.0340) | (0.0108) | (0.0396) | (0.0107) |
| 14.region | -0.195 | -0.152 | -0.0389 | -0.0327*** | -0.0689** | -0.0590*** |
| | (0.224) | (0.260) | (0.0284) | (0.00936) | (0.0331) | (0.00929) |
| 15.region | -0.125 | 0.314 | 0.0975** | 0.0824*** | 0.132*** | 0.0881*** |
| | (0.233) | (0.275) | (0.0382) | (0.00989) | (0.0445) | (0.00976) |
| 16.region | 0.361 | 0.653** | 0.191*** | 0.217*** | 0.210*** | 0.217*** |

Table 5 Determinants of Poverty and Vulnerability in Tanzania

| | (0.223) | (0.281) | (0.0419) | (0.00955) | (0.0488) | (0.00942) |
|--------------------|-----------|-----------|------------|------------|-----------|------------|
| 17.region | -0.462** | -0.563** | -0.0954*** | -0.0577*** | -0.145*** | -0.102*** |
| | (0.215) | (0.246) | (0.0339) | (0.00906) | (0.0395) | (0.00896) |
| 18.region | -0.735*** | -0.680*** | -0.00117 | -0.0204** | 0.0266 | -0.0374*** |
| | (0.221) | (0.247) | (0.0335) | (0.00930) | (0.0390) | (0.00919) |
| 19.region | -0.264 | -0.615** | -0.0455 | -0.0876*** | -0.0775** | -0.155*** |
| | (0.217) | (0.248) | (0.0304) | (0.00896) | (0.0355) | (0.00888) |
| 20.region | -0.0773 | -0.348 | 0.0476 | 0.0499*** | 0.0478 | 0.0317*** |
| | (0.263) | (0.306) | (0.0397) | (0.0110) | (0.0463) | (0.0109) |
| 21.region | -0.285 | -0.357 | -0.116*** | -0.0502*** | -0.195*** | -0.0973*** |
| | (0.248) | (0.284) | (0.0247) | (0.0103) | (0.0288) | (0.0103) |
| 51.region | 0.250 | 0.258 | 0.223*** | -0.00337 | 0.255*** | -0.000438 |
| | (0.263) | (0.290) | (0.0538) | (0.0113) | (0.0627) | (0.0110) |
| 52.region | -0.427 | -0.291 | 0.366*** | -0.0316*** | 0.385*** | 0.0168 |
| | (0.344) | (0.307) | (0.0422) | (0.0120) | (0.0492) | (0.0119) |
| 53.region | -0.274 | -0.921*** | 0.311*** | -0.0483*** | 0.382*** | 0.0420*** |
| | (0.228) | (0.221) | (0.0376) | (0.00876) | (0.0438) | (0.00848) |
| 54.region | 0.495* | 0.183 | | -0.00359 | | -0.00753 |
| | (0.272) | (0.292) | | (0.0117) | | (0.0113) |
| Constant | 1.297*** | 1.556*** | 0.104*** | -0.431*** | 0.168*** | -0.336*** |
| | (0.434) | (0.479) | (0.0253) | (0.0172) | (0.0294) | (0.0170) |
| Observations | 5,294 | 5,294 | 5,294 | 5,294 | 5,294 | 5,294 |
| Number of hhid2 | 3,116 | 3,116 | 3,116 | 3,116 | 3,116 | 3,116 |
| R-squared | | | 0.657 | | 0.709 | |
| Standard errors in | | | | | | |

parentheses *** p<0.01, ** p<0.05, *

Uganda

As is observed in Table 6, the magnitude of inequality was high in Uganda as well and it gradually increased from 2009 to 2011. In particular, urban inequality in terms of Gini coefficient expanded by 4% from 2010 to 2011.

In Table 7, statistics of poverty in Uganda are summarised. Overall, poverty increased from 2009 to 2011 in both rural and urban areas. In particular, urban poverty increased from 2010 to 2011 (19.2% to 25.1% for headcount ratio based on \$1.25; 39.2% to 48.1% for headcount ratio based on \$2).

Table 8 shows the econometric results of determinants of poverty and vulnerability. Age of household head is positively associated with vulnerability in a few cases. Surprisingly, educational level of household head is positive and significant for poverty and vulnerability, while its square is negative and significant in all the cases. A larger household tends to be poorer and more vulnerable, while the household headed by a female member tends to be

p<0.1

poorer and more vulnerable. While increase of female share in the household reduces poverty and vulnerability, the increase in female members below 15 years or above 60 years tends to increase poverty and vulnerability significantly.

| Table | 6: | Ineq | uality | in | Uganda |
|-------|----|------|--------|----|--------|
| | | | | | |

| | | Gini Coefficient | Percentile ra | tios for distribu | tion of expm | nc. 2005: all va | lidobs | |
|------|------------|---------------------|---------------|-------------------|--------------|------------------|---------|---------|
| | | Coernelent | p90/p10 | p90/p50 | p10/p50 | p75/p25 | p75/p50 | p25/p50 |
| 2009 | Rural | 0.392792 | 1 1 | 1 1 | 1 | 1 1 | 1 | 1 |
| | Urban | 0.449661 | 7.274 | 2.955 | 0.406 | 2.763 | 1.705 | 0.617 |
| | Population | 0.463589 | 7.274 | 2.955 | 0.406 | 2.763 | 1.705 | 0.617 |
| | | | | | | | | |
| 2010 | | Gini Coefficient | p90/p10 | p90/p50 | p10/p50 | p75/p25 | p75/p50 | p25/p50 |
| | Rural | 0.416263 | | | | | | |
| | Urban | 0.435774 | 7.317 | 2.917 | 0.399 | 2.768 | 1.694 | 0.612 |
| | Population | 0.462197 | | | | | | |
| | | Gini | | | | | | |
| 2011 | | Coefficient | p90/p10 | p90/p50 | p10/p50 | p75/p25 | p75/p50 | p25/p50 |
| | Rural | 0.426867 | | | | | | |
| | Urban | 0.476041 | 7.021 | 2.963 | 0.422 | 2.691 | 1.715 | 0.637 |
| | Population | 0.466359 | | | | | | |

Table 7: Poverty and vulnerability incidence in Uganda

| | | | Pove | erty | | Vulnerability |
|------|----------|------------|---------|---------|---------|---------------|
| | | | FGT(0): | FGT(1): | FGT(2): | Mean |
| 2009 | US\$1.25 | Urban | 20.8 | 7.1 | 3.4 | 2.3 |
| | | Rural | 52.8 | 19.2 | 9.4 | 39.2 |
| | | Population | 44.6 | 16.1 | 7.9 | 29.7 |
| | US\$2.00 | Urban | 36.5 | 15.0 | 8.3 | 17.0 |
| | | Rural | 79.0 | 37.4 | 21.6 | 69.9 |
| | | Population | 68.0 | 31.6 | 18.1 | 56.3 |
| 2010 | US\$1.25 | Urban | 19.2 | 6.6 | 3.2 | 3.3 |
| | | Rural | 55.1 | 21.5 | 11.0 | 41.3 |
| | | Population | 47.2 | 18.2 | 9.3 | 32.9 |
| | US\$2.00 | Urban | 39.2 | 15.2 | 8.0 | 19.2 |
| | | Rural | 79.6 | 39.5 | 23.6 | 72.0 |
| | | Population | 70.7 | 34.1 | 20.2 | 60.3 |
| 2011 | US\$1.25 | Urban | 25.1 | 8.5 | 4.1 | 2.9 |

| | Rural | 55.5 | 20.9 | 10.3 | 38.8 |
|----------|------------|------|------|------|------|
| | Population | 49.2 | 18.4 | 9.1 | 31.5 |
| US\$2.00 | Urban | 48.1 | 18.7 | 10.1 | 20.3 |
| | Rural | 78.4 | 38.8 | 23.0 | 68.6 |
| | Population | 72.2 | 34.7 | 20.3 | 58.6 |

FGT (0): headcount ratio (proportion poor)

FGT(1): average normalised poverty gap

FGT (2): average squared normalised poverty gap

Table 8 Determinants of Poverty and Vulnerability

| | Case (1) | Case (2) | Case (3) | Case (4) | Case (5) | Case (6) |
|------------------------------|-------------|------------|---------------|---------------|---------------|---------------|
| | Random- | Random- | Fixed- | Random- | Fixed- | Random- |
| | Effects | Effects | Effects | Effects | Effects | Effects |
| | Probit | Probit | Model | Model | Model | Model |
| Dep. Var | Poverty | Poverty | Vulnerability | Vulnerability | Vulnerability | Vulnerability |
| | \$1.25 | \$2.00 | \$1.25 | \$1.25 | \$2.00 | \$2.00 |
| head_age | -0.00231 | 0.0122 | 0.00427 | 0.00335* | 0.00449 | 0.00532*** |
| | (0.00921) | (0.00976) | (0.00398) | (0.00172) | (0.00369) | (0.00155) |
| head_age2 | -1.60e-05 | -0.000138 | -2.86e-05 | -3.95e-05** | -3.01e-05 | -5.16e-05** |
| | (8.97e-05) | (9.54e-05) | (3.75e-05) | (1.68e-05) | (3.49e-05) | (1.52e-05) |
| head_edu | 0.0206 | 0.0444** | 0.0264*** | 0.0166*** | 0.0513*** | 0.0368*** |
| | (0.0167) | (0.0176) | (0.00385) | (0.00277) | (0.00357) | (0.00253) |
| head_edu2 | -0.00910*** | -0.0103*** | -0.00359*** | -0.00285*** | -0.00641*** | -0.00543** |
| | (0.00139) | (0.00138) | (0.000313) | (0.000217) | (0.000290) | (0.000198) |
| Hsize | 0.123*** | 0.145*** | 0.0435*** | 0.0398*** | 0.0413*** | 0.0386*** |
| | (0.00807) | (0.00934) | (0.00244) | (0.00137) | (0.00226) | (0.00124) |
| head_sex (whether male head) | -0.151*** | -0.0881 | -0.0687*** | -0.0279*** | -0.0353* | -0.0340*** |
| _ 、 , | (0.0574) | (0.0633) | (0.0210) | (0.0105) | (0.0195) | (0.00956) |
| 1.region | 0.243 | 0.549*** | -0.000307 | -0.126*** | 0.0236 | -0.0588*** |
| C | (0.149) | (0.126) | (0.0210) | (0.0174) | (0.0195) | (0.0159) |
| 2.region | 1.266*** | 1.511*** | 0.335** | 0.168*** | 0.353** | 0.359*** |
| C | (0.154) | (0.138) | (0.161) | (0.0203) | (0.149) | (0.0185) |
| 3.region | 1.434*** | 1.693*** | · · · | 0.268*** | | 0.439*** |
| C | (0.153) | (0.137) | | (0.0200) | | (0.0182) |
| 4.region | 1.090*** | 1.365*** | | 0.160*** | | 0.374*** |
| C | (0.154) | (0.137) | | (0.0205) | | (0.0186) |
| rural | 0.729*** | 0.860*** | 0.129*** | 0.184*** | 0.228*** | 0.250*** |
| | (0.0647) | (0.0646) | (0.0229) | (0.0113) | (0.0213) | (0.0103) |
| burden_share_female | 1.151*** | 1.595*** | 0.901*** | 0.907*** | 0.872*** | 0.819*** |
| | (0.160) | (0.170) | (0.0406) | (0.0270) | (0.0377) | (0.0246) |
| female_share | -0.539*** | -0.420*** | -0.219*** | -0.179*** | -0.285*** | -0.208*** |
| | (0.154) | (0.155) | (0.0411) | (0.0254) | (0.0382) | (0.0231) |
| 2010.year | 0.0296 | -0.000796 | 0.00837 | 0.00970 | 0.0131** | 0.0154*** |
| | (0.0423) | (0.0469) | (0.00618) | (0.00603) | (0.00574) | (0.00557) |
| 2011.year | 0.121*** | 0.0931* | -0.0165** | -0.00637 | -0.0143** | -0.00963* |
| · | (0.0424) | (0.0479) | (0.00656) | (0.00612) | (0.00609) | (0.00565) |
| Constant | -2.013*** | -2.114*** | -0.306*** | -0.323*** | -0.106 | -0.261*** |
| | (0.266) | (0.267) | (0.108) | (0.0447) | (0.101) | (0.0405) |

| Observations | 8,478 | 8,478 | 8,372 | 8,372 | 8,372 | 8,372 |
|----------------|-------|-------|-------|-------|-------|-------|
| Number of HHID | 3,220 | 3,220 | 3,196 | 3,196 | 3,196 | 3,196 |
| R-squared | | | 0.201 | | 0.270 | |

Standard errors in parentheses

*** p<0.01, ** p<0.05, * p<0.1

5. Concluding Observations

This paper has set out the backgrounds for the construction of new rural and urban poverty and inequality estimates using LSMS data with focus on methodological details as well as on their advantages or disadvantages. First, we have reviewed the recent regional estimates based on the US\$1.25 per day poverty line as well as those based on Multidimensional Poverty Index (MPI) for both rural and urban areas. It has been found that the level of poverty is much higher in rural areas than in urban areas across different regions regardless of the definitions of poverty. Second, we have summarised estimates of poverty and inequality for Tanzania and Uganda based on recent panel data constructed by LSMS.

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| Country | Period (t) | | Rural Poverty Rate (RPR) | | Urban Poverty Rate (UPR) | | Aggregate Poverty Rate (APR) | | Change APR | Share of Rural in Population (SRP) | | Contribution of Rural Sector Aggregate Poverty Reduction (%) | | |
|-------------|------------|------|-----------------------------|------------|-----------------------------|------------|---------------------------------|------------|---------------|---------------------------------------|---------|--|--------------------------|-----------------------|
| Country y | t-1 | t | RPR (t- 1) | RPR (t) | UPR (t- 1) | UPR (t) | APR (t-1) | APR (t) | | SRP (t- 1) | SRP (t) | Migratio n of Non Poor | Neutral Migratio n | Poor Migratio n |
| | | | , | | E | | ope and Cen | tral Asia | | , | | | | |
| Albania | 2002 | 2008 | 0.296 | 0.146 | 0.195 | 0.101 | 0.2518 | 0.1234 | -0.128 | 0.5623 | 0.4989 | 72.94 | 58.30 | 23.48 |
| Azerbaijan | 2001 | 2008 | 0.425 | 0.185 | 0.557 | 0.148 | 0.4931 | 0.1654 | -0.328 | 0.4841 | 0.4700 | 36.25 | 34.42 | 31.95 |
| Moldova | 2009 | 2010 | 0.363 | 0.303 | 0.126 | 0.104 | 0.2536 | 0.2096 | -0.044 | 0.5382 | 0.5306 | 78.67 | 72.44 | 61.50 |
| Montenegro | 2006 | 2008 | 0.176 | 0.089 | 0.074 | 0.024 | 0.1124 | 0.0482 | -0.064 | 0.3762 | 0.3726 | 51.52 | 50.53 | 45.93 |
| Serbia | 2004 | 2007 | 0.202 | 0.098 | 0.104 | 0.043 | 0.1489 | 0.0677 | -0.081 | 0.4586 | 0.4493 | 59.83 | 57.52 | 48.38 |
| Tajikistan | 2003 | 2007 | 0.738 | 0.55 | 0.688 | 0.494 | 0.7248 | 0.5352 | -0.190 | 0.7355 | 0.7354 | 72.95 | 72.93 | 72.92 |
| Ukraine | 2002 | 2008 | 0.351 | 0.047 | 0.247 | 0.02 | 0.2809 | 0.0286 | -0.252 | 0.3260 | 0.3167 | 39.44 | 38.16 | 35.78 |
| | South Asia | | | | | | | | | | | | | |
| Bangladesh | 1992 | 2010 | 0.587 | 0.3516 | 0.427 | 0.2128 | 0.5541 | 0.3129 | -0.241 | 0.7944 | 0.7211 | 88.21 | 70.37 | 57.82 |
| India | 1993 | 2009 | 0.5246 | 0.3428 | 0.4077 | 0.2893 | 0.4940 | 0.3264 | -0.168 | 0.7077 | 0.6907 | 80.25 | 74.94 | 70.13 |
| Pakistan | 2002 | 2006 | 0.393 | 0.27 | 0.227 | 0.131 | 0.3371 | 0.2217 | -0.115 | 0.6632 | 0.6524 | 73.22 | 69.52 | 63.81 |
| | | | | | | East Asi | a and the P | acific | | | | | | |
| China | 1996 | 2009 | 0.4948 | 0.2057 | 0.0887 | 0.0064 | 0.3651 | 0.1103 | -0.255 | 0.6806 | 0.5748 | 85.75 | 65.21 | 44.24 |
| Indonesia | 1996 | 2011 | 0.4675 | 0.1497 | 0.376 | 0.174 | 0.4338 | 0.1620 | -0.272 | 0.5327 | 0.4931 | 64.46 | 57.67 | 49.92 |
| Lao PDR | 1992 | 2008 | 0.487 | 0.317 | 0.331 | 0.174 | 0.4617 | 0.2729 | -0.189 | 0.8379 | 0.6917 | 99.99 | 62.29 | 22.58 |
| | | | | | Ν | Iiddle Ea | st and Nort | h Africa | | | | | | |
| Morocco | 2001 | 2007 | 0.251 | 0.145 | 0.076 | 0.048 | 0.1571 | 0.0910 | -0.066 | 0.4634 | 0.4435 | 78.71 | 71.14 | 48.58 |
| Yemen, Rep. | 1998 | 2005 | 0.425 | 0.401 | 0.323 | 0.207 | 0.3992 | 0.3449 | -0.054 | 0.7474 | 0.7106 | 60.07 | 31.37 | -7.46 |
| | | | | | Lat | in Amerio | ca and the (| Caribbe a | n | | | | | |
| Costa Rica | 2004 | 2007 | 0.388 | 0.212 | 0.256 | 0.175 | 0.3073 | 0.1888 | -0.118 | 0.3883 | 0.3730 | 60.42 | 55.43 | 47.54 |
| Ecuador | 1999 | 2008 | 0.751 | 0.597 | 0.364 | 0.226 | 0.5196 | 0.3538 | -0.166 | 0.4021 | 0.3444 | 58.10 | 31.98 | 23.32 |
| Guatemala * | 2000 | 2006 | 0.745 | 0.705 | 0.271* | 0.3 * | 0.531 | 0.5122 | -0.019 | 0.5487 | 0.5240 | 208.60 * | 110.95 * | 77.52 |
| Honduras | 2003 | 2004 | 0.776 | 0.757 | 0.627 | 0.614 | 0.7054 | 0.6883 | -0.017 | 0.5263 | 0.5199 | 86.88 | 57.87 | 49.49 |
| Mexico | 1996 | 2004 | 0.807 | 0.574 | 0.615 | 0.411 | 0.6656 | 0.4501 | -0.215 | 0.2636 | 0.2401 | 34.77 | 25.96 | 23.85 |
| Paraguay | 2001 | 2004 | 0.627 | 0.442 | 0.397 | 0.344 | 0.4983 | 0.3853 | -0.113 | 0.4404 | 0.4216 | 79.47 | 69.04 | 62.83 |
| | | | | | | Sub-S | aharan Afr | ica | | | | | | |

Appendix 1: Contribution of Rural Sector to Aggregate Poverty Reduction

| Cameroon | 1996 | 2007 | 0.596 | 0.55 | 0.414 | 0.122 | 0.5174 | 0.3372 | -0.180 | 0.5683 | 0.5027 | 34.52 | 12.83 | -1.88 |
|------------|------|------|-------|-------|--------|---------|--------|--------|--------|--------|--------|----------|-------|-------|
| Ghana | 1992 | 2006 | 0.636 | 0.392 | 0.277 | 0.108 | 0.4999 | 0.2545 | -0.245 | 0.6208 | 0.5160 | 78.49 | 51.32 | 35.78 |
| Malawi | 1998 | 2004 | 0.665 | 0.559 | 0.549 | 0.254 | 0.6487 | 0.5134 | -0.135 | 0.8593 | 0.8503 | 71.00 | 66.61 | 64.40 |
| Mali | 2001 | 2010 | 0.648 | 0.506 | 0.241 | 0.189 | 0.5313 | 0.3973 | -0.134 | 0.7132 | 0.6572 | 96.76 | 69.69 | 54.98 |
| Mozambique | 1996 | 2003 | 0.713 | 0.553 | 0.62 | 0.515 | 0.6881 | 0.5417 | -0.146 | 0.7320 | 0.7036 | 90.78 | 76.93 | 71.35 |
| Nigeria * | 2004 | 2010 | 0.734 | 0.69 | 0.522* | 0.512 * | 0.6385 | 0.6028 | -0.036 | 0.5493 | 0.5100 | 143.78 * | 62.90 | 33.59 |
| Rwanda * | 2006 | 2011 | 0.642 | 0.487 | 0.232* | 0.221* | 0.5699 | 0.4361 | -0.134 | 0.8221 | 0.8088 | 100.09 * | 93.72 | 90.17 |
| Senegal | 2001 | 2011 | 0.651 | 0.571 | 0.412 | 0.331 | 0.5542 | 0.4689 | -0.085 | 0.5950 | 0.5744 | 69.55 | 53.84 | 45.42 |
| Togo | 2006 | 2011 | 0.751 | 0.734 | 0.372 | 0.346 | 0.6159 | 0.5865 | -0.029 | 0.6435 | 0.6198 | 96.30 | 35.86 | 15.81 |
| Uganda | 2002 | 2009 | 0.427 | 0.272 | 0.144 | 0.091 | 0.3915 | 0.2453 | -0.146 | 0.8745 | 0.8522 | 96.84 | 90.33 | 81.60 |

Notes: 1. Based on Data from PovCal.Net (Downloaded from http://povcal.net.Note) and WDI Database in 2011.

2. In case of Guatemala, Nigeria and Rwanda (denoted as *), contribution of rural sector to aggregate poverty reduction exceeds 100%. This is due to the fact that urban poverty increased in Guatemala and only marginally decreased in Nigeria and Rwanda.

Source: Imai et at. (2014).

| | Country | | MPI data source | | | Population | | Multidimensional poverty - Urban | | | Multidimensional poverty - Rural | | | |
|---------------------|----------------------------------|------------------------------|-----------------|---------|--|------------|-------|--|---|---|--|--|--|--|
| ISO country code | | World region | wiri data | source | Multidimensional Poverty Index (MPI) National | | (%) | Multidimensional Poverty Index (MPI = H*A) | Headcount ratio: Population in multidimensional poverty (H) | Intensity of deprivation among the poor (A) | Multidimensional Poverty Index (MPI = H*A) | Headcount Ratio: Population in multidimensional poverty (H) | Intensity of deprivation among the poor (A) | |
| | | | Survey | Year | | Urban | Rural | Range 0 to 1 | % Population | Average % of weighted deprivations | Range 0 to 1 | % Population | Average % of weighted deprivations | |
| cup i | | East Asia and the | | | 0.074 | | | 0.004 | | | 0.010 | • • | 240 | |
| CHN | China | Pacific East Asia and the | WHS | 2002 | 0.056 | 88.4 | 11.6 | 0.004 | 1.2 | 37.5 | 0.010 | 2.8 | 36.8 | |
| IDN | Indonesia | Pacific | DHS | 2012 | 0.066 | 34.0 | 66.0 | 0.015 | 3.9 | 38.9 | 0.063 | 14.9 | 42.3 | |
| KHM | Cambodia | East Asia and the Pacific | DHS | 2010 | 0.212 | 60.4 | 39.6 | 0.020 | 5.4 | 37.6 | 0.096 | 20.2 | 47.6 | |
| LAO | Lao People's Democratic Republic | East Asia and the Pacific | MICS/DHS | 2011/12 | 0.174 | 55.5 | 44.5 | 0.021 | 4.6 | 45.4 | 0.104 | 24.6 | 42.2 | |
| MNG | Mongolia | East Asia and the Pacific | MICS | 2005 | 0.065 | 30.3 | 69.7 | 0.045 | 11.4 | 40.0 | 0.223 | 43.3 | 51.4 | |
| PHL | Philippines | East Asia and the Pacific | DHS | 2008 | 0.064 | 40.5 | 59.5 | 0.112 | 25.2 | 44.6 | 0.341 | 65.9 | 51.7 | |
| THA | Thailand | East Asia and the Pacific | MICS | 2005/06 | 0.006 | 17.7 | 82.3 | 0.183 | 39.0 | 47.0 | 0.440 | 83.2 | 52.9 | |
| TLS | Tim or-Leste | East Asia and the Pacific | DHS | 2009/10 | 0.360 | 30.6 | 69.4 | 0.194 | 41.0 | 47.4 | 0.479 | 88.5 | 54.2 | |
| VNM | Viet Nam | East Asia and the Pacific | MICS | 2011 | 0.017 | 37.9 | 62.1 | 0.262 | 49.0 | 53.5 | 0.610 | 94.7 | 64.4 | |
| VUT | Vanuatu | East Asia and the Pacific | MICS | 2007 | 0.129 | 16.4 | 83.6 | 0.276 | 53.9 | 51.2 | 0.621 | 94.7 | 65.6 | |
| ALB | Albania | Europe and Central Asia | DHS | 2008/09 | 0.005 | 66.1 | 33.9 | 0.000 | 0.0 | 36.1 | 0.000 | 0.0 | 33.3 | |
| ARM | Armenia | Europe and Central Asia | DHS | 2010 | 0.001 | 53.6 | 46.4 | 0.000 | 0.1 | 36.9 | 0.001 | 0.3 | 35.8 | |
| AZE | Azerbaijan | Europe and Central Asia | DHS | 2006 | 0.021 | 65.8 | 34.2 | 0.000 | 0.1 | 35.6 | 0.001 | 0.4 | 36.9 | |
| BIH | Bosnia and Herzegovina | Europe and Central Asia | MICS | 2011/12 | 0.002 | 38.3 | 61.7 | 0.001 | 0.4 | 35.5 | 0.006 | 1.3 | 44.6 | |
| BLR | Belarus | Europe and Central Asia | MICS | 2005 | 0.000 | 30.0 | 70.0 | 0.002 | 0.5 | 37.2 | 0.006 | 1.7 | 36.2 | |
| CZE | Czech Republic | Europe and Central Asia | WHS | 2002/03 | 0.010 | 68.1 | 31.9 | 0.006 | 1.8 | 35.2 | 0.017 | 5.0 | 33.3 | |
| EST | Estonia | Europe and Central Asia | WHS | 2002/05 | 0.026 | 71.0 | 29.0 | 0.008 | 2.5 | 33.3 | 0.022 | 5.4 | 41.2 | |
| GEO | | Europe and Central | | 2005 | 0.003 | | | | 2.3 | 37.2 | 0.022 | 6.7 | 37.2 | |
| | Georgia | Asia Europe and Central | MICS | | | 74.6 | 25.4 | 0.009 | | | | | | |
| HRV | Croatia | Asia Europe and Central | WHS | 2003 | 0.016 | 25.8 | 74.2 | 0.014 | 3.6 | 39.5 | 0.049 | 13.1 | 37.3 | |
| HUN | Hungary | Asia Europe and Central | WHS | 2003 | 0.016 | 62.8 | 37.2 | 0.015 | 4.3 | 34.9 | 0.061 | 14.4 | 42.6 | |
| KAZ | Kazakhstan | Asia Europe and Central | MICS | 2010/11 | 0.001 | 28.3 | 71.7 | 0.019 | 4.7 | 39.5 | 0.083 | 20.5 | 40.6 | |
| KGZ | Ky rgy zstan | Asia Europe and Central | MICS | 2005/06 | 0.019 | 24.0 | 76.0 | 0.020 | 5.5 | 36.3 | 0.093 | 20.7 | 45.0 | |
| LVA | Latvia | Asia | WHS | 2003 | 0.006 | 69.4 | 30.6 | 0.028 | 7.7 | 36.0 | 0.120 | 28.9 | 41.7 | |
| MDA | Moldova, Republic of | Europe and Central | DHS | 2005 | 0.007 | 56.8 | 43.2 | 0.034 | 8.3 | 40.9 | 0.159 | 36.0 | 44.2 | |

Appendix 2: Multidimensional Poverty Index (MPI) for Rural and Urban Areas at Country Levels

| | | Asia | | | | | | | | | | | |
|-----|---------------------------------|--------------------------------|----------|---------|-------|------|------|-------|------|------|-------|------|------|
| | Macedonia, The former Yugoslav | Europe and Central | | | | | | | | | | | |
| MKD | Republic of | Asia Europe and Central | MICS | 2011 | 0.002 | 44.8 | 55.2 | 0.040 | 9.2 | 43.4 | 0.198 | 39.4 | 50.2 |
| MNE | Montenegro | Asia | MICS | 2005/06 | 0.006 | 84.3 | 15.7 | 0.044 | 10.9 | 40.7 | 0.216 | 45.9 | 47.1 |
| RUS | Russian Federation | Europe and Central Asia | WHS | 2003 | 0.005 | 53.5 | 46.5 | 0.120 | 26.7 | 45.1 | 0.357 | 66.6 | 53.6 |
| SRB | Serbia | Europe and Central Asia | MICS | 2010 | 0.000 | 35.6 | 64.4 | 0.150 | 34.4 | 43.7 | 0.393 | 70.2 | 56.0 |
| | | Europe and Central | | | | | | | | | | | |
| SVK | Slovakia | Asia Europe and Central | WHS | 2003 | 0.000 | 44.8 | 55.2 | 0.164 | 35.0 | 46.8 | 0.412 | 75.0 | 54.9 |
| TJK | Tajikistan | Asia Europe and Central | DHS | 2012 | 0.054 | 31.3 | 68.7 | 0.189 | 38.2 | 49.6 | 0.445 | 81.9 | 54.3 |
| TUR | Turkey | Asia Europe and Central | DHS | 2003 | 0.028 | 34.3 | 65.7 | 0.226 | 47.0 | 48.1 | 0.504 | 85.7 | 58.8 |
| UKR | Ukraine | Asia Europe and Central | DHS | 2007 | 0.008 | 44.4 | 55.6 | 0.247 | 53.1 | 46.4 | 0.573 | 93.2 | 61.5 |
| UZB | Uzbekistan | Asia | MICS | 2006 | 0.008 | 31.4 | 68.6 | 0.262 | 52.0 | 50.4 | 0.592 | 91.4 | 64.7 |
| BLZ | Belize | Latin America and Caribbean | MICS | 2011 | 0.018 | 34.2 | 65.8 | 0.002 | 0.6 | 36.5 | 0.006 | 1.7 | 37.4 |
| BOL | Bolivia, Plurinational State of | Latin America and Caribbean | DHS | 2008 | 0.089 | 45.1 | 54.9 | 0.003 | 0.8 | 37.7 | 0.007 | 1.9 | 37.7 |
| BRA | Brazil | Latin America and Caribbean | PNDS | 2006 | 0.011 | 31.2 | 68.8 | 0.004 | 1.1 | 33.9 | 0.008 | 2.1 | 38.6 |
| COL | Colombia | Latin America and Caribbean | DHS | 2010 | 0.022 | 43.8 | 56.2 | 0.006 | 1.6 | 37.9 | 0.016 | 4.7 | 33.4 |
| DOM | Dominican Republic | Latin America and Caribbean | DHS | 2007 | 0.018 | 62.2 | 37.8 | 0.007 | 1.7 | 40.2 | 0.020 | 5.7 | 35.4 |
| ECU | Ecuador | Latin America and Caribbean | WHS | 2003 | 0.009 | 82.4 | 17.6 | 0.007 | 2.1 | 34.3 | 0.021 | 5.3 | 38.8 |
| GTM | Guatemala | Latin America and Caribbean | WHS | 2003 | 0.127 | 68.6 | 31.4 | 0.011 | 2.9 | 38.6 | 0.035 | 8.8 | 40.2 |
| GUY | Guyana | Latin America and Caribbean | DHS | 2009 | 0.030 | 27.4 | 72.6 | 0.014 | 3.9 | 35.7 | 0.036 | 9.1 | 39.7 |
| | | Latin America and | | | | | | | | | | | |
| HND | Honduras | Caribbean Latin America and | DHS | 2011/12 | 0.072 | 53.6 | 46.4 | 0.014 | 3.9 | 35.8 | 0.038 | 9.6 | 40.0 |
| HTI | Haiti | Caribbean Latin America and | DHS | 2012 | 0.248 | 58.6 | 41.4 | 0.014 | 3.9 | 37.0 | 0.051 | 11.4 | 44.3 |
| MEX | Mexico | Caribbean Latin America and | ENSANUT | 2012 | 0.011 | 49.8 | 50.2 | 0.039 | 10.2 | 38.7 | 0.193 | 43.3 | 44.5 |
| NIC | Nicaragua | Caribbean Latin America and | DHS | 2011/12 | 0.128 | 34.2 | 65.8 | 0.077 | 16.1 | 48.0 | 0.292 | 58.7 | 49.8 |
| PER | Peru | Caribbean Latin America and | DHS-Cont | 2012 | 0.043 | 63.7 | 36.3 | 0.105 | 24.5 | 42.8 | 0.324 | 57.5 | 56.4 |
| PRY | Paraguay | Caribbean Latin America and | WHS | 2002/03 | 0.064 | 14.0 | 86.0 | 0.113 | 24.9 | 45.3 | 0.348 | 61.8 | 56.3 |
| SUR | Suriname | Caribbean | MICS | 2010 | 0.024 | 38.9 | 61.1 | 0.161 | 36.2 | 44.4 | 0.408 | 77.1 | 52.9 |
| TTO | Trinidad and Tobago | Latin America and Caribbean | MICS | 2006 | 0.020 | 13.5 | 86.5 | 0.201 | 41.7 | 48.2 | 0.480 | 83.9 | 57.2 |
| URY | Uruguay | Latin America and Caribbean | WHS | 2002/03 | 0.006 | 29.1 | 70.9 | 0.252 | 52.6 | 47.9 | 0.586 | 93.5 | 62.7 |
| AFG | Afghanistan | South Asia | MICS | 2010/11 | 0.353 | 53.2 | 46.8 | 0.000 | 0.0 | 0.0 | 0.000 | 0.0 | 0.0 |
| BGD | Bangladesh | South Asia | DHS | 2011 | 0.253 | 55.7 | 44.3 | 0.001 | 0.3 | 36.3 | 0.005 | 1.5 | 35.2 |
| BTN | Bhutan | South Asia | MICS | 2010 | 0.119 | 60.6 | 39.4 | 0.004 | 1.1 | 39.1 | 0.009 | 2.1 | 44.5 |
| IND | India | South Asia | DHS | 2005/06 | 0.283 | 67.2 | 32.8 | 0.015 | 4.3 | 35.4 | 0.065 | 15.6 | 41.3 |
| LKA | Sri Lanka | South Asia | WHS | 2003 | 0.021 | 56.7 | 43.3 | 0.022 | 5.7 | 38.3 | 0.118 | 24.1 | 49.2 |
| MDV | Maldives | South Asia | DHS | 2009 | 0.018 | 22.8 | 77.2 | 0.039 | 9.8 | 39.3 | 0.193 | 43.4 | 44.4 |
| NPL | Nepal | South Asia | DHS | 2009 | 0.217 | 32.3 | 67.7 | 0.087 | 20.0 | 43.3 | 0.299 | 55.7 | 53.6 |
| PAK | Pakistan | South Asia | DHS | 2012/13 | 0.230 | 48.0 | 52.0 | 0.091 | 19.9 | 45.6 | 0.316 | 66.4 | 47.5 |
| BDI | Burundi | Sub-Saharan Africa | DHS | 2010 | 0.454 | 55.0 | 45.0 | 0.001 | 0.1 | 39.8 | 0.002 | 0.4 | 38.0 |
| BEN | Benin | Sub-Saharan Africa | DHS | 2006 | 0.412 | 61.7 | 38.3 | 0.001 | 0.2 | 33.3 | 0.002 | 0.5 | 39.1 |
| | | | | | | | | | | | | | |

| BFA | Burkina Faso | Sub-Saharan Africa | DHS | 2010 | 0.535 | 78.2 | 21.8 | 0.001 | 0.3 | 33.8 | 0.004 | 1.2 | 35.5 |
|-----|-----------------------------------|--------------------|------|---------|-------|------|------|-------|------|------|-------|------|------|
| CAF | Central African Republic | Sub-Saharan Africa | MICS | 2010 | 0.430 | 77.7 | 22.3 | 0.004 | 1.1 | 37.4 | 0.010 | 2.8 | 34.5 |
| CIV | Cote d'Ivoire | Sub-Saharan Africa | DHS | 2011/12 | 0.310 | 55.4 | 44.6 | 0.004 | 1.2 | 37.2 | 0.010 | 2.9 | 36.6 |
| CMR | Cameroon | Sub-Saharan Africa | DHS | 2011 | 0.248 | 29.6 | 70.4 | 0.005 | 1.3 | 37.2 | 0.011 | 3.0 | 35.9 |
| COD | Congo, Democratic Republic of the | Sub-Saharan Africa | MICS | 2010 | 0.392 | 31.1 | 68.9 | 0.005 | 1.6 | 33.3 | 0.012 | 3.2 | 38.7 |
| COG | Congo, Republic of | Sub-Saharan Africa | DHS | 2011/12 | 0.181 | 92.4 | 7.6 | 0.006 | 1.6 | 34.7 | 0.013 | 3.8 | 34.7 |
| ETH | Ethiopia | Sub-Saharan Africa | DHS | 2011 | 0.564 | 64.6 | 35.4 | 0.008 | 2.2 | 37.7 | 0.024 | 6.8 | 35.8 |
| GAB | Gabon | Sub-Saharan Africa | DHS | 2012 | 0.070 | 84.9 | 15.1 | 0.008 | 2.2 | 38.5 | 0.025 | 6.3 | 38.8 |
| GHA | Ghana | Sub-Saharan Africa | MICS | 2011 | 0.139 | 69.1 | 30.9 | 0.009 | 2.4 | 38.3 | 0.028 | 7.0 | 39.9 |
| GIN | Guinea | Sub-Saharan Africa | DHS | 2005 | 0.506 | 55.0 | 45.0 | 0.009 | 2.5 | 37.1 | 0.029 | 7.4 | 38.6 |
| GMB | Gambia | Sub-Saharan Africa | MICS | 2005/06 | 0.324 | 42.6 | 57.4 | 0.010 | 2.5 | 37.6 | 0.034 | 8.5 | 39.5 |
| GNB | Guinea-Bissau | Sub-Saharan Africa | MICS | 2006 | 0.462 | 64.5 | 35.5 | 0.011 | 3.1 | 35.3 | 0.035 | 8.5 | 41.4 |
| KEN | Kenya | Sub-Saharan Africa | DHS | 2008/09 | 0.229 | 59.1 | 40.9 | 0.019 | 4.9 | 39.4 | 0.087 | 17.3 | 50.1 |
| LBR | Liberia | Sub-Saharan Africa | DHS | 2007 | 0.485 | 48.1 | 51.9 | 0.022 | 5.2 | 41.6 | 0.113 | 27.1 | 41.6 |
| LSO | Lesotho | Sub-Saharan Africa | DHS | 2009 | 0.156 | 22.9 | 77.1 | 0.025 | 6.5 | 38.0 | 0.119 | 25.7 | 46.5 |
| MDG | Madagascar | Sub-Saharan Africa | DHS | 2008/09 | 0.357 | 24.0 | 76.0 | 0.038 | 9.7 | 38.8 | 0.191 | 43.0 | 44.4 |
| MLI | Mali | Sub-Saharan Africa | DHS | 2006 | 0.558 | 50.4 | 49.6 | 0.041 | 9.6 | 42.5 | 0.209 | 46.6 | 44.8 |
| MOZ | Mozambique | Sub-Saharan Africa | DHS | 2011 | 0.389 | 18.1 | 81.9 | 0.051 | 12.0 | 42.6 | 0.227 | 51.1 | 44.4 |
| MRT | Mauritania | Sub-Saharan Africa | MICS | 2007 | 0.352 | 48.2 | 51.8 | 0.057 | 13.8 | 41.4 | 0.238 | 48.4 | 49.2 |
| MWI | Malawi | Sub-Saharan Africa | DHS | 2010 | 0.334 | 19.4 | 80.6 | 0.066 | 15.6 | 42.1 | 0.247 | 53.4 | 46.3 |
| NAM | Namibia | Sub-Saharan Africa | DHS | 2006/07 | 0.187 | 40.4 | 59.6 | 0.066 | 15.3 | 43.3 | 0.252 | 53.9 | 46.7 |
| NER | Niger | Sub-Saharan Africa | DHS | 2012 | 0.605 | 12.8 | 87.2 | 0.069 | 15.4 | 44.8 | 0.268 | 56.0 | 47.9 |
| NGA | Nigeria | Sub-Saharan Africa | MICS | 2011 | 0.240 | 37.4 | 62.6 | 0.074 | 17.6 | 41.9 | 0.269 | 55.6 | 48.4 |
| RWA | Rwanda | Sub-Saharan Africa | DHS | 2010 | 0.350 | 96.3 | 3.7 | 0.127 | 27.2 | 46.6 | 0.366 | 72.0 | 50.8 |
| SEN | Senegal | Sub-Saharan Africa | DHS | 2010/11 | 0.439 | 23.9 | 76.1 | 0.129 | 27.7 | 46.5 | 0.374 | 73.2 | 51.0 |
| SLE | Sierra Leone | Sub-Saharan Africa | MICS | 2010 | 0.388 | 14.9 | 85.1 | 0.133 | 29.2 | 45.5 | 0.385 | 74.7 | 51.5 |
| STP | Sao Tome and Principe | Sub-Saharan Africa | DHS | 2008/09 | 0.154 | 22.8 | 77.2 | 0.154 | 34.6 | 44.7 | 0.396 | 73.7 | 53.8 |
| SWZ | Swaziland | Sub-Saharan Africa | MICS | 2010 | 0.086 | 27.8 | 72.2 | 0.168 | 31.5 | 53.4 | 0.414 | 77.0 | 53.7 |
| TCD | Chad | Sub-Saharan Africa | WHS | 2003 | 0.344 | 16.6 | 83.4 | 0.175 | 39.7 | 44.2 | 0.427 | 75.8 | 56.4 |
| TGO | Togo | Sub-Saharan Africa | MICS | 2010 | 0.250 | 22.5 | 77.5 | 0.176 | 37.3 | 47.3 | 0.429 | 78.0 | 55.0 |
| TZA | Tanzania, United Republic of | Sub-Saharan Africa | DHS | 2010 | 0.332 | 21.8 | 78.2 | 0.226 | 45.8 | 49.4 | 0.506 | 84.4 | 60.0 |
| UGA | Uganda | Sub-Saharan Africa | DHS | 2011 | 0.367 | 17.9 | 82.1 | 0.230 | 46.4 | 49.5 | 0.519 | 89.8 | 57.8 |
| ZAF | South Africa | Sub-Saharan Africa | NIDS | 2012 | 0.044 | 38.4 | 61.6 | 0.293 | 55.9 | 52.4 | 0.645 | 96.2 | 67.0 |
| ZMB | Zambia | Sub-Saharan Africa | DHS | 2007 | 0.328 | 37.5 | 62.5 | 0.337 | 68.5 | 49.2 | 0.651 | 96.9 | 67.2 |
| ZWE | Zimbabwe | Sub-Saharan Africa | DHS | 2010/11 | 0.172 | 30.4 | 69.6 | 0.360 | 64.8 | 55.5 | 0.669 | 96.2 | 69.5 |
| | | | | | | | | | | | | | |

Source: Alkire et al. (2014). http://www.ophi.org.uk/multidimensional-poverty-index/mpi-2014/mpi-data/

Appendix 3: Deriving Vulnerability Measure

Vulnerability measure as an expected poverty is specified as:

$$VEP_{it} \equiv V_{it} = Pr(c_{i,t+1} \le z)$$
(A1)

where vulnerability of household *i* at time *t*, V_{it} , is the probability that the *i*-th household's level of consumption at time t+1, $c_{i,t+1}$, will be below the poverty line, z.

Three limitations, amongst others, should be noted in our measure of vulnerability. First, the present analysis is confined to a consumption (used synonymously with income) threshold of poverty. Second, our measure of vulnerability in terms of the probability of a household's consumption falling below the poverty threshold in the future is subject to the choice of a threshold. Third, while income/consumption volatility underlies vulnerability, the resilience in mitigating welfare losses depends on assets defined broadly-including human, physical and social capital. A household with inadequate physical or financial asset or savings, for example, may find it hard to overcome loss of income. This may translate into lower nutritional intake and rationing out of its members from the labor market (Dasgupta, 1997; Foster, 1995). Lack of physical assets may also impede accumulation of profitable portfolios under risk and generate poverty traps.

The consumption function is estimated by the equation (A2).¹¹

$$\ln c_i = X_i \lambda + e_i \tag{A2}$$

where c_i is mean per capita consumption (i.e. food and non-food consumption expenditure) for the household and X is a vector of observable household characteristics and other determinants of consumption. It is further assumed that the structure of the economy is relatively stable over time and, hence, future consumption stems solely from the uncertainty about the idiosyncratic shocks, e_i . It is also assumed that the variance of the disturbance term depends on:

$$\sigma_{e,i}^2 = X_i \theta \tag{A3}$$

The estimates of β and θ are obtained using a three-step feasible generalized least squares (FGLS)¹². Using the estimates $\hat{\beta}$ and $\hat{\theta}$, we can compute the expected log consumption and the variance of log consumption for each household as follows.

$$E[\ln C_i | X_i] = X_i \hat{\beta} \tag{A4}$$

$$V[\ln C_i | X_i] = X_i \theta \tag{A5}$$

¹¹ We have used White-Huber sandwich estimator to overcome heteroscedasticity in the sample.

¹² See Chaudhuri (2003), Chaudhuri et al. (2002), and Hoddinott and Quisumbing (2003) for technical details. This is summarised in Appendix 3.

By assuming lnc_i as normally distributed and letting $\Phi(\cdot)$ denote the cumulative density function of the standard normal distribution, the estimated probability that a household will be poor in the future (say, at time t+1) is given by:

$$\hat{\text{VEP}}_{i} \equiv \hat{v}_{i} = \hat{P}r\left(\ln c_{i} < \ln z | X_{i}\right) = \Phi\left(\frac{\ln z - X_{i}\hat{\beta}}{\sqrt{X_{i}\hat{\theta}}}\right)$$
(A6)

This is an *ex ante* vulnerability measure that can be estimated with cross-sectional data. Note that this expression also yields the probability of a household at time *t* becoming poor at t+1 given the distribution of consumption at *t*.

A merit of this vulnerability measure is that it can be estimated with cross-sectional data (e.g. Imai et al., 2011). However, it correctly reflects a household's vulnerability only if the distribution of consumption across households, given the household characteristics at time t, represents time-series variation of household consumption. Hence this measure requires a large sample in which some households experience positive shocks while others suffer from negative shocks. Also, the measure is unlikely to reflect unexpected large negative shocks (e.g., Asian financial crisis), if we use the cross-section data for a normal year.